

Asymptotic behavior of stochastic algorithms with statistical applications Part 1

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ETICS Annual Research School, Fréjus, 2019

1 Introduction

- Sample mean and variance.
- Recursive estimation of mean and variance.
- Quantile of a continuous distribution.
- Recursive estimation of quantile.

2 Convergence of martingales

- Definition and Examples.
- On Doob's convergence theorem.
- Square integrable martingales.
- Robbins-Siegmund theorem.
- Strong law of large numbers for martingales.
- Central limit theorem for martingales.

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Sample mean and variance.

Let (X_n, X) be a sequence of square integrable **independent and identically distributed** random variables with $\mathbb{E}[X] = m$, $\text{Var}(X) = \sigma^2$. The **sample mean** and the **sample variance** are defined by

$$\bar{X}_n = \frac{1}{n} \sum_{k=1}^n X_k,$$

and

$$S_n^2 = \frac{1}{n} \sum_{k=1}^n (X_k - \bar{X}_n)^2.$$

Goal

→ Recursively estimate the unknown mean and variance

$$\theta = \begin{pmatrix} m \\ \sigma^2 \end{pmatrix}.$$

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Two recursive equations.

We have

$$(n+1)\bar{X}_{n+1} = \sum_{k=1}^n X_k + X_{n+1}.$$

Consequently,

$$(n+1)\bar{X}_{n+1} = n\bar{X}_n + X_{n+1} = (n+1)\bar{X}_n + X_{n+1} - \bar{X}_n,$$

which implies that

$$\bar{X}_{n+1} = \bar{X}_n + \frac{1}{n+1} (X_{n+1} - \bar{X}_n).$$

We also have

$$S_n^2 = \frac{1}{n} \sum_{k=1}^n X_k^2 - \bar{X}_n^2.$$

Consequently,

$$\begin{aligned} (n+1)S_{n+1}^2 &= \sum_{k=1}^{n+1} X_k^2 - (n+1)\bar{X}_{n+1}^2, \\ &= \sum_{k=1}^{n+1} X_k^2 - (n+1)\left(\bar{X}_n + \frac{1}{n+1}(X_{n+1} - \bar{X}_n)\right)^2, \\ &= \sum_{k=1}^{n+1} X_k^2 - (n+1)\bar{X}_n^2 - 2\bar{X}_n(X_{n+1} - \bar{X}_n) - \xi_{n+1} \end{aligned}$$

where

$$\xi_{n+1} = \frac{1}{n+1}(X_{n+1} - \bar{X}_n)^2.$$

We also have

$$S_n^2 = \frac{1}{n} \sum_{k=1}^n X_k^2 - \bar{X}_n^2.$$

Consequently,

$$\begin{aligned} (n+1)S_{n+1}^2 &= \sum_{k=1}^{n+1} X_k^2 - (n+1)\bar{X}_{n+1}^2, \\ &= \sum_{k=1}^{n+1} X_k^2 - (n+1)\left(\bar{X}_n + \frac{1}{n+1}(X_{n+1} - \bar{X}_n)\right)^2, \\ &= \sum_{k=1}^{n+1} X_k^2 - (n+1)\bar{X}_n^2 - 2\bar{X}_n(X_{n+1} - \bar{X}_n) - \xi_{n+1} \end{aligned}$$

where

$$\xi_{n+1} = \frac{1}{n+1}(X_{n+1} - \bar{X}_n)^2.$$

We also have

$$S_n^2 = \frac{1}{n} \sum_{k=1}^n X_k^2 - \bar{X}_n^2.$$

Consequently,

$$\begin{aligned} (n+1)S_{n+1}^2 &= \sum_{k=1}^{n+1} X_k^2 - (n+1)\bar{X}_{n+1}^2, \\ &= \sum_{k=1}^{n+1} X_k^2 - (n+1)\left(\bar{X}_n + \frac{1}{n+1}(X_{n+1} - \bar{X}_n)\right)^2, \\ &= \sum_{k=1}^{n+1} X_k^2 - (n+1)\bar{X}_n^2 - 2\bar{X}_n(X_{n+1} - \bar{X}_n) - \xi_{n+1} \end{aligned}$$

where

$$\xi_{n+1} = \frac{1}{n+1}(X_{n+1} - \bar{X}_n)^2.$$

Two recursive equations.

Therefore,

$$\begin{aligned}(n+1)S_{n+1}^2 &= \sum_{k=1}^n X_k^2 - n\bar{X}_n^2 + X_{n+1}^2 - 2\bar{X}_n X_{n+1} + \bar{X}_n^2 - \xi_{n+1}, \\ &= nS_n^2 + (X_{n+1} - \bar{X}_n)^2 - \xi_{n+1}.\end{aligned}$$

Hence

$$(n+1)S_{n+1}^2 = (n+1)S_n^2 + (X_{n+1} - \bar{X}_n)^2 - S_n^2 - \xi_{n+1},$$

leading to

$$S_{n+1}^2 = S_n^2 + \frac{1}{n+1} \left((X_{n+1} - \bar{X}_n)^2 - S_n^2 \right) - \frac{1}{(n+1)^2} (X_{n+1} - \bar{X}_n)^2.$$

Two recursive equations.

Therefore,

$$\begin{aligned}(n+1)S_{n+1}^2 &= \sum_{k=1}^n X_k^2 - n\bar{X}_n^2 + X_{n+1}^2 - 2\bar{X}_n X_{n+1} + \bar{X}_n^2 - \xi_{n+1}, \\ &= nS_n^2 + (X_{n+1} - \bar{X}_n)^2 - \xi_{n+1}.\end{aligned}$$

Hence

$$(n+1)S_{n+1}^2 = (n+1)S_n^2 + (X_{n+1} - \bar{X}_n)^2 - S_n^2 - \xi_{n+1},$$

leading to

$$S_{n+1}^2 = S_n^2 + \frac{1}{n+1} \left((X_{n+1} - \bar{X}_n)^2 - S_n^2 \right) - \frac{1}{(n+1)^2} (X_{n+1} - \bar{X}_n)^2.$$

A recursive matrix equation.

Denote

$$\hat{\theta}_n = \begin{pmatrix} \bar{X}_n \\ S_n^2 \end{pmatrix}.$$

It follows from the previous calculation that

$$\hat{\theta}_{n+1} = \hat{\theta}_n + \frac{1}{n+1} F(\hat{\theta}_n, X_{n+1}) + \frac{1}{(n+1)} r_{n+1}$$

where

$$F(\hat{\theta}_n, X_{n+1}) = \begin{pmatrix} X_{n+1} - \bar{X}_n \\ (X_{n+1} - \bar{X}_n)^2 - S_n^2 \end{pmatrix}$$

and

$$r_{n+1} = \begin{pmatrix} 0 \\ -\frac{1}{(n+1)} (X_{n+1} - \bar{X}_n)^2 \end{pmatrix}.$$

A recursive matrix equation.

However,

$$\mathbb{E}[\mathbf{F}(\hat{\theta}_n, \mathbf{X}_{n+1}) | \mathcal{F}_n] = \mathbf{f}(\hat{\theta}_n) + \mathbf{s}_n$$

where $\mathbf{f}(\hat{\theta}_n) = \theta - \hat{\theta}_n$ and

$$\mathbf{s}_n = \begin{pmatrix} 0 \\ (m - \bar{X}_n)^2 \end{pmatrix}.$$

Consequently, we obtain the **martingale decomposition**

$$\hat{\theta}_{n+1} = \hat{\theta}_n + \frac{1}{n+1} \left(\mathbf{f}(\hat{\theta}_n) + \varepsilon_{n+1} + \mathbf{R}_{n+1} \right)$$

where (ε_n) is a **martingale difference sequence**, $\mathbb{E}[\varepsilon_{n+1} | \mathcal{F}_n] = 0$ and the remainder $\mathbf{R}_{n+1} = r_{n+1} + \mathbf{s}_n$ is negligible.

A first warm-up result.

Theorem

Assume that (X_n, X) is a sequence of **iid** random variables such that $\mathbb{E}[X^4]$ is finite. Denote $\mathbb{E}[(X - m)^3] = \mu^3$ and $\mathbb{E}[(X - m)^4] = \tau^4$. Then, we have the almost sure convergence

$$\lim_{n \rightarrow \infty} \hat{\theta}_n = \theta \quad \text{a.s.}$$

In addition, we also have the asymptotic normality

$$\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{\mathcal{L}} \mathcal{N}(\mathbf{0}, \Gamma)$$

where

$$\Gamma = \begin{pmatrix} \sigma^2 & \mu^3 \\ \mu^3 & \tau^4 - \sigma^4 \end{pmatrix}.$$

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Quantile of a continuous distribution.

Let X be a **continuous** random variable with **unknown** distribution function F . Assume that F is **continuous and strictly increasing**.

Definition

For any α in $]0, 1[$, the quantile of order α of X is the unique solution θ_α of the equation $F(x) = \alpha$,

$$F(\theta_\alpha) = \alpha.$$

For the Exponential $\mathcal{E}(\lambda)$ distribution with $\lambda > 0$,

$$\theta_\alpha = -\frac{1}{\lambda} \log(1 - \alpha).$$

Goal

→ Recursively estimate the unknown quantile θ_α .

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Let (X_n) be a sequence of **iid** random variables sharing the same distribution as X . We estimate θ_α by the recursive estimator

$$\hat{\theta}_{n+1} = \hat{\theta}_n - \frac{1}{n+1} (Y_{n+1} - \alpha)$$

where

$$Y_{n+1} = F(\hat{\theta}_n, X_{n+1}) = \mathbf{I}_{\{X_{n+1} \leq \hat{\theta}_n\}}.$$

We clearly have $\mathbb{E}[Y_{n+1} | \mathcal{F}_n] = F(\hat{\theta}_n)$ leading to the **martingale decomposition**

$$\hat{\theta}_{n+1} = \hat{\theta}_n - \frac{1}{n+1} (F(\hat{\theta}_n) + \varepsilon_{n+1} - \alpha)$$

where (ε_n) is a **martingale difference sequence**, $\mathbb{E}[\varepsilon_{n+1} | \mathcal{F}_n] = 0$.

A second warm-up result.

Denote by f the probability density function of X .

Theorem

We have the almost sure convergence

$$\lim_{n \rightarrow \infty} \hat{\theta}_n = \theta_\alpha \quad \text{a.s.}$$

Moreover, *as soon as $f(\theta_\alpha) > 1/2$* , we have the asymptotic normality

$$\sqrt{n}(\hat{\theta}_n - \theta_\alpha) \xrightarrow{\mathcal{L}} \mathcal{N}\left(0, \frac{\alpha(1-\alpha)}{2f(\theta_\alpha) - 1}\right).$$

A second warm-up result.

Consider the **slow down Robbins-Monro algorithm** given by

$$\hat{\theta}_{n+1} = \hat{\theta}_n - \gamma_n (Y_{n+1} - \alpha)$$

where

$$\gamma_n = \frac{1}{n^c} \quad \text{with} \quad \frac{1}{2} < c < 1.$$

At time $n \geq 1$, compute de Cesaro mean

$$\bar{\theta}_n = \frac{1}{n} \sum_{k=1}^n \hat{\theta}_k.$$

A second warm-up result.

We already saw that

$$\bar{\theta}_{n+1} = \bar{\theta}_n + \frac{1}{n+1} (\hat{\theta}_{n+1} - \bar{\theta}_n).$$

Theorem

We have the almost sure convergence

$$\lim_{n \rightarrow \infty} \bar{\theta}_n = \theta_\alpha \quad \text{a.s.}$$

Moreover, we also have the asymptotic normality

$$\sqrt{n}(\bar{\theta}_n - \theta_\alpha) \xrightarrow{\mathcal{L}} \mathcal{N}\left(\mathbf{0}, \frac{\alpha(1-\alpha)}{f^2(\theta_\alpha)}\right).$$

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Let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space with a filtration $\mathbb{F} = (\mathcal{F}_n)$ where \mathcal{F}_n is the σ -algebra of events occurring up to time n .

Definition

Let (M_n) be a sequence of integrable random variables defined on $(\Omega, \mathcal{A}, \mathbb{P})$ such that, for all $n \geq 0$, M_n is \mathcal{F}_n -measurable.

- ① (M_n) is a martingale **MG** if for all $n \geq 0$,

$$\mathbb{E}[M_{n+1} | \mathcal{F}_n] = M_n \quad \text{a.s.}$$

- ② (M_n) is a submartingale **SMG** if for all $n \geq 0$,

$$\mathbb{E}[M_{n+1} | \mathcal{F}_n] \geq M_n \quad \text{a.s.}$$

- ③ (M_n) is a supermartingale **SMG** if for all $n \geq 0$,

$$\mathbb{E}[M_{n+1} | \mathcal{F}_n] \leq M_n \quad \text{a.s.}$$

Let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space with a filtration $\mathbb{F} = (\mathcal{F}_n)$ where \mathcal{F}_n is the σ -algebra of events occurring up to time n .

Definition

Let (M_n) be a sequence of integrable random variables defined on $(\Omega, \mathcal{A}, \mathbb{P})$ such that, for all $n \geq 0$, M_n is \mathcal{F}_n -measurable.

- 1 (M_n) is a martingale **MG** if for all $n \geq 0$,

$$\mathbb{E}[M_{n+1} | \mathcal{F}_n] = M_n \quad \text{a.s.}$$

- 2 (M_n) is a submartingale **SMG** if for all $n \geq 0$,

$$\mathbb{E}[M_{n+1} | \mathcal{F}_n] \geq M_n \quad \text{a.s.}$$

- 3 (M_n) is a supermartingale **SMG** if for all $n \geq 0$,

$$\mathbb{E}[M_{n+1} | \mathcal{F}_n] \leq M_n \quad \text{a.s.}$$

Let $(\Omega, \mathcal{A}, \mathbb{P})$ be a probability space with a filtration $\mathbb{F} = (\mathcal{F}_n)$ where \mathcal{F}_n is the σ -algebra of events occurring up to time n .

Definition

Let (M_n) be a sequence of integrable random variables defined on $(\Omega, \mathcal{A}, \mathbb{P})$ such that, for all $n \geq 0$, M_n is \mathcal{F}_n -measurable.

- ① (M_n) is a martingale **MG** if for all $n \geq 0$,

$$\mathbb{E}[M_{n+1} | \mathcal{F}_n] = M_n \quad \text{a.s.}$$

- ② (M_n) is a submartingale **SMG** if for all $n \geq 0$,

$$\mathbb{E}[M_{n+1} | \mathcal{F}_n] \geq M_n \quad \text{a.s.}$$

- ③ (M_n) is a supermartingale **SMG** if for all $n \geq 0$,

$$\mathbb{E}[M_{n+1} | \mathcal{F}_n] \leq M_n \quad \text{a.s.}$$

Martingales with sums.

Example (Sums)

Let (X_n) be a sequence of integrable and independent random variables such that, for all $n \geq 1$, $\mathbb{E}[X_n] = m$. Denote

$$S_n = \sum_{k=1}^n X_k.$$

We clearly have

$$S_{n+1} = S_n + X_{n+1}.$$

Consequently, (S_n) is a sequence of integrable random variables with

$$\begin{aligned}\mathbb{E}[S_{n+1} | \mathcal{F}_n] &= S_n + \mathbb{E}[X_{n+1} | \mathcal{F}_n], \\ &= S_n + \mathbb{E}[X_{n+1}],\end{aligned}$$

Martingales with sums.

Example (Sums)

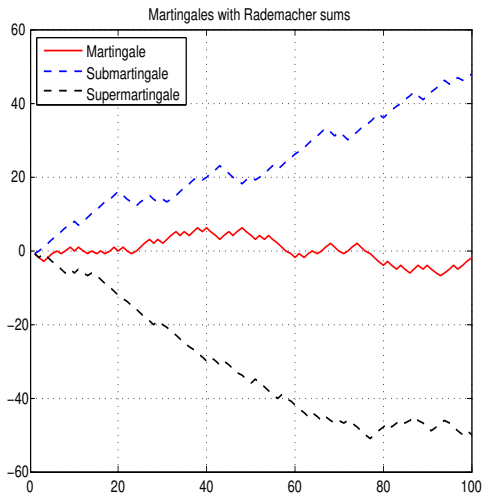
$$\mathbb{E}[\mathbf{S}_{n+1} \mid \mathcal{F}_n] = \mathbf{S}_n + m.$$

- (S_n) is a **martingale** if $m = 0$,
- (S_n) is a **submartingale** if $m \geq 0$,
- (S_n) is a **supermartingale** if $m \leq 0$.

→ It holds for Rademacher $\mathcal{R}(p)$ distribution with $0 < p < 1$ where

$$m = 2p - 1.$$

Martingales with Rademacher sums.



Martingales with products.

Example (Products)

Let (X_n) be a sequence of positive, integrable and independent random variables such that, for all $n \geq 1$, $\mathbb{E}[X_n] = m$. Denote

$$P_n = \prod_{k=1}^n X_k.$$

We clearly have

$$P_{n+1} = P_n X_{n+1}.$$

Consequently, (P_n) is a sequence of integrable random variables with

$$\begin{aligned}\mathbb{E}[P_{n+1} | \mathcal{F}_n] &= P_n \mathbb{E}[X_{n+1} | \mathcal{F}_n], \\ &= P_n \mathbb{E}[X_{n+1}],\end{aligned}$$

Martingales with products.

Example (Products)

$$\mathbb{E}[P_{n+1} | \mathcal{F}_n] = mP_n.$$

- (P_n) is a **martingale** if $m = 1$,
- (P_n) is a **submartingale** if $m \geq 1$,
- (P_n) is a **supermartingale** if $m \leq 1$.

→ It holds for Exponential $\mathcal{E}(\lambda)$ distribution with $\lambda > 0$ where

$$m = \frac{1}{\lambda}.$$

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Doob's convergence theorem.

Theorem (Doob)

Let (M_n) be a **MG**, **sMG**, or **SMG** bounded in \mathbb{L}^1 which means

$$\sup_{n \geq 0} \mathbb{E}[|M_n|] < +\infty.$$

Then, we have the almost sure convergence

$$\lim_{n \rightarrow \infty} M_n = M_\infty \quad \text{a.s.}$$

where M_∞ is an integrable random variable.



Joseph Leo Doob



Jacques Neveu

Convergence of martingales.

Theorem

Let (M_n) be a **MG** bounded in \mathbb{L}^p with $p \geq 1$, which means that

$$\sup_{n \geq 0} \mathbb{E}[|M_n|^p] < +\infty.$$

- 1 If $p > 1$, (M_n) converges almost surely to an integrable random variable M_∞ . The convergence is also true in \mathbb{L}^p .
- 2 If $p = 1$, (M_n) converges almost surely to an integrable random variable M_∞ . The convergence holds in \mathbb{L}^1 **as soon as** (M_n) is **uniformly integrable** that is

$$\lim_{a \rightarrow \infty} \sup_{n \geq 0} \mathbb{E}[|M_n| \mathbf{I}_{\{|M_n| \geq a\}}] = 0.$$

Chow's Theorem.

Theorem (Chow)

Let (M_n) be a **MG** such that for $1 \leq a \leq 2$ and for all $n \geq 1$,

$$\mathbb{E}[|M_n|^a] < \infty.$$

Denote, for all $n \geq 1$, $\Delta M_n = M_n - M_{n-1}$ and assume that

$$\sum_{n=1}^{\infty} \mathbb{E}[|\Delta M_n|^a | \mathcal{F}_{n-1}] < \infty \quad \text{a.s.}$$

Then, we have the almost sure convergence

$$\lim_{n \rightarrow \infty} M_n = M_{\infty} \quad \text{a.s.}$$

where M_{∞} is an integrable random variable.

Exponential Martingale.

Example (Exponential Martingale)

Let (X_n) be a sequence of independent random variable sharing the same $\mathcal{N}(0, 1)$ distribution. For all $t \in \mathbb{R}^*$, let $S_n = X_1 + \dots + X_n$ and denote

$$M_n(t) = \exp\left(tS_n - \frac{nt^2}{2}\right).$$

It is clear that $(M_n(t))$ is a **MG** which converges a.s. to zero. However, $\mathbb{E}[M_n(t)] = \mathbb{E}[M_1(t)] = 1$. It means that $(M_n(t))$ does not converge in \mathbb{L}^1 .

Autoregressive Martingale.

Example (Autoregressive Martingale)

Let (X_n) be the autoregressive process given for all $n \geq 0$ by

$$X_{n+1} = \theta X_n + (1 - \theta)\varepsilon_{n+1}$$

where the initial state $X_0 = p$, $0 < p < 1$ and the parameter $0 < \theta < 1$. Assume that $\mathcal{L}(\varepsilon_{n+1} | \mathcal{F}_n)$ is the Bernoulli $\mathcal{B}(X_n)$ distribution. We can show that $0 < X_n < 1$ and that (X_n) is a **MG** satisfying

$$\lim_{n \rightarrow \infty} X_n = X_\infty \quad \text{a.s.}$$

The convergence also holds in \mathbb{L}^r for all $r \geq 1$. Finally, we can prove that X_∞ has the Bernoulli $\mathcal{B}(p)$ distribution.

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Increasing process.

Definition

Let (M_n) be a square integrable **MG** that is for all $n \geq 1$,

$$\mathbb{E}[M_n^2] < \infty.$$

The **increasing process** associated with (M_n) is given by $\langle M \rangle_0 = 0$ and, for all $n \geq 1$,

$$\langle M \rangle_n = \sum_{k=1}^n \mathbb{E}[\Delta M_k^2 | \mathcal{F}_{k-1}]$$

where $\Delta M_k = M_k - M_{k-1}$.

Example (Increasing Process)

Let (X_n) be a sequence of square integrable and independent random variables such that, for all $n \geq 1$, $\mathbb{E}[X_n] = m$ and $\text{Var}(X_n) = \sigma^2 > 0$. Denote

$$M_n = \sum_{k=1}^n (X_k - m).$$

Then, (M_n) is a martingale and its increasing process reduces to

$$\langle M \rangle_n = \sigma^2 n.$$

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Theorem (Robbins-Siegmund)

Let (V_n) , (A_n) and (B_n) be three positive sequences adapted to $\mathbb{F} = (\mathcal{F}_n)$. Assume that V_0 is integrable and, for all $n \geq 0$,

$$\mathbb{E}[V_{n+1} | \mathcal{F}_n] \leq V_n + A_n - B_n \quad \text{a.s.}$$

Assume also that

$$\sum_{n=0}^{\infty} A_n < +\infty \quad \text{a.s.}$$

- 1 The sequence (V_n) converges a.s. to a random variable V_∞ .
- 2 We also have

$$\sum_{n=0}^{\infty} B_n < +\infty \quad \text{a.s.}$$

Corollary

Let (V_n) , (A_n) , (B_n) and (a_n) be four positive sequences adapted to $\mathbb{F} = (\mathcal{F}_n)$. Assume that V_0 is integrable and, for all $n \geq 0$,

$$\mathbb{E}[V_{n+1} | \mathcal{F}_n] \leq V_n(1 + a_n) + A_n - B_n \quad \text{a.s.}$$

Assume also that

$$\sum_{n=0}^{\infty} a_n < +\infty, \quad \sum_{n=0}^{\infty} A_n < +\infty \quad \text{a.s.}$$

- 1 The sequence (V_n) converges a.s. to a random variable V_∞ .
- 2 We also have

$$\sum_{k=0}^n B_k < +\infty \quad \text{a.s.}$$

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Strong law of large numbers for martingales.

Theorem (Strong Law of large numbers)

Let (M_n) be a square integrable **MG** and denote

$$\langle M \rangle_\infty = \lim_{n \rightarrow \infty} \langle M \rangle_n .$$

① Assume that $\langle M \rangle_\infty < \infty$ a.s. Then, we have

$$\lim_{n \rightarrow \infty} M_n = M_\infty \quad \text{a.s.}$$

② Assume that $\langle M \rangle_\infty = \infty$ a.s. Then, we have

$$\lim_{n \rightarrow \infty} \frac{M_n}{\langle M \rangle_n} = 0 \quad \text{a.s.}$$

→ If it exists a positive sequence (a_n) increasing to infinity such that $\langle M \rangle_n = o(a_n)$ a.s., then we have $M_n = o(a_n)$ a.s.

Strong law of large numbers for martingales, continued

Theorem (Strong Law of large numbers)

Let (M_n) be a square integrable **MG** such that

$$\lim_{n \rightarrow \infty} \langle M \rangle_n = \infty \quad \text{a.s.}$$

① For any positive γ , we have

$$\frac{M_n^2}{\langle M \rangle_n} = o\left(\left(\log \langle M \rangle_n\right)^{1+\gamma}\right) \quad \text{a.s.}$$

② If the increments of (M_n) have conditional moments of order > 2 ,

$$\frac{M_n^2}{\langle M \rangle_n} = O\left(\log \langle M \rangle_n\right) \quad \text{a.s.}$$

Example on sums.

Let (X_n) be a sequence of square integrable and independent random variables such that, for all $n \geq 1$, $\mathbb{E}[X_n] = m$ and $\text{Var}(X_n) = \sigma^2 > 0$. We already saw that

$$M_n = \sum_{k=1}^n (X_k - m)$$

is square integrable **MG** with $\langle M \rangle_n = \sigma^2 n$. It follows from the **SLLN** for martingales that $M_n = o(n)$ a.s. which means that

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n X_k = m \quad \text{a.s.}$$

More precisely, for any positive γ ,

$$\left(\frac{M_n}{n}\right)^2 = \left(\frac{1}{n} \sum_{k=1}^n X_k - m\right)^2 = o\left(\frac{(\log n)^{1+\gamma}}{n}\right) \quad \text{a.s.}$$

Outline

1 Introduction

- Sample mean and variance.
- Recursive estimation of mean and variance.
- Quantile of a continuous distribution.
- Recursive estimation of quantile.

2 Convergence of martingales

- Definition and Examples.
- On Doob's convergence theorem.
- Square integrable martingales.
- Robbins-Siegmund theorem.
- Strong law of large numbers for martingales.
- **Central limit theorem for martingales.**

Central limit theorem for martingales.

Theorem (Central Limit Theorem)

Let (M_n) be a square integrable **MG** and let (a_n) be a sequence of positive real numbers increasing to infinity. Assume that

- ① *It exists a deterministic limit $L \geq 0$ such that*

$$\frac{\langle M \rangle_n}{a_n} \xrightarrow{\mathcal{P}} L.$$

- ② **Lindeberg's condition.** *For all positive ε ,*

$$\frac{1}{a_n} \sum_{k=1}^n \mathbb{E}[|\Delta M_k|^2 \mathbf{1}_{\{|\Delta M_k| \geq \varepsilon \sqrt{a_n}\}} | \mathcal{F}_{k-1}] \xrightarrow{\mathcal{P}} 0$$

where $\Delta M_k = M_k - M_{k-1}$.

Central limit theorem for martingales.

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Central limit theorem for martingales, continued.

Theorem (Central Limit Theorem)

Then, we have the asymptotic normality

$$\frac{1}{\sqrt{a_n}} M_n \xrightarrow{\mathcal{L}} \mathcal{N}(0, L).$$

Moreover, if $L > 0$, we also have

$$\sqrt{a_n} \left(\frac{M_n}{\langle M \rangle_n} \right) \xrightarrow{\mathcal{L}} \mathcal{N}(0, L^{-1}).$$

→ Lyapunov's condition implies Lindeberg's condition : For $b > 2$,

$$\sum_{k=1}^n \mathbb{E}[|\Delta M_k|^b | \mathcal{F}_{k-1}] = O(a_n) \quad \text{a.s.}$$

