A Spatial Analysis of Multivariate Output from Regional Climate Models

NZZ.ch

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Outline

- Examples of spatial climate data
- Uni-/Multivariate spatial models
- General approaches
- Markovian modeling
- Example: Regional temperature & precipitation change

Outlook



Spatial data

General circulation model data



-30-24-18-12 -6 0 6 12 18 24 30



Source: www.cisl.ucar.edu

CCSM3 DJF tempemperature change

2080-2100 vs 1980-2000







Spatial data

Regional climate model data





Spatial data

Typical features

- Large datasets
- Complex nonstationarities
- Unknown dependencies
- Difficulty visualizing the results



Spatial, additive mixed effects model:

data = signal + noise

= fixed effects + trend + spatial term + noise



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```
data = signal + noise
= fixed effects + trend + spatial term + noise
```

or

$$Y(\mathbf{s}) = \mathbf{X}\beta + \alpha(\mathbf{s}) + \gamma(\mathbf{s}) + \varepsilon(\mathbf{s}) \qquad \mathbf{s} \in \mathcal{D} \subset \mathbb{R}^d, \ d \ge 1$$

with

 $Y(\mathbf{s})$: observations

- $X\beta$: fixed effects and trend
- $\alpha(\mathbf{s})$: spline component (trend)
- $\gamma(\mathbf{s})$: zero mean spatial Gaussian process
- $\varepsilon(\mathbf{s})$: iid Gaussian noise, orthogonal to $\gamma(\mathbf{s})$



Parameters $\boldsymbol{\theta} = \left(\boldsymbol{\beta}^{\mathsf{T}}, \boldsymbol{\theta}_{\alpha}{}^{\mathsf{T}}, \lambda_{\alpha}, \boldsymbol{\theta}_{\gamma}{}^{\mathsf{T}}, \sigma^{2}\right)^{\mathsf{T}}$:

- **X** β : coefficients β
- $\alpha(\mathbf{s})$: basis function coefficients $\boldsymbol{\theta}_{\alpha}$; smoothing parameter λ_{α}
- $\gamma(\mathbf{s})$: parameters $\boldsymbol{ heta}_{\gamma}$ decribing the covariance function
- $\varepsilon(\mathbf{s})$: variance σ^2



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Statistical tasks:

- estimation of θ
- smoothing or prediction
- uncertainty assessment
- model validation



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Multivariate modeling: setting

Spatial, additive mixed effects model:

$$Y_{1}(\mathbf{s}) = \mathbf{X}_{1}\beta_{1} + \alpha_{1}(\mathbf{s}) + \gamma_{1}(\mathbf{s}) + \varepsilon_{1}(\mathbf{s})$$

$$\vdots$$

$$Y_{p}(\mathbf{s}) = \mathbf{X}_{p}\beta_{p} + \alpha_{p}(\mathbf{s}) + \gamma_{p}(\mathbf{s}) + \varepsilon_{p}(\mathbf{s}) \qquad \mathbf{s} \in \mathcal{D} \subset \mathbb{R}^{d}, \ d \ge 1$$

with

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Multivariate modeling: setting

Modeling the spatial processes themselves:





Inlet: dependency modeling

Common process(es):

$$\begin{split} & \mathbf{X} \sim \mathcal{N}_{n}(\boldsymbol{\mu}_{\mathbf{X}}, \boldsymbol{\Sigma}_{\mathbf{X}}) \\ & \mathbf{Y} \sim \mathcal{N}_{n}(\boldsymbol{\mu}_{\mathbf{Y}}, \boldsymbol{\Sigma}_{\mathbf{Y}}) \end{split} \qquad \mathbf{Z} \sim \mathcal{N}_{n}(\mathbf{0}, \boldsymbol{\Sigma}_{\mathbf{Z}}) \\ & \sim \qquad \begin{pmatrix} \mathbf{X} + \mathbf{Z} \\ \mathbf{Y} + \mathbf{Z} \end{pmatrix} \sim \mathcal{N}_{2n} \left(\begin{pmatrix} \boldsymbol{\mu}_{\mathbf{X}} \\ \boldsymbol{\mu}_{\mathbf{Y}} \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma}_{\mathbf{X}} + \boldsymbol{\Sigma}_{\mathbf{Z}} & \boldsymbol{\Sigma}_{\mathbf{Z}} \\ \boldsymbol{\Sigma}_{\mathbf{Z}} & \boldsymbol{\Sigma}_{\mathbf{Y}} + \boldsymbol{\Sigma}_{\mathbf{Z}} \end{pmatrix} \right) \end{split}$$



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Cross-correlation model:

$$\begin{aligned} \mathbf{X} &\sim \mathcal{N}_{n}(\boldsymbol{\mu}_{\mathbf{X}}, \boldsymbol{\Sigma}_{\mathbf{X}}) \\ \mathbf{Y} &\sim \mathcal{N}_{n}(\boldsymbol{\mu}_{\mathbf{Y}}, \boldsymbol{\Sigma}_{\mathbf{Y}}) \end{aligned} \quad \operatorname{Cov}(\mathbf{X}, \mathbf{Y}) = \boldsymbol{\Sigma}_{\mathbf{X}\mathbf{Y}} \\ &\sim & \begin{pmatrix} \mathbf{X} \\ \mathbf{Y} \end{pmatrix} \sim \mathcal{N}_{2n} \left(\begin{pmatrix} \boldsymbol{\mu}_{\mathbf{X}} \\ \boldsymbol{\mu}_{\mathbf{Y}} \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma}_{\mathbf{X}} & \boldsymbol{\Sigma}_{\mathbf{X}\mathbf{Y}} \\ \boldsymbol{\Sigma}_{\mathbf{X}\mathbf{Y}}^{\mathsf{T}} & \boldsymbol{\Sigma}_{\mathbf{Y}} \end{pmatrix} \right) \end{aligned}$$



Inlet: random field modeling

Spatial process (GRF):





Covariance matrix: Σ



Inlet: random field modeling

Spatial process (GRF):





Covariance matrix: Σ



Lattice data (GMRF):



 $E[y_i|y_{-i}] = \beta \sum_{\substack{j \text{ neighbor of } i}} y_j$ $Var[y_i|y_{-i}] = \tau^2$

Gaussianity and regularity conditions:

$$\Sigma = \tau^2 (\mathbf{I} - \mathbf{B})^{-1}$$

Inlet: random field modeling



Spatial process:



- Iterative approaches
 - + Flexible, numerically feasible
 - Uncertainties
- Maximum likelihood
 - + Uncertainties, asymptotics
 - Numerical issues
- Bayesian hierarchical models
 - + Flexible, uncertainties
 - MCMC



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Backfitting: $1 2 3 4 \blacktriangleleft$

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Tapering: (1) (3); (2) (4) <

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Multivariate modeling: backfitting

Recall:

$$Y_{1}(\mathbf{s}) = \mathbf{X}_{1}\beta_{1} + \alpha_{1}(\mathbf{s}) + \gamma_{1}(\mathbf{s}) + \varepsilon_{1}(\mathbf{s})$$

:
$$Y_{p}(\mathbf{s}) = \mathbf{X}_{p}\beta_{p} + \alpha_{p}(\mathbf{s}) + \gamma_{p}(\mathbf{s}) + \varepsilon_{p}(\mathbf{s})$$



Multivariate modeling: backfitting

Recall:

$$Y_{1}(\mathbf{s}) = \mathbf{X}_{1}\beta_{1} + \alpha_{1}(\mathbf{s}) + \gamma_{1}(\mathbf{s}) + \varepsilon_{1}(\mathbf{s})$$
$$\vdots$$
$$Y_{p}(\mathbf{s}) = \mathbf{X}_{p}\beta_{p} + \alpha_{p}(\mathbf{s}) + \gamma_{p}(\mathbf{s}) + \varepsilon_{p}(\mathbf{s})$$

Extending the 'classical' backfitting approach to dependent data:

repeat until convergence
 repeat until convergence
 estimate fixed effects
 for all 'stochastic' effects
 estimate parameters
 predict smooth field



Multivariate modeling: backfitting

- ► Intuitive, stable.
- Computationally easy to implement, handles very large datasets. Limitation of handling is one single $\alpha_i(\mathbf{s})$ or $\gamma_i(\mathbf{s})$ field.
- Known covariance structure:
 Equivalence after convergence and convergence.
- Unknown covariance structure:
 'Nothing' can be said.
- Uncertainties . . .

See Furrer, Sain (2009) StCo; Heersink, Furrer (subm.) LAA.



Multivariate modeling: tapering

Univariate case: tapering based on infill asymptotics and equivalent Gaussian measures.



Multivariate modeling: tapering

Univariate case: tapering based on infill asymptotics and equivalent Gaussian measures.

Idea:

choose an asymptotic framework such that original and tapered covariance matrix are asymptotically equivalent.

Then the difference in the likelihoods tends to zero almost surely.



Difficulty in modeling flexibly joint multivariate processes

Idea for lattice data:

For conditional autoregressive models (CARs), consider one three-dimensional lattice instead of several two-dimensional lattices.

See Sain, Furrer, Cressie (2011) AOAS.









Univariate CAR:

$$E[y_i|y_{-i}] = \mu_i + \sum_{j \neq i} \beta_{ij}(y_j - \mu_j) \qquad \text{Var}[y_i|y_{-i}] = \tau_i^2 + \text{regularity conditions}$$

$$\mathsf{E}[\mathbf{y}_i|\mathbf{y}_{-i}] = \boldsymbol{\mu}_i + \sum_{j \neq i} \mathsf{B}_{ij}(\mathbf{y}_j - \boldsymbol{\mu}_j) \qquad \mathsf{Var}[\mathbf{y}_i|\mathbf{y}_{-i}] = \mathbf{T}_i + \mathsf{regularity co}$$

+ regularity conditions

 Multivariate CAR, alternative formulation: (following slide)









$$\begin{split} \mathsf{E}[y_{ij}|y_{-\{ij\}}] &= \mu_{ij} + \sum_{k \neq i} \beta_{ijkj}(y_{kj} - \mu_{kj}) \\ &+ \sum_{\ell \neq j} \beta_{iji\ell}(y_{i\ell} - \mu_{i\ell}) \\ &+ \sum_{k,\ell \neq i,j} \beta_{ijk\ell}(y_{k\ell} - \mu_{k\ell}) \\ \mathsf{Var}[y_{ij}|y_{-\{ij\}}] &= \tau_{ij}^2 \end{split}$$

+ regularity conditions



Overparameterized! Simplify to:

- constant variance
- constant dependencies

+ symmetry



Overparameterized! Simplify to:

- constant variance
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For example:

 $b_{iji\ell} = \rho_{j\ell}\tau_j\tau_\ell \qquad b_{ijk\ell} = \phi_{j\ell}\tau_j\tau_\ell$

results in:





plus p variance parameters: $\{ au_j\}$



- ► Falls within the framework of a unidimensional lattice model
- Guarantees sparse precision matrices
- Flexibly modeling multivariate spatial dependencies
- MCMC is (often) required and may be difficult to tame
- Possibility to implement asymmetric cross-dependencies, ...



"Multivariate" RCM experiment:

- NCAR/DOE Parallel Climate Model to drive the NCAR/Penn State Mesoscale Model (MM5)
- One control run from 1995–2015 and three future runs (ensemble members) from 2040–2060 (1% annual increase in the amount of greenhouse gases)
- Difference between future and control twenty-year winter (DJF) and summer (JJA) average temperature and average total precipitation
- Spatial fields with $44 \times 56 = 2464$ grid boxes





Differences in DJF temperature ($^{\circ}$ K) and in total precipitation (in).



Hierarchical model:

Data level

data_{var, run} = fixed effects_{var} + random effects_{var, run} + error Process level

fixed effects_{var} = $lat_{var} + lon_{var} + elevation_{var}$ random effects_{var, run} = intercept_{var, run} + MGMRF_{var, run}

Prior level

conjugate or uniform over valid parameter range

→ run MCMC beast







University of Zurich[™] in temperature.









Approximate 95% contours for the average change in temperature and precipitation for five consolidated metropolitan areas.





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Collaboration with: Stephan Sain, NCAR Noel Cressie, OSU Reto Knutti, ETHZ Simon Wood, Bath

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URPP Systems Biology / Functional Genomics





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