

Statistical Learning for Computer Experiments - Application to Aeronautics

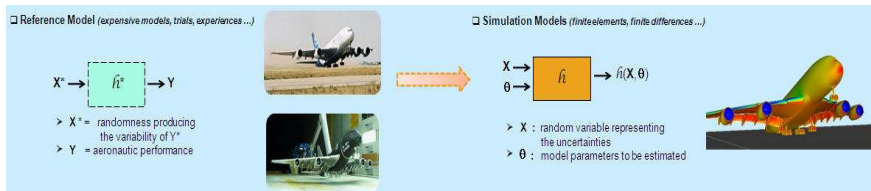
— Ateliers GdR MASCOT-NUM, 17 of May 2011 —

Nabil Rachdi, PhD student at Institut de Mathématiques de Toulouse
Advised by: JC Fort (Paris V), Thierry Klein (Toulouse III)
and Fabien Mangeant (EADS IW)



Context

■ From Real life to Simulated Life...

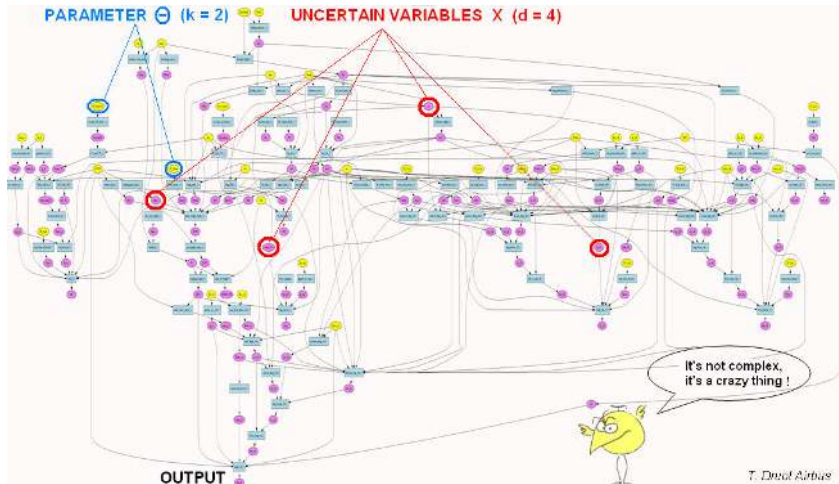


- Y = Variable of Interest (**uncertain !**)
- ρ^* = Quantity of Interest (*quantile, pdf, exceed. probability ...*)
- **Challenge :**
From ref. data Y_1, \dots, Y_n or $(\mathbf{X}_1^*, Y_1), \dots, (\mathbf{X}_n^*, Y_n)$ (**n limited !**)
→ **Choose h and θ to predict ρ^* with simulation model(s) h**

Examples

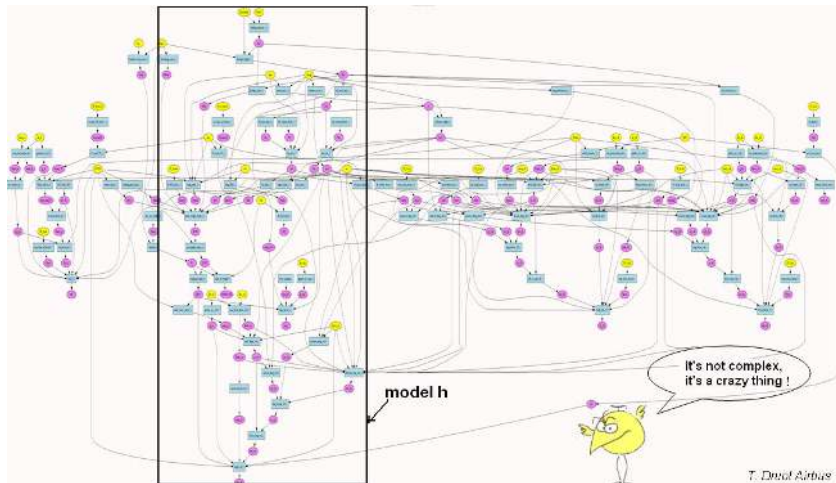
- $Y = TOFL$ (= *Take-Off Field Length*)
 - Quantity of Interest : $\mathbb{P}(TOFL > tofl_{req})$
 - Ref. data $TOFL_1, \dots, TOFL_n$ providing from tests, former aircrafts etc...
 - $\Rightarrow n$ too small for evaluating $\mathbb{P}(TOFL > tofl_{req})$
 - h = aeronautic model with parameters & uncertainties
- $Y = Range$ (= *distance an aircraft can travel*)
 - Quantity of Interest : $\mathbb{P}(Range < range_{req})$
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Numerical Simulations under Uncertainties

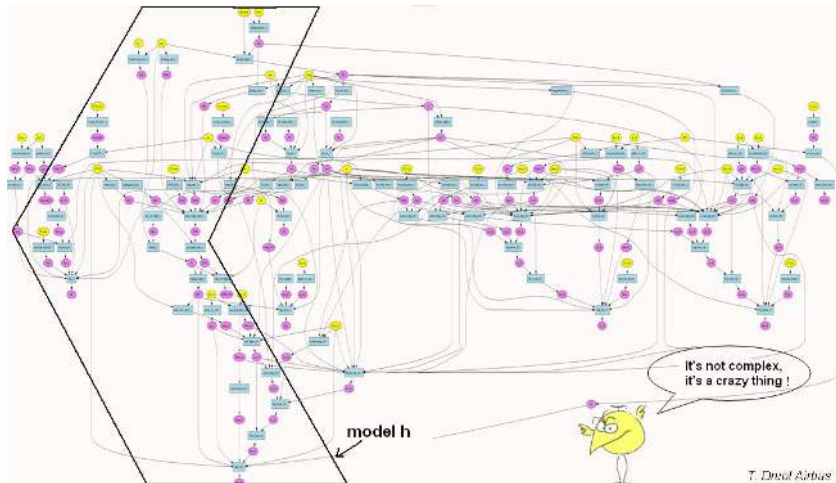


T. Druot Airbus

Numerical Simulations under Uncertainties

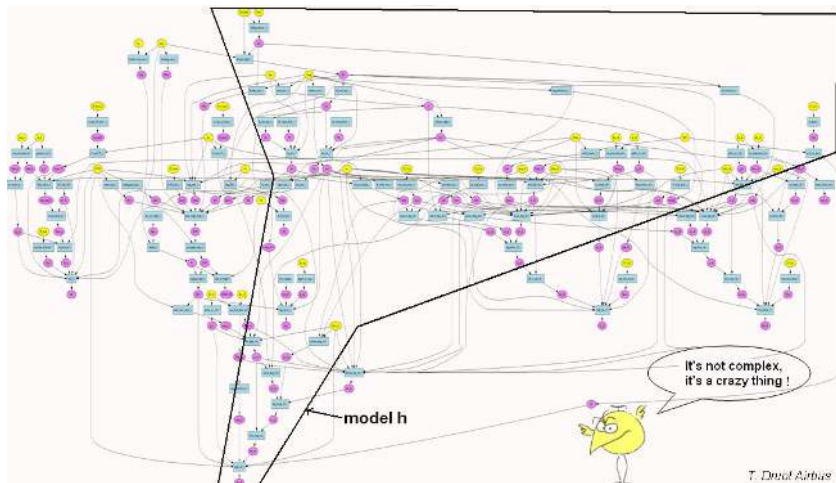


Numerical Simulations under Uncertainties



T. Druot Airbus

Numerical Simulations under Uncertainties



General Framework

- Reference data : $Z_1 = (\mathbf{X}_1, Y_1), \dots, Z_n = (\mathbf{X}_n, Y_n)$ with (**unknown**) dist. Q^Z and denote by Q the marginal dist. of Y
 - $\mathbf{X}_1, \dots, \mathbf{X}_n$ may be **unobserved** (too complex, \neq input codes etc...)
- Models : $\{\mathbf{x} \in \mathcal{X} \mapsto h(\mathbf{x}, \boldsymbol{\theta}) \in \mathcal{Y}, \boldsymbol{\theta} \in \Theta\}$
 - mathematical models : $h(\mathbf{x}, \boldsymbol{\theta}) = \sum_{i=1}^{l=q} \phi(\mathbf{x}) \boldsymbol{\theta}$ etc ...
 - physical/simulation models : $h(\mathbf{x}, \boldsymbol{\theta})$ is the result of a computer code
- Uncertainty : equip \mathcal{X} with a prob. measure P^x - $\mathbf{x} \rightarrow \mathbf{X} \in (\mathcal{X}, P^x)$
 - stochastic codes, Monte-Carlo codes, uncertain variables etc...

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 2 kind of predictions ...
 - Given a new input \mathbf{X} , predict the output Y
 - probabilistic feature on Y (mean, quantile, pdf, exceedance prob., etc...)

Remark :

prediction = param. estimation + computation under the param.

Motivations

- Consider the classical modeling
 - $Y_i = h(\mathbf{X}_i, \boldsymbol{\theta}^*) + \varepsilon_i, \quad i = 1, \dots, n$
 - $\varepsilon_i \sim \mathcal{N}(0, 1)$ independent of \mathbf{X}_i ,
- Suppose $\mathbf{X}_1, \dots, \mathbf{X}_n$ observed
 - classical statistical learning (regression etc...)

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- **Questions** –
- If the \mathbf{X}_i 's are not observed ? How to calibrate ?
(e.g Monte-Carlo codes, input code \neq experimental conditions etc...)
- even if they are observed, should we always use regression parameters for prediction?
- meaning of ...

for example $\mathbb{P}(h(\mathbf{X}, \hat{\boldsymbol{\theta}}_{reg}) > s), pdf_{h(\mathbf{X}, \hat{\boldsymbol{\theta}}_{reg})}$

... *duality between estimation procedure and target prediction*

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- **Robust Prediction**

Calibration - Parameter estimation

- Feature Space \mathcal{F} associated to a feature of Q^Z :
 - $(\mathbf{x} \mapsto \mathbb{E}(Y/\mathbf{X} = \mathbf{x})) \in \mathcal{F} \subset \{\rho : \mathcal{X} \rightarrow \mathcal{Y}\}$
 - $\mathbb{E}(Y) \in \mathcal{F} \subset \mathbb{R}$
 - (pdf of Y) $\in \mathcal{F} \subset \{\rho : \mathcal{Y} \rightarrow \mathbb{R}\}$
 - etc ...
- \mathcal{F} -Contrast function Ψ :

$$\Psi : \mathcal{F} \longrightarrow L_1(Q^Z)$$

$$\rho \longmapsto \Psi(\rho, \cdot) : (\mathbf{x}, y) \in \mathcal{X} \times \mathcal{Y} \longmapsto \Psi(\rho, (\mathbf{x}, y))$$

- Ψ -Risk (to be minimized !) :

$$\mathcal{R}_\Psi(\rho) := \mathbb{E}_{Q^Z} \Psi(\rho, Z) \xrightarrow{\rho = \rho(\boldsymbol{\theta})} \mathcal{R}_\Psi(\boldsymbol{\theta}) := \mathbb{E}_{Q^Z} \Psi(\rho(\boldsymbol{\theta}), Z)$$

Rmk : think \mathcal{R}_Ψ as a "distance" between model and true features

$$\mathcal{R}_\Psi(\boldsymbol{\theta}) \approx \mathcal{D}_\Psi(\rho(\boldsymbol{\theta}), \rho^*)$$

ex. $\mathcal{D}(\mathbb{E}(h(\mathbf{X}, \boldsymbol{\theta})), \mathbb{E}(Y))$, $\mathcal{D}(\text{pdf}_{h(\mathbf{X}, \boldsymbol{\theta})}, \text{pdf}_Y)$, $\mathcal{D}(h(\cdot, \boldsymbol{\theta}), \mathbb{E}(Y/\mathbf{X} = \cdot)) \dots$

Example of contrasts (*"the way of minimizing"*)

- $\mathcal{F} \subset \{\rho : \mathcal{X} \rightarrow \mathcal{Y}\}$

regression contrast

$$\Psi(\rho, (\mathbf{x}, y)) = (y - \rho(\mathbf{x}))^2$$

- $\mathcal{F} \subset \mathbb{R} : \rho = \mathbb{E}(Y), \mathbb{P}(Y > s)$ etc...

mean contrast

$$\Psi(\rho, (\mathbf{x}, y)) = \Psi(\rho, y) = (y - \rho)^2$$

- $\mathcal{F} \subset \{\text{density functions on } \mathcal{Y}\}$

log-contrast

$$\Psi(\rho, (\mathbf{x}, y)) = \Psi(\rho, y) = -\log \rho(y)$$

L_2 -contrast

$$\Psi(\rho, (\mathbf{x}, y)) = \Psi(\rho, y) = \|\rho\|_2^2 - 2\rho(y)$$

- etc...

Application to computer experiments

Recall: $(\mathbf{X}_i, Y_i)_{1..n}, \{\mathbf{x} \in \mathcal{X} \mapsto h(\mathbf{x}, \boldsymbol{\theta}) \in \mathcal{Y}, \boldsymbol{\theta} \in \Theta\}, \mathbf{X} \sim P^{\mathbf{x}} \rightarrow \boldsymbol{\theta}?$

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- Learning Procedures

- Regression: $\theta_{reg} = \text{Argmin}_{\theta \in \Theta} \mathbb{E}_{Q^{\mathbf{x}}} (Y - h(\mathbf{X}, \theta))^2$

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" $\mathcal{D}(\text{pdf}_{h(\mathbf{X}, \theta)}, \text{pdf}_Y)$ "
- (where $Y \sim Q$ and $\rho_{\theta} = \text{pdf of } h(\mathbf{X}, \theta)$ **both unknown** !)

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■ Learning Algorithms (depend on database: $(\mathbf{X}_i, Y_i)_{1..n}$ or Y_1, \dots, Y_n)

- Regression (if \mathbf{X}_i observed !):
 $\hat{\theta}_{reg} = \text{Argmin}_{\theta \in \Theta} \frac{1}{n} \sum_{i=1}^n (Y_i - h(\mathbf{X}_i, \theta))^2$ (well studied)

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- Density contrast for computer experiments (N. Rachdi et al. 2010):

$$\hat{\theta}_{log} = \text{Argmin}_{\theta \in \Theta} - \sum_{i=1}^n \log \left(\sum_{j=1}^m K_b(Y_i - h(\mathbf{X}_j, \theta)) \right)$$

where $\mathbf{X}_1, \dots, \mathbf{X}_m$ i.i.d from $P^{\mathbf{x}}$, $K()$ is a kernel, b bandwidth (ρ_{θ} was estimated by a kernel smoothing)

General Result

Consider the general procedure

$$\hat{\theta} = \underset{\theta \in \Theta}{\text{Argmin}} \sum_{i=1}^n \Psi(\rho^m(\theta), Y_i) \left(\approx \underset{\theta \in \Theta}{\text{Argmin}} \mathcal{D}_{\Psi}(\rho^m(\theta), \rho^n) \right)$$

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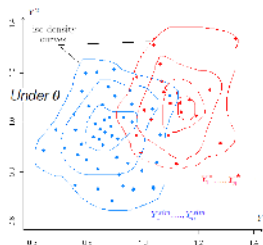
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- **In blue:** Simulated data
- **In red:** Reference data



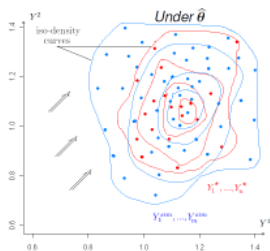
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Theorem: Oracle Inequality (N. Rachdi *et al* 2010)

Under some conditions on the contrast Ψ and under tightness conditions, for all $\varepsilon > 0$, with probability at least $1 - 2\varepsilon$ it holds

$$\mathcal{R}_{\Psi}(\hat{\theta}) \leq \inf_{\theta \in \Theta} (\mathcal{R}_{\Psi}(\theta)) + \frac{K_{(\tilde{\rho}, \Psi)}^{\varepsilon}}{\sqrt{n}} \left(1 + \sqrt{\frac{n}{m}} (K_{(\tilde{\rho}, h)}^{\varepsilon} + B_m) \right)$$

where $K_{(\tilde{\rho}, \Psi)}^{\varepsilon}$, $K_{(\tilde{\rho}, h)}^{\varepsilon}$ some *concentration constants* and B_m a bias factor

Resume

- From **experimental data** $(\mathbf{X}_i, Y_i)_{i=1..n}$ and **simulated data** $h(\mathbf{X}_1, \boldsymbol{\theta}), \dots, h(\mathbf{X}_m, \boldsymbol{\theta})$, we propose others estimation procedures adapted to the *quantity of interest* we want to predict.
- In practice, regression parameters $(\hat{\boldsymbol{\theta}}_{reg})$ may be used to predict a lot of quantities:
 - quantile
 - exceedance probability
 - density function
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Question : How to quantify the estimation procedure error ?

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Question : How to quantify the estimation procedure error ?

- It can be investigated in terms of "*distance between contrasts*"

(N. Rachdi "*A note about predicting with computer experiments*" (In preparation))

⇒ key point: each quantity of interest is viewed as an *Argmin* of some Ψ -Risk \mathcal{R}_Ψ

Academic example

- Let consider

$$Y_i = \sin(X_i) + 0.01 \varepsilon_i \quad i = 1, \dots, n$$

- $X_i \sim \mathcal{N}(0, 1)$
- $\varepsilon_i \sim \mathcal{N}(0, 1)$ independent of X_i

- Model

$$h(\mathbf{X}, \boldsymbol{\theta}) = \theta_0 + \theta_1 X + \theta_2 X^3, \quad \mathbf{X} \sim \mathcal{N}(0, 1)$$

- Goal: Predict the pdf of Y

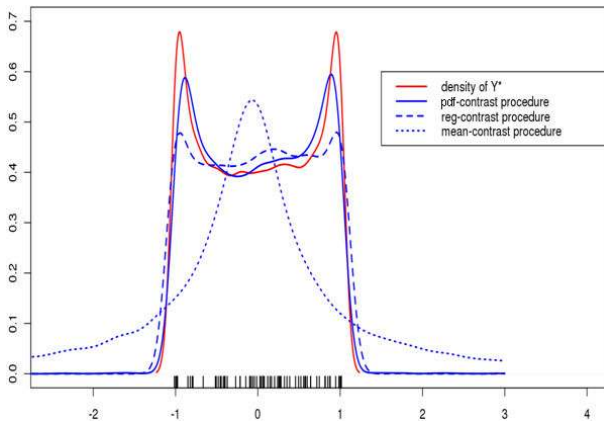
⇒ for this

- Compute some $\hat{\boldsymbol{\theta}}$
- Compute the prediction propagating Uncertainties of \mathbf{X} through the model $\mathbf{x} \mapsto h(\mathbf{x}, \hat{\boldsymbol{\theta}})$

Predictions

- We compute density predictions by propagating uncertainties through models

$h(\mathbf{x}, \hat{\theta}_{pdf})$ (solid line), $h(\mathbf{x}, \hat{\theta}_{reg})$ (dashed line), $h(\mathbf{x}, \hat{\theta}_{mean})$ (dotted line)



Perspectives

■ Academic

- Constants improvement in inequalities
- Central Limit Theorems for the calibration parameter $\hat{\theta}_\Psi$
- Functional analysis of *contrast functions*

■ Industrial Applications

- Run the learning algorithms with *real* computer codes...
- EADS Applications
 - **Non Destructive Testing** (*Prediction*)
combine reference data and simulated data for POD estimation
 - **Electromagnetism** (*Inverse problem*)
Characterize slot parameters from sensors data and uncertain numerical models

Thank you for your attention !