



Towards more interpretable kernel-based sensitivity analysis

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0 ■ Introduction

GSA in support to metamodel construction


➤ In all four work packages, there is a need to construct **metamodels** for **high-dimensional design problems**.


- Let $\mathbf{X} := [X_1, \dots, X_d]$ be a random vector with **independent** components ($d \approx 100$).
- Let $Y := g(\mathbf{X})$ where $g: \mathcal{X}_1 \times \dots \times \mathcal{X}_d \rightarrow \mathcal{Y}$ is a **computationally-expensive** simulation code.
- $\mathbf{Z} = (\mathbf{X}, Y)$ is the **augmented vector** containing the input and output variables.

 **The design of experiments (DoE) consists of a number of input-output observations.**

➤ The metamodel \hat{g} is constructed from $\mathbf{Z}_{\text{obs}} := \{(\mathbf{X}^{(i)}, Y^{(i)})\}_{1 \leq i \leq N_{\text{sim}}}$ with $N_{\text{sim}} \leq 10d \rightarrow$ **SMALL DATA**.

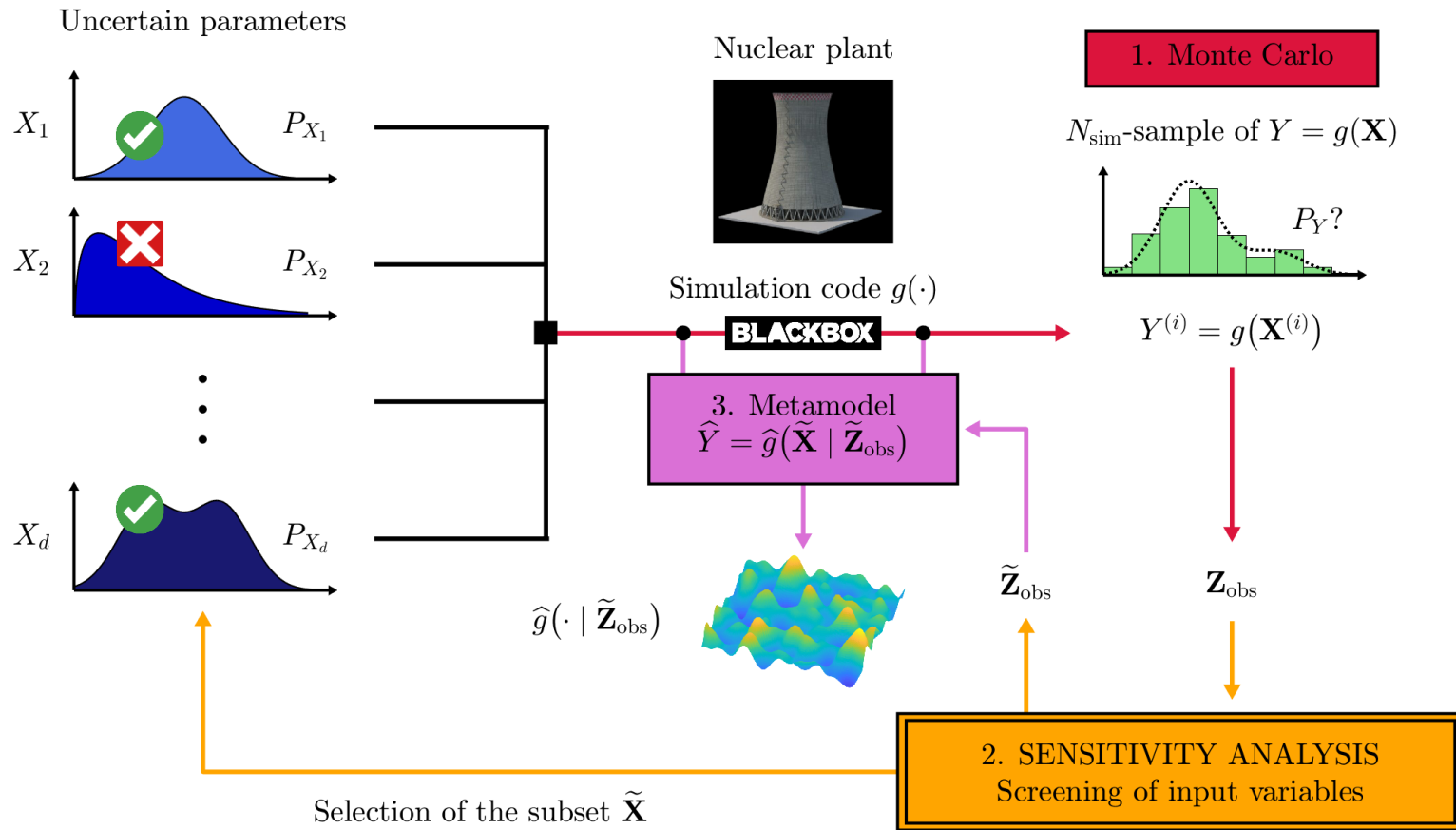
➤ For a nice coverage of the input domain of variation, the DoE must be space-filling \rightarrow **GIVEN DATA**.

 Classical metamodeling techniques (such as **GP regression**) cannot be used directly.
Curse of dimensionality \rightarrow too many **GP hyperparameters** have to be optimized!

 **Many existing strategies (screening, additive and ANOVA models, linear and nonlinear embeddings).**
 \rightarrow Binois & Wycoff (2022) for a comprehensive review.

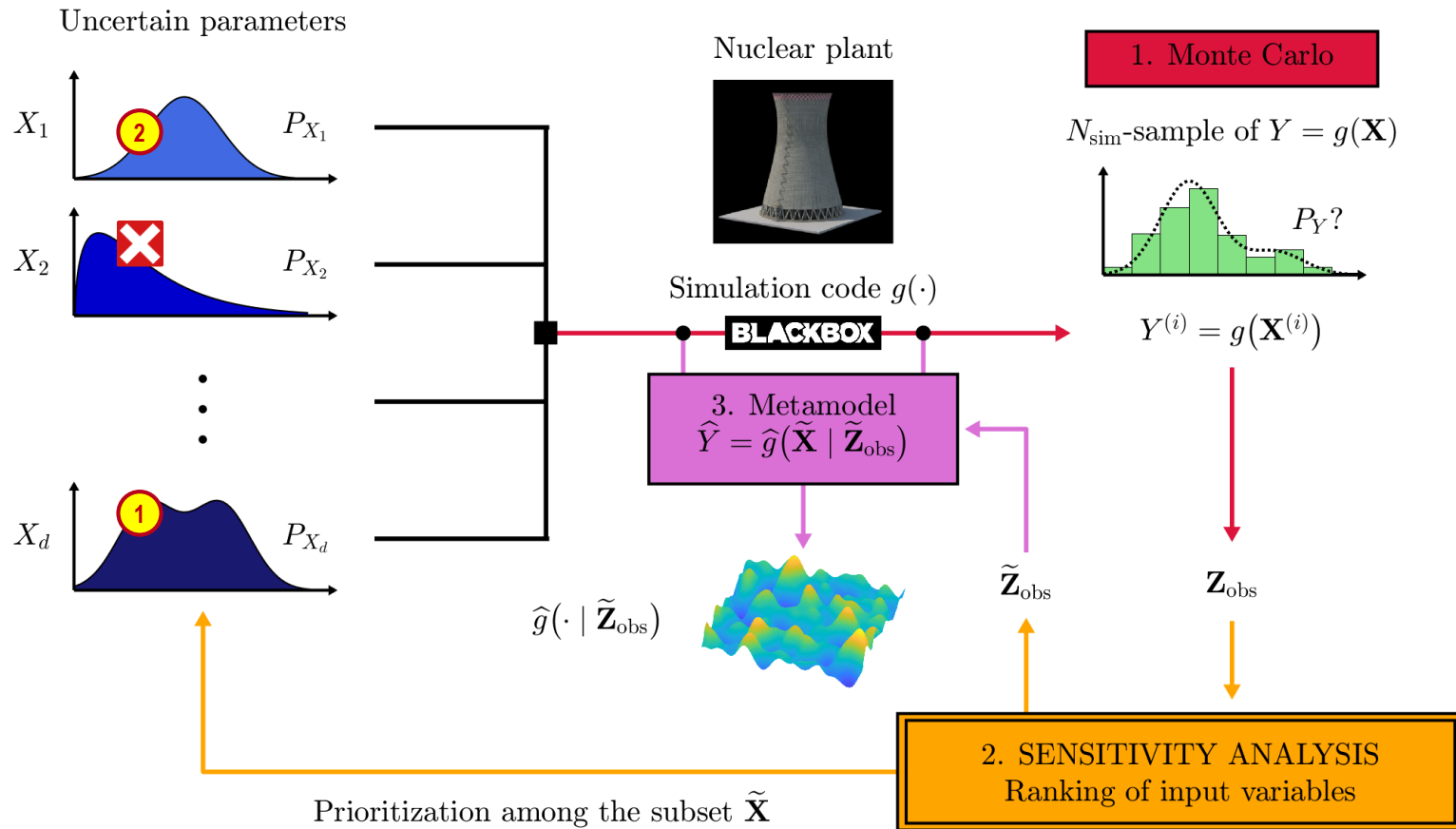
Focus on **SCREENING** \rightarrow preliminary GSA for **variable selection** (and thus **dimension reduction**).

GSA in support to metamodel construction



- Steps 2 and 3 of the **ICSCREAM** methodology → [Iooss & Marrel \(2019\)](#) or [Marrel et al. \(2020\)](#)
 - ✓ Identification of penalizing Configurations using **SCRE**ening **And** **Meta**model
- Performing a **preliminary GSA** has two main advantages.
 - **Screening-oriented GSA** → (crude) dimension reduction by discarding non-influential input variables.

GSA in support to metamodel construction



- Steps 2 and 3 of the **ICSCREAM** methodology → looss & Marrel (2019) or Marrel et al. (2020)
 - ✓ Identification of penalizing **C**onfigurations using **S**CReening **A**nd **M**etamodel
- Performing a **preliminary GSA** has two main advantages.
 - **Screening-oriented GSA** → (crude) dimension reduction by discarding non-influential input variables.
 - **Ranking-oriented GSA** → sequential building process of the GP metamodel.

Summary

1. A few concepts related to kernels
2. Sensitivity measures based on the HSIC
3. A bridge between two opposite worlds: HSIC-ANOVA indices
4. Is it relevant to talk about interactions for HSIC-ANOVA indices?
5. More about Sobolev kernels and their properties
6. Does all this benefit independence testing?



1 ■ A few concepts related to kernels

Fundamentals of reproducing kernel theory

1 Kernel mean embeddings → Muandet et al. (2017)

- Let $\mathcal{M}_1^+(\mathcal{Z})$ be the space of all probability measures defined on $\mathcal{Z} \subseteq \mathbb{R}^p$.
- Let $K: \mathcal{Z} \times \mathcal{Z} \rightarrow \mathbb{R}$ be a kernel and let \mathcal{H} be the induced RKHS.

- Any probability measure $\nu \in \mathcal{M}_1^+(\mathcal{Z})$ can be represented by a (well-defined) function $\mu_\nu \in \mathcal{H}$.

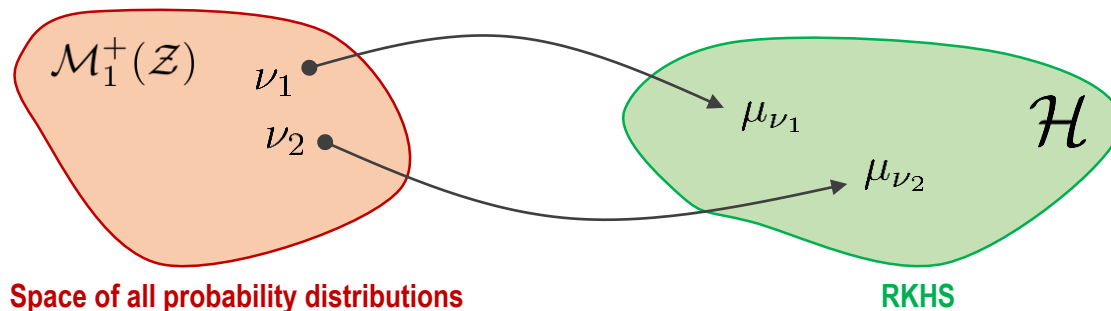
$$\mu_\nu : \mathcal{Z} \rightarrow \mathbb{R}$$

$$z \mapsto \mu_\nu(z) = \mathbb{E}_\nu[K(z, Z)] = \int_{\mathcal{Z}} K(z, \zeta) d\nu(\zeta)$$

Assumptions

- K must be measurable
- $\mathbb{E}_\nu[\sqrt{K(Z, Z)}] < \infty$

- K is said to be a **characteristic kernel** if the map $\nu \mapsto \mu_\nu$ is injective.



- The **dissimilarity** between ν_1 and ν_2 can be measured through the **distance** in \mathcal{H} between μ_{ν_1} and μ_{ν_2} .
 - ✓ Definition of a **kernel-based dissimilarity measure** on $\mathcal{M}_1^+(\mathcal{Z})$.

Fundamentals of reproducing kernel theory

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$$\mu_\nu : \mathcal{Z} \longrightarrow \mathbb{R}$$

$$z \longmapsto \mu_\nu(z) = \mathbb{E}_\nu [K(z, Z)] = \int_{\mathcal{Z}} K(z, \zeta) d\nu(\zeta)$$

Assumptions

- K must be measurable
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➤ K is said to be a **characteristic kernel** if the map $\nu \mapsto \mu_\nu$ is injective.

2 Maximum Mean Discrepancy (MMD) → Gretton *et al.* (2006)

$$\text{MMD}^2(\nu_1, \nu_2) = \|\mu_{\nu_1} - \mu_{\nu_2}\|_{\mathcal{H}}^2 \quad \checkmark \quad \text{Definition resulting from the embedding mechanism}$$

$$= \mathbb{E}_{\nu_1 \otimes \nu_1} [K(Z, Z')] + \mathbb{E}_{\nu_2 \otimes \nu_2} [K(Z, Z')] - 2 \mathbb{E}_{\nu_1 \otimes \nu_2} [K(Z, Z')]$$

✓ Alternative formula paving the way to a simple estimation procedure

Fundamentals of reproducing kernel theory

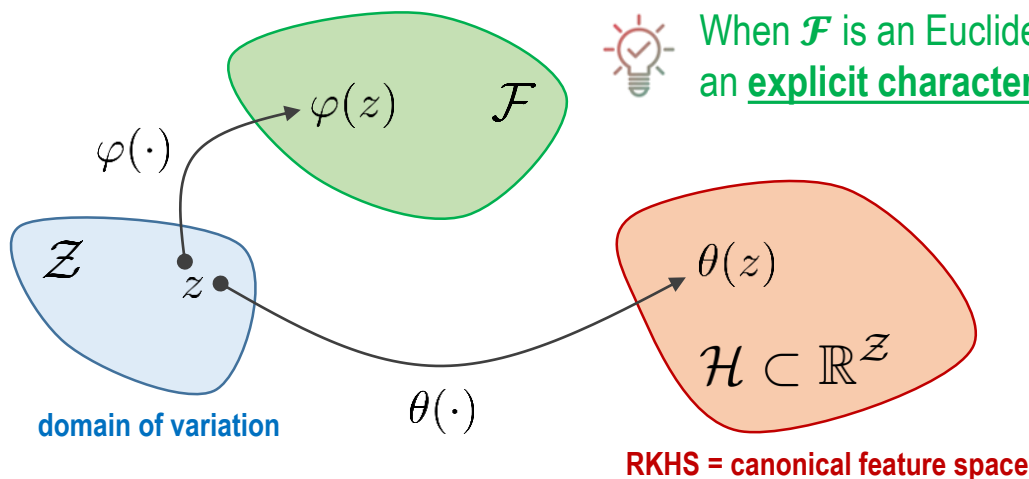
3 Feature maps → Chapter 4 in Steinwart & Christmann (2008)

- Let $K: \mathcal{Z} \times \mathcal{Z} \rightarrow \mathbb{R}$ be a kernel and let \mathcal{H} be the induced RKHS.
- Let us assume that there exist a Hilbert space \mathcal{F} and a map $\varphi: \mathcal{Z} \rightarrow \mathcal{F}$ such that:

$$\forall z, z' \in \mathcal{Z}, K(z, z') = \langle \varphi(z), \varphi(z') \rangle_{\mathcal{F}}$$

\mathcal{F} is called a **feature space**. φ is called a **feature map**. Any object $\varphi(z)$ is called a **feature function**.

- **Existence** of at least one feature map.
 - ✓ The **canonical feature map** $\theta: \mathcal{Z} \rightarrow \mathcal{H}$ is thus defined by $\theta(z) := K(\cdot, z)$ for any $z \in \mathcal{Z}$.
- **Non-uniquity** of the feature map.
 - ✓ There may exist a feature space where the kernel action is **much easier to understand**.



When \mathcal{F} is an Euclidean space or a sequence space, an **explicit characterization of the RKHS** can be derived.



Most often, $\theta(\cdot)$ is **NOT informative!**

Fundamentals of reproducing kernel theory

4 Feature-based characterization of the RKHS → Chapter 4 in Steinwart & Christmann (2008)

 First, let us examine **two particular kernels!**

Example 1 ➤ The **polynomial kernel** with position parameter $c \geq 0$ and exponent $m \in \mathbb{N}^*$.

initial definition

$$K_{\text{poly}}(x, x') := (xx' + c)^m = \sum_{k=0}^m \binom{m}{k} x^k (x')^k c^{m-k}$$

$$= \langle \varphi_{\text{poly}}(x), \varphi_{\text{poly}}(x') \rangle_{\mathbb{R}^{m+1}} \quad \text{with}$$

$$\varphi_{\text{poly}}(x) = \left[(\sqrt{c})^{m-k} \sqrt{\binom{m}{k}} x^k \right]_{0 \leq k \leq m}$$

finite number of polynomial features

✓ The **binomial theorem** reveals a feature map φ_{poly} from \mathbb{R} to the **Euclidean** space \mathbb{R}^{m+1} .

Fundamentals of reproducing kernel theory

4 Feature-based characterization of the RKHS → Chapter 4 in Steinwart & Christmann (2008)

 First, let us examine **two particular kernels!**

Example 2 ➤ The **Gaussian kernel** with scale parameter $\gamma > 0$.

initial definition

$$K_\gamma(x, x') := e^{-\frac{1}{2}\left(\frac{x-x'}{\gamma}\right)^2} = e^{-\frac{1}{2}\left(\frac{x}{\gamma}\right)^2} e^{-\frac{1}{2}\left(\frac{x'}{\gamma}\right)^2} \sum_{k=0}^{\infty} \frac{1}{k!} \left(\frac{x}{\gamma}\right)^k \left(\frac{x'}{\gamma}\right)^k$$

$$= \langle \varphi_\gamma(x), \varphi_\gamma(x') \rangle_{\ell^2} \quad \text{with} \quad \varphi_\gamma(x) := e^{-\frac{1}{2}\left(\frac{x}{\gamma}\right)^2} \left[\frac{1}{\sqrt{k}} \left(\frac{x}{\gamma}\right)^k \right]_{k \geq 0}$$

infinite number of damped polynomial features

✓ The **Taylor series expansion** reveals a feature map φ_γ from \mathbb{R} into the **Hilbert** space $\ell^2(\mathbb{N})$.

Fundamentals of reproducing kernel theory

4 Feature-based characterization of the RKHS → Chapter 4 in Steinwart & Christmann (2008)

- As shown in these two examples, a **kernel expansion** allows to identify a **feature map**.
 - ✓ More importantly, it provides **all-in-one characterization** of the RKHS.

→ Let $K: \mathcal{Z} \times \mathcal{Z} \rightarrow \mathbb{R}$ be a kernel and let \mathcal{H} be the induced RKHS.

- It is assumed that it can be expanded as a sum (or series) of **symmetric** and **separable** functions.

$$\forall z, z' \in \mathcal{Z}, \quad K(z, z') = \sum_{i \in I} g_i(z) g_i(z')$$

Polynomial kernel → $I = \{0, \dots, m\}$

Gaussian kernel → $I = \mathbb{N}$

- ✓ The functions $(g_i)_{i \in I}$ are the **features**. They must be linearly independent (in the ℓ^2 -sense).

$$\textcircled{1} \quad \mathcal{H} = \left\{ h \in \mathbb{R}^{\mathcal{Z}} \mid h(\cdot) = \sum_{i \in I} a_i g_i(\cdot) \text{ with } (a_i)_{i \in I} \in \ell^2(I, \mathbb{R}) \right\}$$

$$\textcircled{2} \quad \langle \cdot, \cdot \rangle_{\mathcal{H}} : \quad \mathcal{H} \quad \times \quad \mathcal{H} \quad \longrightarrow \quad \mathbb{R}$$
$$\left(h_1(\cdot) = \sum_{i \in I} a_i g_i(\cdot) \text{ , } h_2(\cdot) = \sum_{i \in I} b_i g_i(\cdot) \right) \longmapsto \sum_{i \in I} a_i b_i$$

- ③ The functions $(g_i)_{i \in I}$ form an **orthonormal basis (ONB)** of \mathcal{H} .



2 ■ Sensitivity measures based on the HSIC

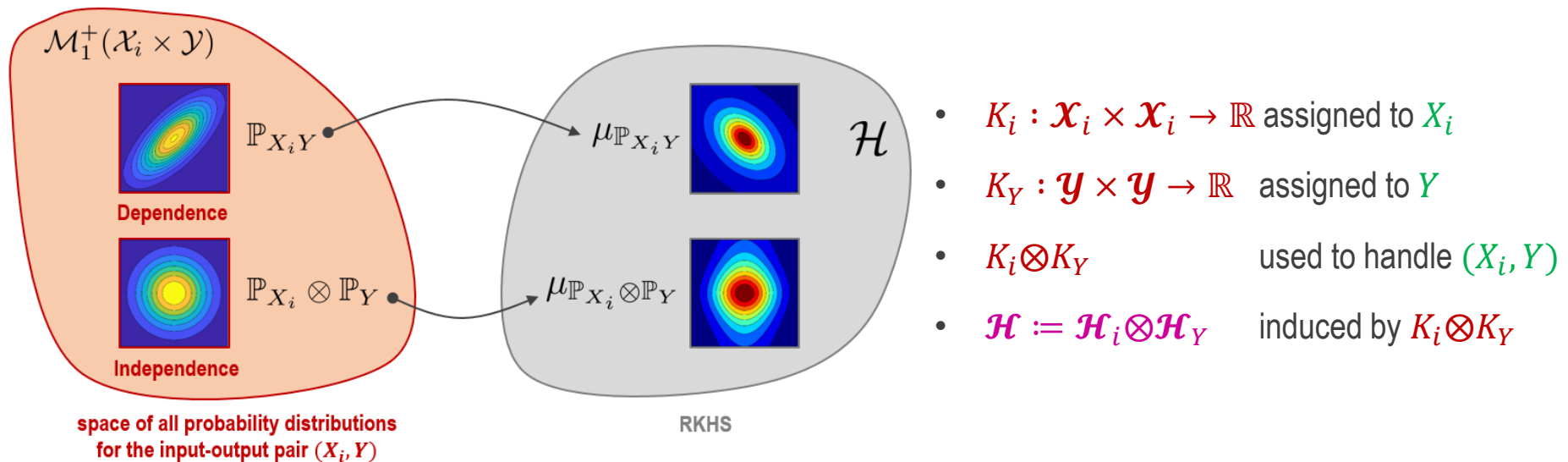
Several views on HSIC indices

1 Kernel-based dependences measures → Da Veiga (2015)

- $\mathbb{P}_{X_i Y}$ → Joint distribution of (X_i, Y) → True influence of X_i on Y
- $\mathbb{P}_{X_i} \otimes \mathbb{P}_Y$ → Independence within (X_i, Y) → Hypothetical lack of influence

$$S_i^\Delta := \Delta(\mathbb{P}_{X_i Y}, \mathbb{P}_{X_i} \otimes \mathbb{P}_Y)$$

🤔 How to measure the discrepancy?
What about using the MMD?



$$\text{HSIC}(X_i, Y) := \text{MMD}^2(\mathbb{P}_{X_i Y}, \mathbb{P}_{X_i} \otimes \mathbb{P}_Y) = \|\mu_{\mathbb{P}_{X_i Y}} - \mu_{\mathbb{P}_{X_i} \otimes \mathbb{P}_Y}\|_{\mathcal{H}}^2$$

🤔 How to estimate HSIC indices from a given dataset?

Several views on HSIC indices

2 Efficient estimation → Gretton *et al.* (2005, 2007) and Serfling (2009)

- The alternative formula of the MMD allows to rewrite the HSIC only in terms of kernel-based moments.

$$\text{HSIC}(X_i, Y) = \mathbb{E} \left[K_i(X_i, X'_i) K_Y(Y, Y') \right] + \mathbb{E} \left[K_i(X_i, X'_i) K_Y(Y'', Y''') \right] - 2 \mathbb{E} \left[K_i(X_i, X'_i) K_Y(Y, Y'') \right]$$

⚠ $(X_i, Y) \perp (X'_i, Y') \perp (X''_i, Y'') \perp (X'''_i, Y''')$ follow the **joint input-output distribution** $\mathbb{P}_{X_i Y}$.

- **U-statistics** and **V-statistics** are well-adapted to estimate HSIC indices from a given DoE.

$$N_{\text{sim}} = n$$

$$\hat{H}_i^U = \frac{1}{(n)_2} \sum_{1 \leq p \neq q \leq n} K_i(X_i^{(p)}, X_i^{(q)}) K_Y(Y^{(p)}, Y^{(q)}) + \frac{1}{(n)_4} \sum_{1 \leq p \neq q \neq r \neq s \leq n} K_i(X_i^{(p)}, X_i^{(q)}) K_Y(Y^{(r)}, Y^{(s)}) - \frac{2}{(n)_3} \sum_{1 \leq p \neq q \neq r \leq n} K_i(X_i^{(p)}, X_i^{(q)}) K_Y(Y^{(p)}, Y^{(r)}) \quad \text{with} \quad (n)_p = p! \binom{n}{p}$$

- \hat{H}_i^U denotes the **U-statistic** estimator of $\text{HSIC}(X_i, Y)$ → **no bias** BUT **no guarantee of positivity**.
- \hat{H}_i^V denotes the **V-statistic** estimator of $\text{HSIC}(X_i, Y)$ → **positivity** BUT **bias**.
- **Consistency** and **existence of a CLT** → convergence at rate $1/\sqrt{n}$.
- **Low computational complexity** → only $\mathcal{O}(n^2)$ operations are required to compute estimates.

Several views on HSIC indices

3 Independence testing → Gretton *et al.* (2007)

- The input kernel $K_i : \mathcal{X}_i \times \mathcal{X}_i \rightarrow \mathbb{R}$ is assumed to be **characteristic** to $\mathcal{M}_1^+(\mathcal{X}_i)$.
- The output kernel $K_Y : \mathcal{Y} \times \mathcal{Y} \rightarrow \mathbb{R}$ is assumed to be **characteristic** to $\mathcal{M}_1^+(\mathcal{Y})$.

$$X_i \perp Y \iff \text{HSIC}(X_i, Y) = 0$$

- **Testing independence** between X_i and Y is equivalent to **testing the nullity** of the HSIC.

$$(H_0) : \text{HSIC}(X_i, Y) = 0 \quad \text{vs.} \quad (H_1) : \text{HSIC}(X_i, Y) > 0$$

- **Test statistic** → either \hat{H}_i^U or \hat{H}_i^V
- **Test procedure** → selected according to the **sample size** and the chosen **test statistic**
 - ✓ **Asymptotic test procedure** → Sejdinovic *et al.* (2013) and Zhang *et al.* (2018)
 - ✓ **Permutation-based test procedure** → De Lozzo & Marrel (2016)
 - ✓ **Sequential permutation-based test procedure** → El Amri & Marrel (2022)
 - ✓ **Non-asymptotic Gamma test procedure** → El Amri & Marrel (2024)
- **Theoretical guarantees** → control of the **type-I error** + minimization of the **type-II error**
 - ✓ **Type-I error** controlled even when n is small. → Albert *et al.* (2022)
 - ✓ **Type-II error** vanishing asymptotically. → Gretton *et al.* (2007) and Pfister *et al.* (2018)

Sobol' indices vs. HSIC indices



- HSIC indices perfectly meet the needs of **screening-oriented** GSA.
 - ✓ The use of characteristic kernels allows to detect any type of input-output dependence.
 - ✓ Inference is an easy task (no need for specific data, big data or density estimation).

GSA requirements	S_i	T_i	HSIC(X_i, Y)
ANOVA decomposition → RANKING	✓	✓	✗
Characterize independence → SCREENING	✗	✓	✓
Estimation from GIVEN DATA	✓	✗	✓
Estimation from SMALL DATA	✓	✗	✓
Compatibility with DEPENDENT inputs	✗	✗	✓
INVARIANCE through monotonic transformations	✓	✓	✗

Still room to improve HSIC indices?

➤ HSIC indices lack interpretability and they are not tailored to perform **ranking-oriented** GSA.

⚠ Sum not equal to 1.

⚠ No universal bound.

⚠ Different MMD scales.

GSA requirements	S_i	T_i	HSIC(X_i, Y)
ANOVA decomposition → RANKING	✓	✓	✗
Characterize independence → SCREENING	✗	✓	✓
Estimation from GIVEN DATA	✓	✗	✓
Estimation from SMALL DATA	✓	✗	✓
Compatibility with DEPENDENT inputs	✗	✗	✓
INVARIANCE through monotonic transformations	✓	✓	✗



How to do better
on that point?



3 ■ A bridge between two opposite worlds: HSIC-ANOVA indices

Taking inspiration from standard ANOVA...



➤ ANOVA decomposition for Sobol' indices → Sobol' (1993)

- ✓ The output variance $\mathbb{V}(Y)$ is apportioned between **all subsets of inputs**.

$$\mathbb{V}(Y) = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} V_{\mathbf{u}} = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} \sum_{\mathbf{v} \subseteq \mathbf{u}} (-1)^{|\mathbf{u}|-|\mathbf{v}|} \mathbb{V}(\mathbb{E}[Y | X_{\mathbf{u}}])$$

 $X_1 \perp \dots \perp X_d$

➤ First-order and total-order Sobol' indices

- ✓ **First-order** Sobol' indices $(S_i)_{1 \leq i \leq d}$ → **main effects only!**
- ✓ **Total-order** Sobol' indices $(T_i)_{1 \leq i \leq d}$ → **main effects + interactions.**

$$\forall 1 \leq i \leq d, \quad S_i = \frac{\mathbb{V}(\mathbb{E}[Y | X_i])}{\mathbb{V}(Y)} \quad \text{and} \quad T_i = 1 - \frac{\mathbb{V}(\mathbb{E}[Y | \mathbf{X}_{-i}])}{\mathbb{V}(Y)}$$

➤ Constraints imposed on the sub-functions of the Sobol'-Hoeffding decomposition

$$g(\mathbf{x}) = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} \eta_{\mathbf{u}}(\mathbf{x}_{\mathbf{u}}) \quad \text{such that} \quad \forall i \in \mathbf{u}, \quad \int_{\mathcal{X}_i} \eta_{\mathbf{u}}(\mathbf{x}_{\mathbf{u}}) d\mathbb{P}_{X_i}(x_i) = 0$$

... and bringing ANOVA into the HSIC paradigm



➤ HSIC-ANOVA decomposition → Da Veiga (2021)

- ✓ The quantity $\text{HSIC}(\mathbf{X}, Y)$ is apportioned between **all subsets of inputs**.

$$\text{HSIC}(\mathbf{X}, Y) = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} H_{\mathbf{u}} = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} \sum_{\mathbf{v} \subseteq \mathbf{u}} (-1)^{|\mathbf{u}|-|\mathbf{v}|} \text{HSIC}(\mathbf{X}_{\mathbf{v}}, Y)$$

 $X_1 \perp \dots \perp X_d$

➤ First-order and total-order HSIC-ANOVA indices

- ✓ **First-order** HSIC-ANOVA indices $(S_i^{\text{HSIC}})_{1 \leq i \leq d} \rightarrow$ **main effects only!**
- ✓ **Total-order** HSIC-ANOVA indices $(T_i^{\text{HSIC}})_{1 \leq i \leq d} \rightarrow$ **main effects + interactions.**

$$\forall 1 \leq i \leq d, \quad S_i^{\text{HSIC}} := \frac{\text{HSIC}(X_i, Y)}{\text{HSIC}(\mathbf{X}, Y)} \quad \text{and} \quad T_i^{\text{HSIC}} := 1 - \frac{\text{HSIC}(\mathbf{X}_{-i}, Y)}{\text{HSIC}(\mathbf{X}, Y)}$$

➤ Constraints imposed on the input kernels

- ✓ Each input kernel K_i must be an **ANOVA kernel** (\approx a **constant** kernel + an **orthogonal** kernel).

$$K_i(x_i, x'_i) = 1 + k_i(x_i, x'_i) \quad \text{with} \quad \forall x_i \in \mathcal{X}_i, \quad \int_{\mathcal{X}_i} k_i(x_i, x'_i) d\mathbb{P}_{X_i}(x'_i) = 0$$

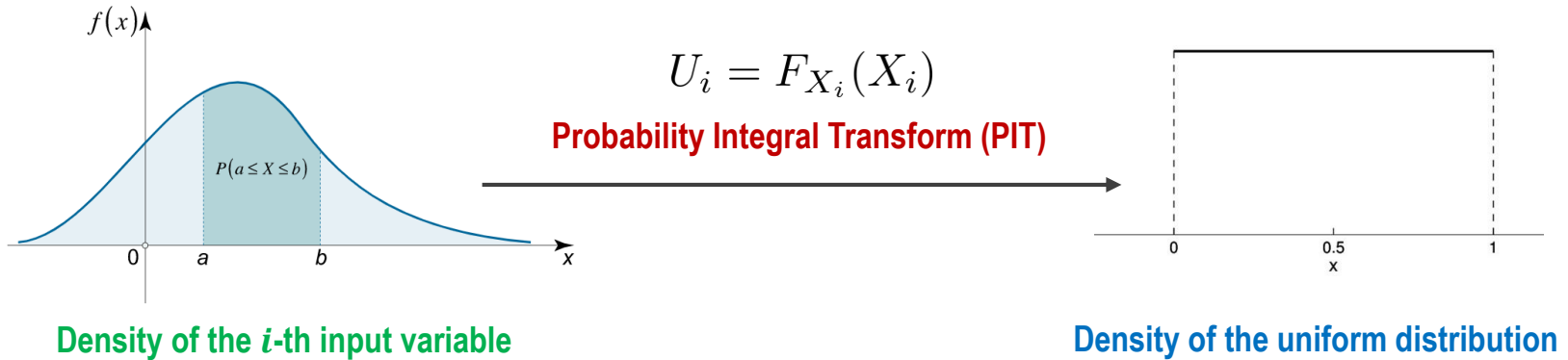
- ✓ $\mathcal{H}_i = \mathbb{R} \oplus \mathcal{G}_i$ where \mathcal{G}_i is only composed of **zero-mean functions** (with respect to \mathbb{P}_{X_i}).

How to find ANOVA kernels?

⚠ For most **parametric families of distributions**, there is no well-known **characteristic ANOVA** kernel.

🤔 How to implement the **HSIC-ANOVA decomposition** in practice?

1. Transform each input distribution \mathbb{P}_{X_i} into a standard uniform distribution $\mathcal{U}([0,1])$.



2. Assign a **Sobolev kernel** K_{Sob}^r to each new input variable $U_i := F_{X_i}(X_i)$.

$$\forall u, u' \in [0, 1], \quad K_{\text{Sob}}^r(u, u') := 1 + \sum_{i=1}^r \frac{B_i(u) B_i(u')}{(i!)^2} + \frac{(-1)^{r+1}}{(2r)!} B_{2r}(|u - u'|)$$

- ✓ $r \in \mathbb{N}^*$ is an **integer parameter** indicating the **degree of smoothness** of the RKHS.
- ✓ The functions $(B_i)_{i \geq 1}$ are the **Bernoulli polynomials** $\rightarrow \int_0^1 B_i(u) du = 0$.



A grey area around HSIC-ANOVA indices?

1. How do they measure sensitivity? How to distinguish between main effects and interactions?
2. Are they able to characterize independence?

GSA requirements	T_i	HSIC(X_i, Y)	S_i^{HSIC} ?	T_i^{HSIC}
ANOVA decomposition → RANKING	✓	✗	?	?
Characterize independence → SCREENING	✓	✓	?, ?	?, ?
Estimation from GIVEN DATA	✗	✓	✓	✓
Estimation from SMALL DATA	✗	✓	✓	✓
Compatibility with DEPENDENT inputs	✗	✓	✗	✗
INVARIANCE through monotonic transformations	✓	✗	✗	✗



4. Is it relevant to talk about interactions for HSIC-ANOVA indices?

Focus on HSIC-ANOVA interactions



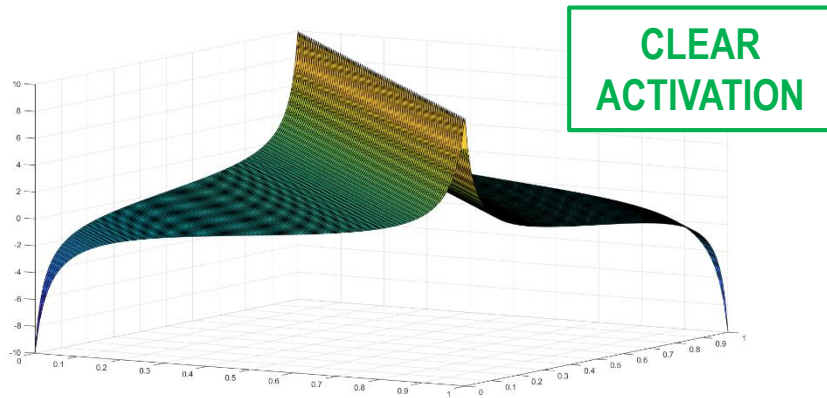
! For most benchmark test cases, HSIC-ANOVA interactions are not significant.

Example → the Ishigami function

$$Y = g(X_1, X_2, X_3) = \sin(X_1) + \sin^2(X_2) + X_3^4 \sin(X_1) \quad \text{with} \quad X_i \sim \mathcal{U}([- \pi, \pi])$$

- Strong interaction between X_1 and X_3 in the **variance-based ANOVA** framework.
- No interaction between X_1 and X_3 in the **HSIC-ANOVA** framework.

Counterexample → Hand-made **pathological functions** (only for $d \approx 2$)



Hull function

$$g(x_1, x_2) = -\tan \left[(2\sqrt{2})a \left| \frac{x_1 + x_2 - 1}{\sqrt{2}} \right| - a \right]$$

$$S_1^{\text{HSIC}} = S_2^{\text{HSIC}} = 17\%$$

$$T_1^{\text{HSIC}} = T_2^{\text{HSIC}} = 83\%$$

! No clear explanation on why those functions lead to strong HSIC-ANOVA interactions.

💡 The feature-based viewpoint on the HSIC allows to break the deadlock.

A detour through cross-covariance operators

1 HSIC indices → Gretton *et al.* (2005)

- Let $K_i : \mathcal{X}_i \times \mathcal{X}_i \rightarrow \mathbb{R}$ be the i -th input kernel (with RKHS denoted by \mathcal{H}_i).
- Let $K_Y : \mathcal{Y} \times \mathcal{Y} \rightarrow \mathbb{R}$ be the output kernel (with RKHS denoted by \mathcal{H}_Y).
- The knowledge of \mathcal{H}_i and \mathcal{H}_Y allows to rewrite $\text{HSIC}(X_i, Y)$ as a kind of **generalized covariance**.

$$\text{HSIC}(X_i, Y) = \sum_k \sum_l |\text{Cov}(v_{ik}(X_i), w_l(Y))|^2 \quad \text{with} \quad \begin{cases} (v_{ik})_k & \text{an ONB of } \mathcal{H}_i \\ (w_l)_l & \text{an ONB of } \mathcal{H}_Y \end{cases}$$

sum of covariances for different patterns catalogues of transformations

- ✓ **Aggregation of covariance terms** obtained after applying sequences of **preliminary basis transformations**.
- ✓ Each pair of non-linear functions $(v_{ik}(\cdot), w_l(\cdot))$ corresponds to a **non-linear dependence pattern**.

Example

➤ **HSIC indices** computed with **Gaussian kernels** → $K_i = K_Y = K_\gamma$

$$K_\gamma(z, z') = e^{-\frac{1}{2} \left(\frac{z-z'}{\gamma} \right)^2} = \sum_{k=0}^{\infty} g_k(z) g_k(z') \quad \text{with} \quad g_k(z) \propto e^{-\frac{1}{2} \left(\frac{z}{\gamma} \right)^2} z^k$$

damped polynomial feature

$$\text{HSIC}(X_i, Y) = \sum_{k=0}^{\infty} \sum_{l=0}^{\infty} \left| \text{Cov}(g_k(X_i), g_l(Y)) \right|^2$$

- ✓ **Infinitely many damped polynomial transformations are applied to both X_i and Y .**

A detour through cross-covariance operators



2 HSIC-ANOVA indices

$$S_1^{\text{HSIC}} + S_2^{\text{HSIC}} + \Delta_{12}^{\text{HSIC}} = 1$$

➤ For the sake of clarity, it is assumed that $d = 2$.

✓ **No loss of generality. Everything remains true in higher dimension!**

A detour through cross-covariance operators

2 HSIC-ANOVA indices

$$S_1^{\text{HSIC}} + S_2^{\text{HSIC}} + \Delta_{12}^{\text{HSIC}} = 1$$

$$S_1^{\text{HSIC}} \propto \sum_i \sum_k |\text{Cov}(u_{1i}(X_1), w_k(Y))|^2$$

dependence patterns captured by k_1 and K_Y

with

$$\begin{cases} (u_{1i})_i & \text{an ONB of } \mathcal{G}_1 \\ (w_k)_k & \text{an ONB of } \mathcal{H}_Y \end{cases}$$

$$\Delta_{12}^{\text{HSIC}} \propto \sum_i \sum_j \sum_k |\text{Cov}(u_{1i}(X_1) u_{2j}(X_2), w_k(Y))|^2$$

dependence patterns captured by $k_1 \otimes k_2$ and K_Y

with

$$\begin{cases} (u_{1i})_i & \text{an ONB of } \mathcal{G}_1 \\ (u_{2j})_j & \text{an ONB of } \mathcal{G}_2 \\ (w_k)_k & \text{an ONB of } \mathcal{H}_Y \end{cases}$$

A detour through cross-covariance operators

2 HSIC-ANOVA indices

$$S_1^{\text{HSIC}} + S_2^{\text{HSIC}} + \Delta_{12}^{\text{HSIC}} = 1$$

$$S_1^{\text{HSIC}} \propto \sum_i \sum_k |\text{Cov}(u_{1i}(X_1), w_k(Y))|^2$$

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$$\Delta_{12}^{\text{HSIC}} \propto \sum_i \sum_j \sum_k |\text{Cov}(u_{1i}(X_1) u_{2j}(X_2), w_k(Y))|^2$$

dependence patterns captured by $k_1 \otimes k_2$ and K_Y

with

$$\begin{cases} (u_{1i})_i & \text{an ONB of } \mathcal{G}_1 \\ (u_{2j})_j & \text{an ONB of } \mathcal{G}_2 \\ (w_k)_k & \text{an ONB of } \mathcal{H}_Y \end{cases}$$



➤ Remember the simplest solution to compute HSIC-ANOVA indices.

✓ Uniform inputs

→

$$U_1 \perp U_2 \sim \mathbf{u}([0,1])$$

✓ Sobolev kernels for the inputs

→

$$K_1 = K_2 = K_{\text{Sob}}^r$$

✓ Gaussian kernel for the output

→

$$K_Y = K_Y$$

$$\forall u, u' \in [0, 1], \quad K_{\text{Sob}}^r(u, u') := 1 + \sum_{k=1}^r \frac{B_k(u) B_k(u')}{(k!)^2} + \frac{(-1)^{r+1}}{(2r)!} B_{2r}(|u - u'|)$$



What are the basis functions in the case of Sobolev kernels?



5 ■ **More about Sobolev kernels and their properties**

Sobolev kernels and their feature maps



Many questions at the beginning of this work...

- 1 What is the RKHS $\mathcal{H}_{\text{Sob}}^r$ induced by K_{Sob}^r ?
- 2 Is K_{Sob}^r a characteristic kernel?
- 3 Is there an explicit and easily interpretable feature map $\varphi_{\text{Sob}}^r : [0,1] \rightarrow \mathcal{F}_{\text{Sob}}^r$?
- 4 How to identify an ONB of $\mathcal{H}_{\text{Sob}}^r$? Is there a link with feature maps?
- 5 How to choose r in practice?

Sobolev kernels and their feature maps



Many questions at the beginning of this work...

1 What is the RKHS $\mathcal{H}_{\text{Sob}}^r$ induced by K_{Sob}^r ? → see Gu (2013) or Kuo *et al.* (2010)

➤ A standard function space: the Sobolev space of order r defined on $[0,1]$ for the L^2 -norm.

$$H^r([0, 1]) := \left\{ h \in \mathbb{R}^{[0,1]} \mid \forall 0 \leq k \leq r, D^k h \in L^2([0, 1]) \right\}$$

➤ A specific inner product:

$$\langle f, g \rangle_{\mathcal{H}_{\text{Sob}}^r} := \sum_{k=0}^{r-1} \left(\int_0^1 D^k f(x) dx \right) \left(\int_0^1 D^k g(x) dx \right) + \int_0^1 D^r f(x) D^r g(x) dx$$

Sobolev kernels and their feature maps



Many questions at the beginning of this work...

1 What is the RKHS $\mathcal{H}_{\text{Sob}}^r$ induced by K_{Sob}^r ?

2 Is K_{Sob}^r a characteristic kernel?

➤ **YES!** Simply because $H^r([0,1])$ is **uniformly dense** in $C([0,1])$.

➤ **Major consequence**

✓ The **HSIC-ANOVA indices** based on **Sobolev kernels** are able to characterize independence.

$$X_i \perp Y \iff S_i^{\text{HSIC}} = 0 \iff T_i^{\text{HSIC}} = 0$$



This is different from what happens for **Sobol' indices**.

$$S_i = 0 \not\Rightarrow X_i \perp Y \quad \text{while} \quad X_i \perp Y \iff T_i = 0$$

Sobolev kernels and their feature maps



Many questions at the beginning of this work...

- 1 What is the RKHS $\mathcal{H}_{\text{Sob}}^r$ induced by K_{Sob}^r ?
- 2 Is K_{Sob}^r a characteristic kernel?
- 3 Is there an explicit and easily interpretable feature map $\varphi_{\text{Sob}}^r : [0,1] \rightarrow \mathcal{F}_{\text{Sob}}^r$?

$$K_{\text{Sob}}^r(x, x') = \langle \varphi_{\text{Sob}}^r(x), \varphi_{\text{Sob}}^r(x') \rangle_{\mathcal{F}_{\text{Sob}}^r}$$

➤ For $r = 1$, the **Mercer expansion** of K_{Sob}^1 is actually known. → [Dick et al. \(2014, 2015\)](#)

$$K_{\text{Sob}}^1(x, x') := 1 + \sum_{k=1}^{\infty} \frac{1}{(k\pi)^2} c_k(x) c_k(x') \quad \text{with} \quad c_k(x) := \sqrt{2} \cos(k\pi x)$$

➤ For $r \geq 2$, a **series expansion** of K_{Sob}^r is also mentioned in the literature. → [Baldeaux et al. \(2009\)](#)

$$K_{\text{Sob}}^r(x, x') := 1 + \sum_{k=1}^r \frac{B_k(x) B_k(x')}{(k!)^2} + \sum_{k=1}^{\infty} \frac{1}{(2k\pi)^{2r}} \left[c_{2k}(x) c_{2k}(x') + s_{2k}(x) s_{2k}(x') \right] \quad \text{with} \quad \begin{cases} c_{2k}(x) := \sqrt{2} \cos(2k\pi x) \\ s_{2k}(x) := \sqrt{2} \sin(2k\pi x) \end{cases}$$

Sobolev kernels and their feature maps



Many questions at the beginning of this work...

- 1 What is the RKHS $\mathcal{H}_{\text{Sob}}^r$ induced by K_{Sob}^r ?
- 2 Is K_{Sob}^r a characteristic kernel?
- 3 Is there an explicit and easily interpretable feature map $\varphi_{\text{Sob}}^r : [0,1] \rightarrow \mathcal{F}_{\text{Sob}}^r$?
- 4 How to identify an ONB of $\mathcal{H}_{\text{Sob}}^r$? Is there a link with feature maps?

➤ Mercer expansion of K_{Sob}^1

→

$$K_{\text{Sob}}^1(x, x') := 1 + \sum_{k=1}^{\infty} \frac{1}{(k\pi)^2} c_k(x) c_k(x')$$

➤ ONB of the RKHS $\mathcal{H}_{\text{Sob}}^1$

→

$$\left\{ \mathbf{1}; \left(\frac{c_k(\cdot)}{k\pi} \right)_{k \geq 1} \right\}$$

➤ Series expansion of K_{Sob}^r

→

$$K_{\text{Sob}}^r(x, x') := 1 + \sum_{k=1}^r \frac{B_k(x) B_k(x')}{(k!)^2} + \sum_{k=1}^{\infty} \frac{1}{(2k\pi)^{2r}} \left[c_{2k}(x) c_{2k}(x') + s_{2k}(x) s_{2k}(x') \right]$$

➤ ONB of the RKHS $\mathcal{H}_{\text{Sob}}^r$

→

$$\left\{ \mathbf{1}; \left(\frac{B_k(\cdot)}{k!} \right)_{1 \leq k \leq r}; \left(\frac{c_{2k}(\cdot)}{(2k\pi)^r} \right)_{k \geq 1}; \left(\frac{s_{2k}(\cdot)}{(2k\pi)^r} \right)_{k \geq 1} \right\}$$

Sobolev kernels and their feature maps



Many questions at the beginning of this work...

- 1 What is the RKHS $\mathcal{H}_{\text{Sob}}^r$ induced by K_{Sob}^r ?
 - 2 Is K_{Sob}^r a characteristic kernel?
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 - 4 How to identify an ONB of $\mathcal{H}_{\text{Sob}}^r$? Is there a link with feature maps?
 - 5 How to choose r in practice?
- Taking $r = 1$ is recommended!
- For $r \geq 2$, $K_{\text{Sob}}^r(x, x') \approx 1 + k_{\text{lin}}(x, x')$ → **poor numerical performance** for screening!

Sobolev kernels and their feature maps



Many questions at the beginning of this work...

- 1 What is the RKHS $\mathcal{H}_{\text{Sob}}^r$ induced by K_{Sob}^r ?
- 2 Is K_{Sob}^r a characteristic kernel?
- 3 Is there an explicit and easily interpretable feature map $\varphi_{\text{Sob}}^r : [0,1] \rightarrow \mathcal{F}_{\text{Sob}}^r$?
- 4 How to identify an ONB of $\mathcal{H}_{\text{Sob}}^r$? Is there a link with feature maps?
- 5 How to choose r in practice?



What is the point of these theoretical results?

- Remember the pure interaction term $\Delta_{12}^{\text{HSIC}}$.
- Apply with $K_1 = K_2 = K_{\text{Sob}}^1$ now that an ONB of $\mathcal{H}_{\text{Sob}}^1$ is explicitly known.

$$\Delta_{12}^{\text{HSIC}} \propto \sum_i \sum_j \sum_k |\text{Cov}(u_{1i}(X_1) u_{2j}(X_2), w_k(Y))|^2 = \boxed{\sum_i \sum_j \sum_k \frac{1}{ij \pi^2} |\text{Cov}(c_i(X_1) c_j(X_2), w_k(Y))|^2}$$



This provides the hint to design a toy case.

How to exacerbate HSIC-ANOVA interactions?



➤ Back to the Ishigami function

- ✓ Additional term chosen to boost HSIC-ANOVA interactions.

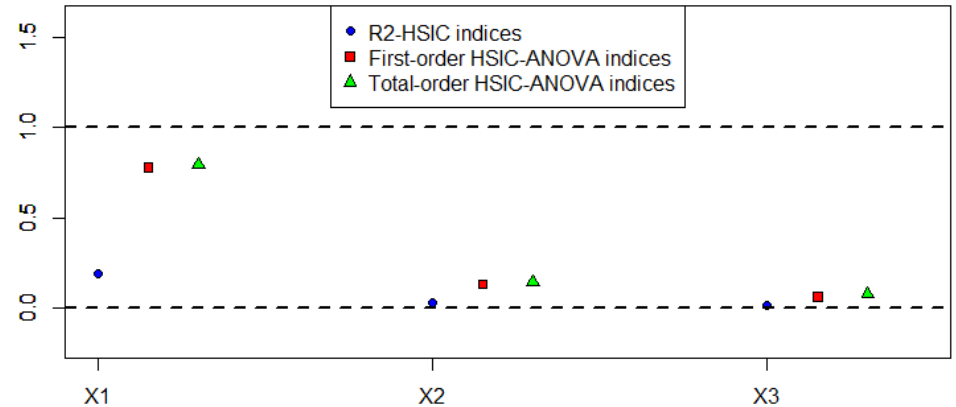
$$Y = g(U_1, U_2, U_3) = \text{ishigami}(X_1, X_2, X_3) + \gamma \cos(\pi U_1) \cos(\pi U_2) \quad \text{with} \quad \begin{aligned} U_i &\sim \mathcal{U}([0,1]) \\ X_i &= \pi(2U_i - 1) \end{aligned}$$

➤ Design parameter

- ✓ $\gamma = 0$

➤ Estimation of sensitivity measures

- ✓ Sample size $n = 500$
- ✓ R²-HSIC indices + HSIC-ANOVA indices



	U_1	U_2	U_3
R ² -HSIC	0.19	0.03	0.01
First-order	0.77	0.13	0.07
Total-order	0.79	0.14	0.08

How to exacerbate HSIC-ANOVA interactions?



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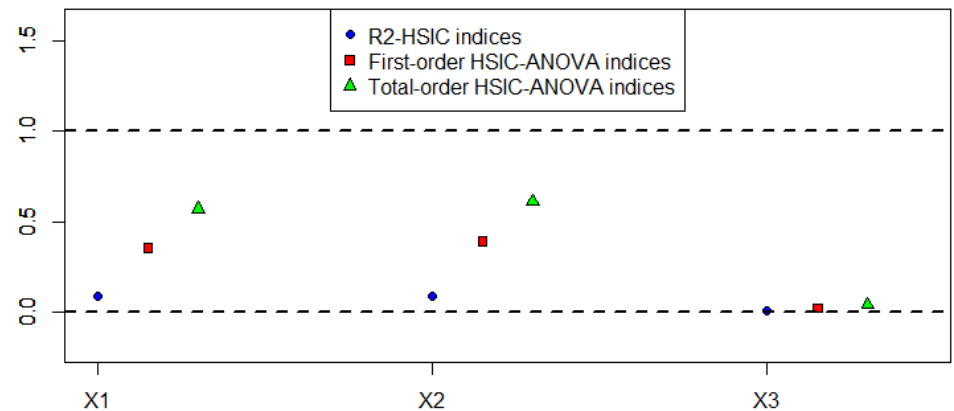
$$Y = g(U_1, U_2, U_3) = \text{ishigami}(X_1, X_2, X_3) + \gamma \cos(\pi U_1) \cos(\pi U_2) \quad \text{with} \quad \begin{aligned} U_i &\sim \mathcal{U}([0,1]) \\ X_i &= \pi(2U_i - 1) \end{aligned}$$

➤ Design parameter

- ✓ $\gamma = 10$

➤ Estimation of sensitivity measures

- ✓ Sample size $n = 500$
- ✓ R²-HSIC indices + HSIC-ANOVA indices



	U_1	U_2	U_3
R ² -HSIC	0.05	0.08	0.01
First-order	0.25	0.40	0.02
Total-order	0.56	0.71	0.04

How to exacerbate HSIC-ANOVA interactions?



➤ Back to the Ishigami function

- ✓ Additional term chosen to boost HSIC-ANOVA interactions.

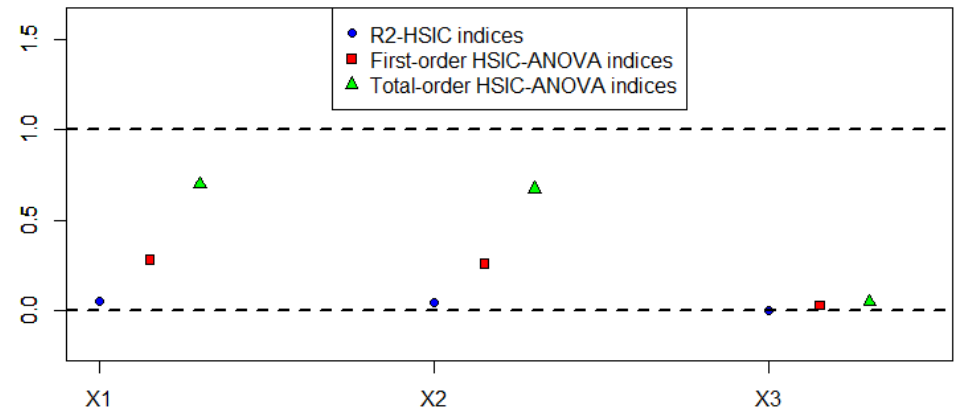
$$Y = g(U_1, U_2, U_3) = \text{ishigami}(X_1, X_2, X_3) + \gamma \cos(\pi U_1) \cos(\pi U_2) \quad \text{with} \quad \begin{aligned} U_i &\sim \mathcal{U}([0,1]) \\ X_i &= \pi(2U_i - 1) \end{aligned}$$

➤ Design parameter

- ✓ $\gamma = 100$

➤ Estimation of sensitivity measures

- ✓ Sample size $n = 500$
- ✓ R²-HSIC indices + HSIC-ANOVA indices



	U_1	U_2	U_3
R ² -HSIC	0.05	0.05	0.01
First-order	0.28	0.23	0.04
Total-order	0.72	0.66	0.05



How to use HSIC-ANOVA in practice?

1. How to build a test of independence? How to extend to the existing test procedures?
2. Is there any advantage to using the total-order HSIC-ANOVA index?

GSA requirements	T_i	HSIC(X_i, Y)	S_i^{HSIC}	T_i^{HSIC}
ANOVA decomposition → RANKING	✓	✗	✓	✓
Characterize independence → SCREENING	✓	✓	✓ (⊕ ?)	✓ (⊕ ?)
Estimation from GIVEN DATA	✗	✓	✓	✓
Estimation from SMALL DATA	✗	✓	✓	✓
Compatibility with DEPENDENT inputs	✗	✓	✗	✗
INVARIANCE through monotonic transformations	✓	✗	✗	✗



6 ■ Does all this benefit independence testing?

Testing independence with HSIC-ANOVA indices



➤ A test of independence consists in testing the null hypothesis $(H_0^i) : X_i \perp Y$.

$$X_i \perp Y \iff S_i^{\text{HSIC}} = 0 \iff T_i^{\text{HSC}} = 0$$

$$\iff \boxed{\begin{array}{l} \text{HSIC}(X_i, Y) = 0 \\ \text{with } K_{\text{Sob}}^1 \otimes K_Y \end{array}} \iff \boxed{\begin{array}{l} \text{HSIC}(\mathbf{X}, Y) - \text{HSIC}(\mathbf{X}_{-i}, Y) = 0 \\ \text{with } K_{\text{Sob}}^1 \otimes \dots \otimes K_{\text{Sob}}^1 \otimes K_Y \end{array}}$$

Numerator of the first-order index

Numerator of the total-order index

Testing independence with HSIC-ANOVA indices

➤ A test of independence consists in testing the null hypothesis $(H_0^i) : X_i \perp Y$.

$$X_i \perp Y \iff S_i^{\text{HSIC}} = 0 \iff T_i^{\text{HSC}} = 0$$

$$\iff \begin{array}{c} \text{HSIC}(X_i, Y) = 0 \\ \text{with } K_{\text{Sob}}^1 \otimes K_Y \end{array} \iff \begin{array}{c} \text{HSIC}(\mathbf{X}, Y) - \text{HSIC}(\mathbf{X}_{-i}, Y) = 0 \\ \text{with } K_{\text{Sob}}^1 \otimes \dots \otimes K_{\text{Sob}}^1 \otimes K_Y \end{array}$$

Numerator of the first-order index

Numerator of the total-order index



V-statistic estimator

$$\hat{S}_i(\mathbf{Z}_{\text{obs}}) = \widehat{\text{HSIC}}_v(X_i, Y)$$

Apply existing test procedures with $K_i = K_{\text{Sob}}^1$

🤔 Is there a reason to hope for higher statistical power?

Actually, NO!

Testing independence with HSIC-ANOVA indices

➤ A test of independence consists in testing the null hypothesis $(H_0^i) : X_i \perp Y$.

$$X_i \perp Y \iff S_i^{\text{HSIC}} = 0 \iff T_i^{\text{HSC}} = 0$$

$$\iff \begin{array}{c} \text{HSIC}(X_i, Y) = 0 \\ \text{with } K_{\text{Sob}}^1 \otimes K_Y \end{array} \iff \begin{array}{c} \text{HSIC}(\mathbf{X}, Y) - \text{HSIC}(\mathbf{X}_{-i}, Y) = 0 \\ \text{with } K_{\text{Sob}}^1 \otimes \dots \otimes K_{\text{Sob}}^1 \otimes K_Y \end{array}$$

Numerator of the first-order index

Numerator of the total-order index

V-statistic estimator

$$\hat{S}_i(\mathbf{Z}_{\text{obs}}) = \widehat{\text{HSIC}}_v(X_i, Y)$$



Much more appealing!

Apply existing test procedures with $K_i = K_{\text{Sob}}^1$

🤔 Is there a reason to hope for higher statistical power?

Actually, NO!

V-statistic estimator

$$\hat{T}_i = \widehat{\text{HSIC}}_v(\mathbf{X}, Y) - \widehat{\text{HSIC}}_v(\mathbf{X}_{-i}, Y)$$

Computing this test statistic is slightly more expensive.

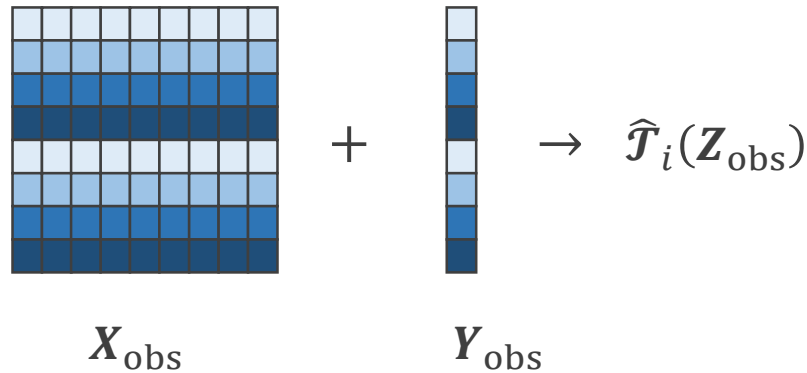
🤔 Is there a reason to hope for higher statistical power?

Let us see!

Testing independence with the total-order index

➤ The distribution of $\hat{\mathcal{J}}_i(\mathbf{Z}_{\text{obs}})$ under (H_0^i) can be simulated from the available data.

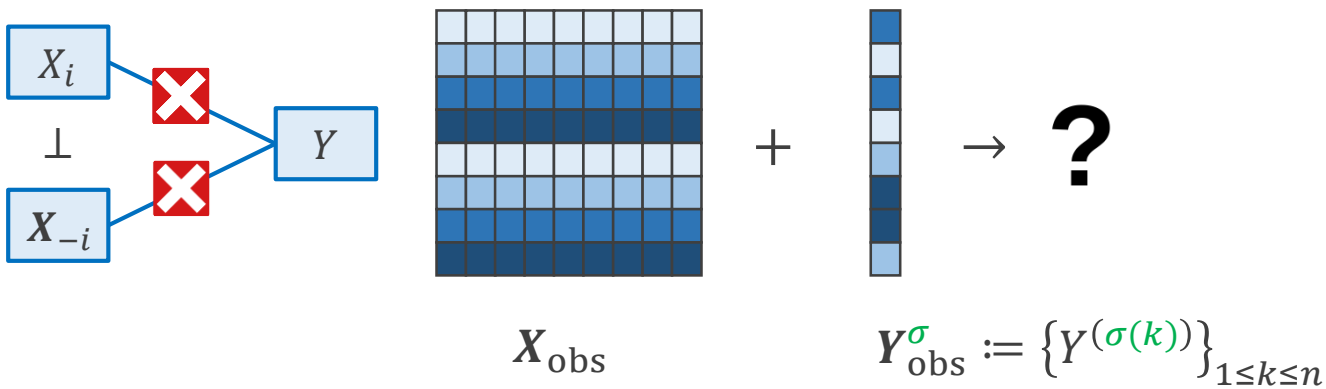
 All the columns of the DoE are required to compute the test statistic.



Testing independence with the total-order index

- The distribution of $\hat{\mathcal{T}}_i(\mathbf{Z}_{\text{obs}})$ under (H_0^i) can be simulated from the available data.

 All the columns of the DoE are required to compute the test statistic.

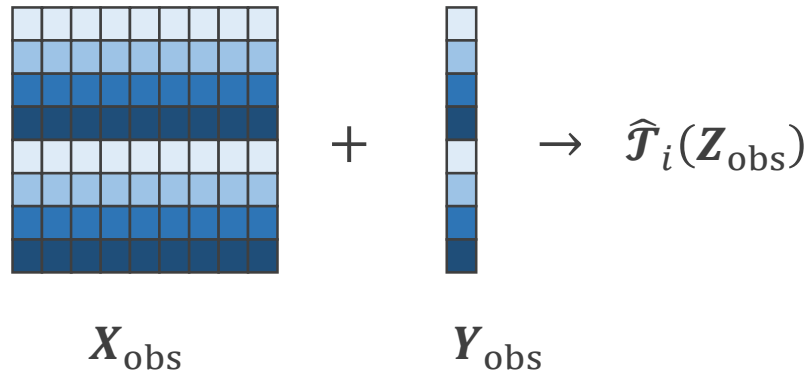


- Permuting \mathbf{Y}_{obs} leads to **eliminate dependence** between the joint observations $(\mathbf{X}^{(k)}, Y^{(k)})$.
 - ✓ This boils down to testing $(H_0) : \mathbf{X} \perp Y$ and this is not what is desired!

Testing independence with the total-order index

➤ The distribution of $\hat{\mathcal{J}}_i(\mathbf{Z}_{\text{obs}})$ under (H_0^i) can be simulated from the available data.

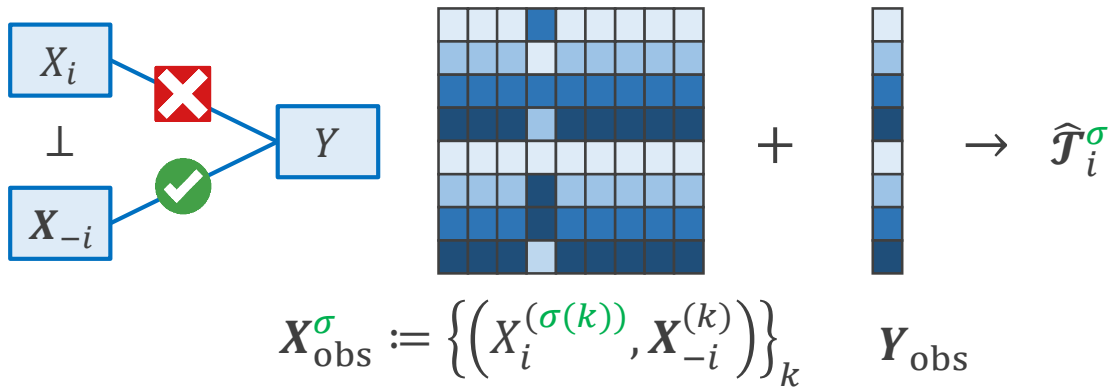
💡 Instead, the trick is to permute the observations of the input variable.



Testing independence with the total-order index

➤ The distribution of $\hat{\mathcal{J}}_i(\mathbf{Z}_{\text{obs}})$ under (H_0^i) can be simulated from the available data.

💡 Instead, the trick is to permute the observations of the input variable.



Testing independence with the total-order index

➤ The distribution of $\hat{\mathcal{T}}_i(\mathbf{Z}_{\text{obs}})$ under (H_0^i) can be simulated from the available data.

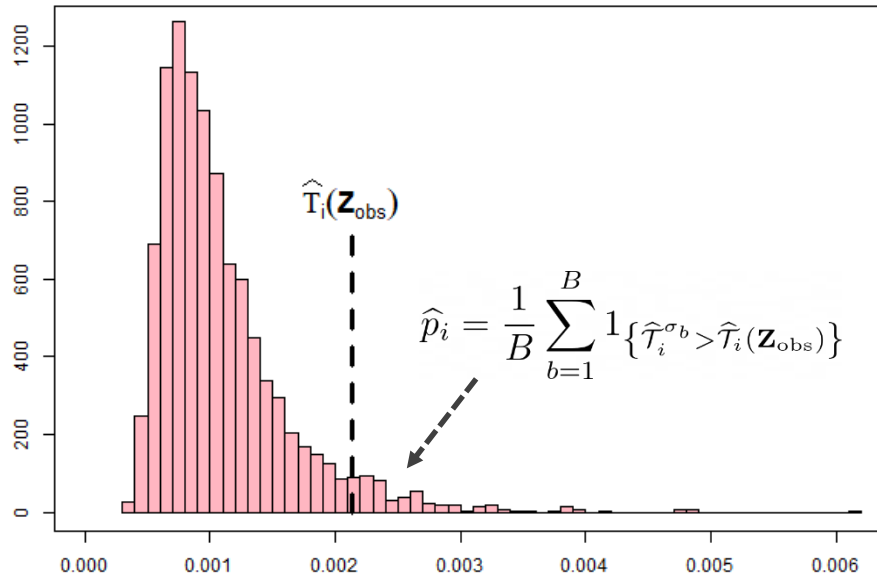


Instead, the trick is to permute the observations of the input variable.

Permutation-based test procedure

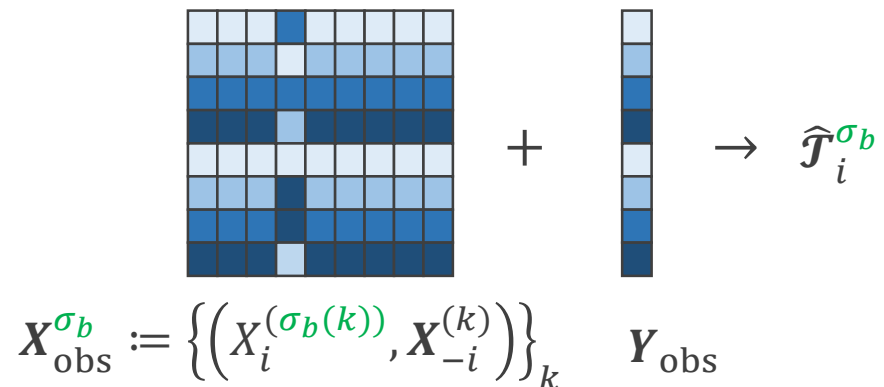
- **Step A** → Perform a sequence $\{\sigma_b\}_{1 \leq b \leq B}$ of **random permutations** on the i -th column of \mathbf{X}_{obs} .
- **Step B** → Compute the value $\hat{\mathcal{T}}_i^{\sigma_b}$ of the test statistic for each permuted design.
- **Step C** → Derive a non-parametric estimate of the p-value $p_i := \mathbb{P}(\hat{\mathcal{T}}_i > \hat{\mathcal{T}}_i(\mathbf{Z}_{\text{obs}}))$.

Simulation of the test statistic under the null hypothesis



- **Default value:** $B \approx 10^3$
- **Complexity:** $(d^2 + 7Bd) n^2$

Permutation scheme



Numerical study of the statistical power



➤ Back to the Ishigami function

- ✓ Additional term chosen to boost HSIC-ANOVA interactions.

$$Y = g(U_1, U_2, U_3) = \text{ishigami}(X_1, X_2, X_3) + \gamma \cos(\pi U_1) \cos(\pi U_2) \quad \text{with} \quad \begin{aligned} U_i &\sim \mathcal{U}([0,1]) \\ X_i &= \pi(2U_i - 1) \end{aligned}$$

➤ Design parameter

- ✓ $\gamma = 0$

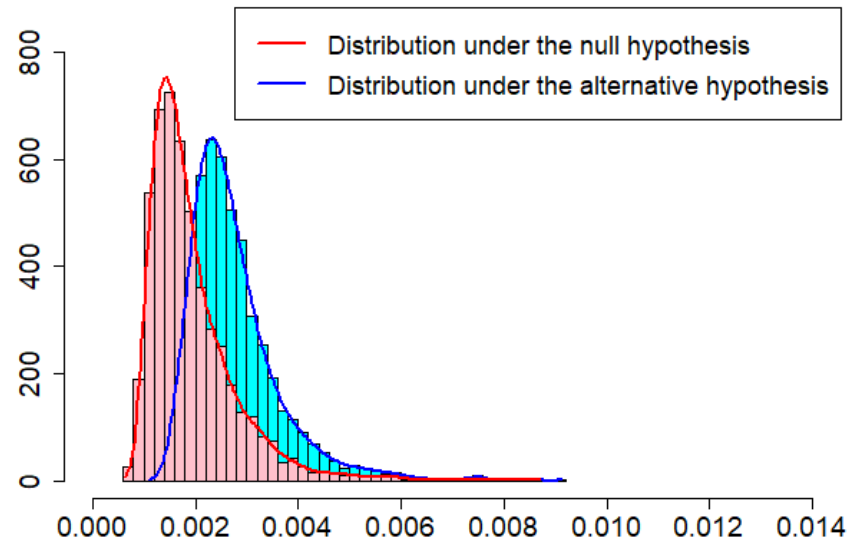
➤ Study of the statistical power

- ✓ Sample size $n = 50$
- ✓ Number of replicates $M = 200$

	U_1	U_2	U_3
HSIC	0.88	0.07	0.22
Total-order	0.87	0.19	0.19

➤ Separation rate

- ✓ Distributions of $\hat{\mathcal{T}}_i(\mathbf{Z}_{\text{obs}})$ under (H_0^i) et (H_1^i)



Numerical study of the statistical power



➤ Back to the Ishigami function

✓ Additional term chosen to boost HSIC-ANOVA interactions.

$$Y = g(U_1, U_2, U_3) = \text{ishigami}(X_1, X_2, X_3) + \gamma \cos(\pi U_1) \cos(\pi U_2) \quad \text{with} \quad \begin{aligned} U_i &\sim \mathcal{U}([0,1]) \\ X_i &= \pi(2U_i - 1) \end{aligned}$$

➤ Design parameter

✓ $\gamma = 10$

➤ Study of the statistical power

✓ Sample size $n = 50$

✓ Number of replicates $M = 200$

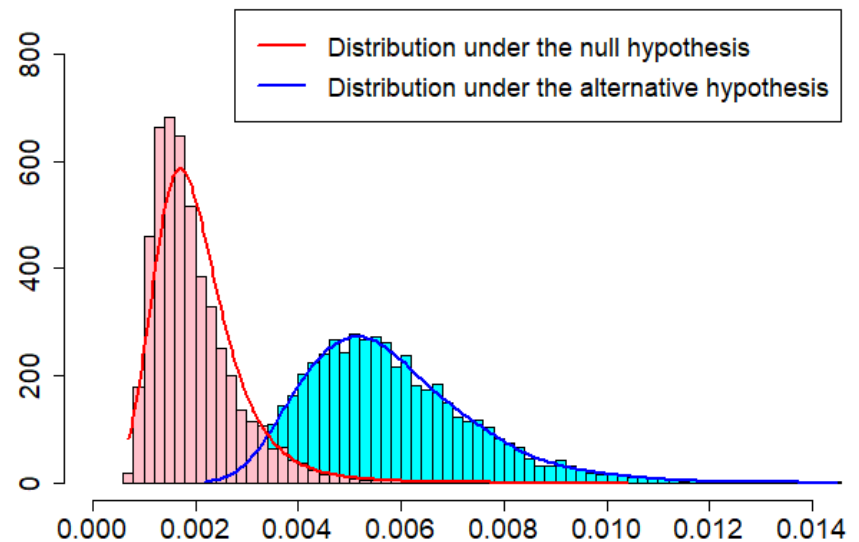
	U_1	U_2	U_3
HSIC	0.59	0.63	0.05
Total-order	0.92	0.94	0.07

✓ Increased power when $S_i^{\text{HSIC}} \ll T_i^{\text{HSIC}}$

✓ Same power when $S_i^{\text{HSIC}} \approx T_i^{\text{HSIC}}$

➤ Separation rate

✓ Distributions of $\hat{T}_i(\mathbf{Z}_{\text{obs}})$ under (H_0^i) et (H_1^i)



Numerical study of the statistical power



➤ Back to the Ishigami function

✓ Additional term chosen to boost HSIC-ANOVA interactions.

$$Y = g(U_1, U_2, U_3) = \text{ishigami}(X_1, X_2, X_3) + \gamma \cos(\pi U_1) \cos(\pi U_2) \quad \text{with} \quad \begin{aligned} U_i &\sim \mathcal{U}([0,1]) \\ X_i &= \pi(2U_i - 1) \end{aligned}$$

➤ Design parameter

✓ $\gamma = 100$

➤ Study of the statistical power

✓ Sample size $n = 50$

✓ Number of replicates $M = 200$

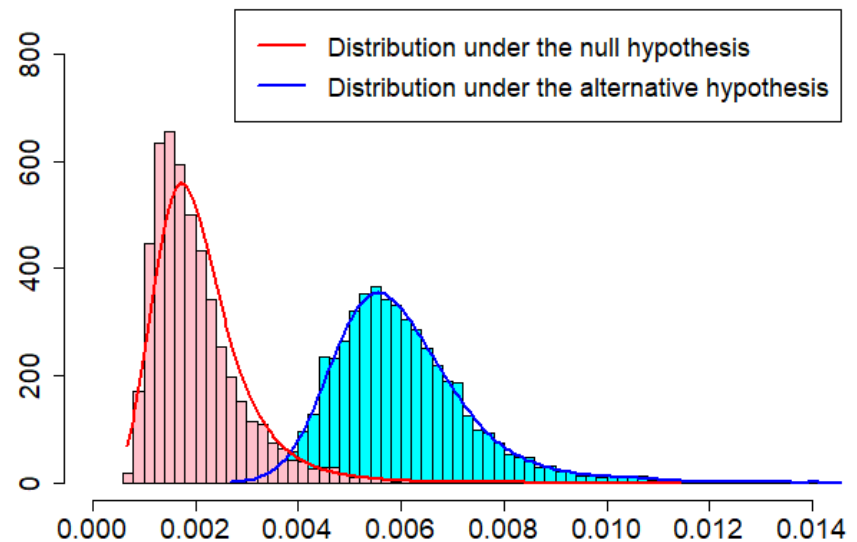
	U_1	U_2	U_3
HSIC	0.65	0.70	0.07
Total-order	1.00	1.00	0.06

✓ Increased power when $S_i^{\text{HSIC}} \ll T_i^{\text{HSIC}}$

✓ Same power when $S_i^{\text{HSIC}} \approx T_i^{\text{HSIC}}$

➤ Separation rate

✓ Distributions of $\hat{T}_i(\mathbf{Z}_{\text{obs}})$ under (H_0^i) et (H_1^i)





Benefits brought by HSIC-ANOVA indices in GSA

- ✓ HSIC-ANOVA indices are fully transparent sensitivity measures able to perform screening and ranking!
- ✓ In many situations, the test of independence based on T_i^{HSIC} is more powerful!

GSA requirements	T_i	HSIC(X_i, Y)	S_i^{HSIC}	T_i^{HSIC}
ANOVA decomposition → RANKING	✓	✗	✓	✓
Characterize independence → SCREENING	✓	✓	✓	✓ ✓
Estimation from GIVEN DATA	✗	✓	✓	✓
Estimation from SMALL DATA	✗	✓	✓	✓
Compatibility with DEPENDENT inputs	✗	✓	✗	✗
INVARIANCE through monotonic transformations	✓	✗	✗	✗



7 ■ Conclusion

Conclusion

- The very recent **HSIC-ANOVA indices** have enabled **significant progress in GSA** since they combine the advantages of **Sobol' indices (variance-based GSA)** and those of **HSIC indices (kernel-based GSA)**.
- The **HSIC-ANOVA decomposition** requires the use of **characteristic ANOVA kernels** for the input variables.
- The way sensitivity is measured by HSIC-ANOVA indices is driven by the kernel feature maps.
- **Variable selection** can be performed with **test procedures** based on HSIC-ANOVA indices.
- Using the total-order HSIC-ANOVA indices leads to **more powerful** test procedures.

Traditional benchmarks

- ✓ Ishigami, Friedman, Morris...

$$S_i^{HSIC} \lesssim T_i^{HSIC}$$

$$\text{Power}(\widehat{\mathcal{S}}_i) \approx \text{Power}(\widehat{\text{HSIC}}_{\mathcal{N}}) \approx \text{Power}(\widehat{\mathcal{T}}_i)$$

Specific benchmarks

- ✓ Hand-made use cases.
- ✓ Test functions in optimization.
- ✓ Flexible metafunction framework.

$$S_i^{HSIC} \ll T_i^{HSIC}$$

$$\text{Power}(\widehat{\mathcal{S}}_i) \ll \text{Power}(\widehat{\text{HSIC}}_{\mathcal{N}}) \ll \text{Power}(\widehat{\mathcal{T}}_i)$$

Conclusion

- The very recent **HSIC-ANOVA indices** have enabled **significant progress in GSA** since they combine the advantages of **Sobol' indices** (variance-based GSA) and those of **HSIC indices** (kernel-based GSA).
- The **HSIC-ANOVA decomposition** requires the use of **characteristic ANOVA kernels** for the input variables.
- The way sensitivity is measured by HSIC-ANOVA indices is driven by the kernel feature maps.
- **Variable selection** can be performed with **test procedures** based on HSIC-ANOVA indices.
- Using the total-order HSIC-ANOVA indices leads to **more powerful** test procedures.

Publications

- **Preprint** → <https://cea.hal.science/cea-04320711/document>
- **Conference paper** → <https://cea.hal.science/cea-03701170v1/document>

Codes

- **Two dedicated routines the R package sensitivity**
 - ✓ **sensiHSIC** → <https://rdr.io/cran/sensitivity/man/sensiHSIC.html>
 - ✓ **testHSIC** → <https://rdr.io/cran/sensitivity/man/testHSIC.html>



8 ■ References

Key papers (1/3)



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