



Weierstraß-Institut für
Angewandte Analysis und Stochastik

Non-Gaussian Component Analysis using Semi Definite Relaxation

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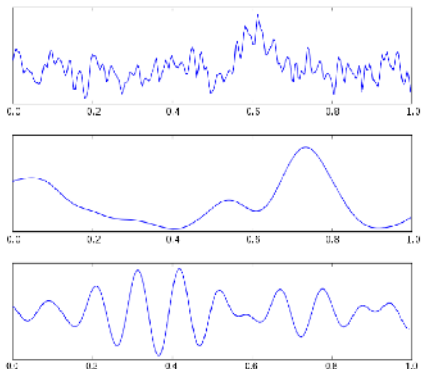
joint with A.Juditsky and E. Diederichs

- 1 Motivation, Data and Problem**
- 2 Semi-parametric framework**
- 3 NGCA procedures**
- 4 Numerical Experiments**

1 Motivation, Data and Problem

- Examples: EEG and microarray data
- Robust risk management
- High dimensional clustering
- Conformational Changes of Biomolecules

Motivating Example: Highdimensional EEG Data



noisy mixed signal of 26 sensors;
top: raw; middle: δ -waves (sleep); bottom: α -waves

R_t , a vector of observed log-returns for a big portfolio.

Problem:

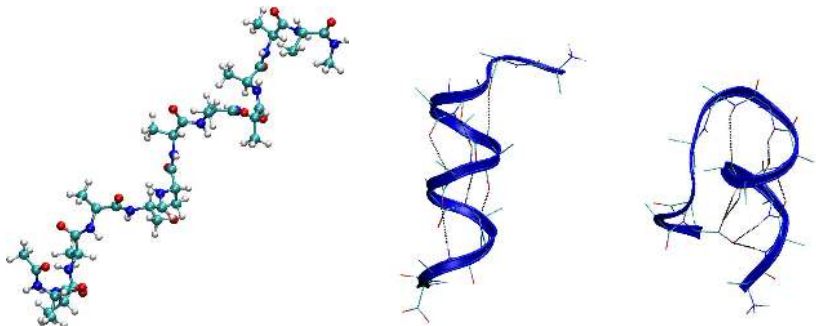
- ▶ Assumption of normal returns do not capture large losses caused by external shocks.
 - ▶ Requires to include non-normal heavy tailed components in the distribution of the portfolio returns.
-
- Chen, Härdle, VS (2010) [GHICA – Risk analysis with GH distributions and independent components](#), J. Empirical Finance, 17 (2010) pp. 255–269.
 - Chen, Härdle, VS (2007) [Portfolio value at risk based on independent components analysis](#), J. Comp. Appl. Math., 205 (2007) pp. 594–607.

Given a sample X_1, \dots, X_n from a measure \mathbb{P} on \mathbb{R}^d , identify the clustering (multimodality) structure of \mathbb{P} .

Curse of dimensionality: non-parametric methods poor for d large.

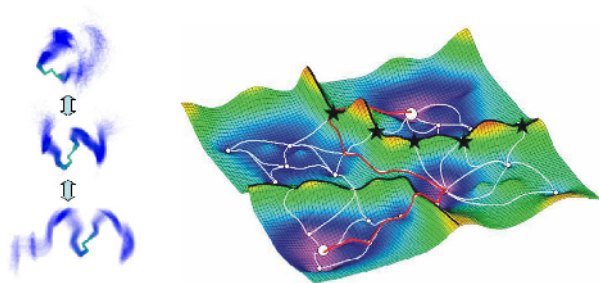
NGCA approach: Gaussian component does not contribute to clustering, focus on non-Gaussian part (with a clear multimodal structure).

Example: folding states of 12-alanine



Most probable large scale shapes of 12-alanine, α -helix and β -sheet

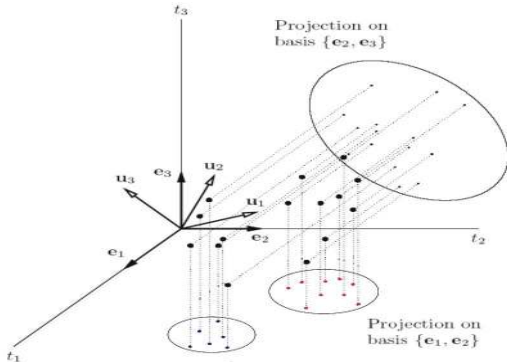
- a. **small and fast** variations around stable geometric mean due to random perturbations of the molecule from the solvent
- b. rare flipping between **long-living** geometric mean configurations of a molecule, called **conformations**



conformational changes of 12-alanine as transition in the landscape of potential energy

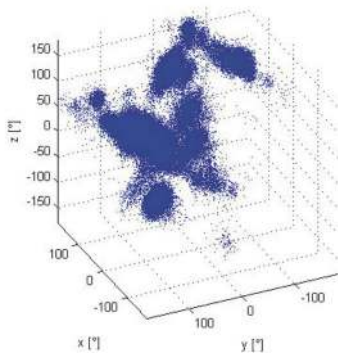
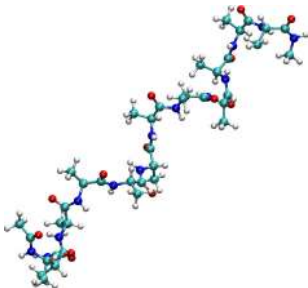
General Picture of Dimension Reduction for Biomolecules

Observation: In conformational dynamics the detection of rare folding events coincides with structural data analysis.



Aim: find a linear combination of rotational angles (dieder angles) spanning a low dimensional conformational subspace.

Example: cluster structure in reduced data (12-alanine)



multimodal component of 12-alanine

Drawbacks of standard methods to detect the cluster structure:

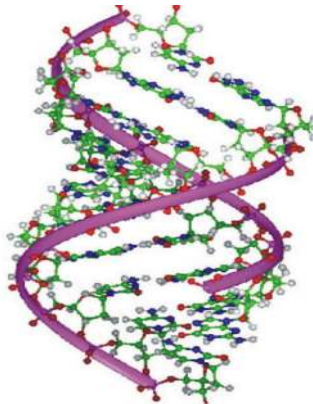
1. Conformational changes are realized by small variations s.t. PCA fails to detect them.
2. Direct Perron Cluster Analysis is unreliable for $10 \leq d$.
3. Fitting of HMM via EM-algorithm is computationally very expensive for $35 \leq d$ and the EM-algorithm has only local convergence.

Strategy for metastability analysis of highdimensional biomolecules:

- a. Reduce highdimensional data with **SDNGCA**.
- b. Fit a **HMM with Gaussian via EM** output and sufficient high number M of hidden states to the data
- c. Consider the resulting Viterbi path, describing the macroscopic dynamics as a realization of a **Markov jump process**.
- d. Perform **Perron Cluster Analysis** (PCCA) to detect the metastable states.

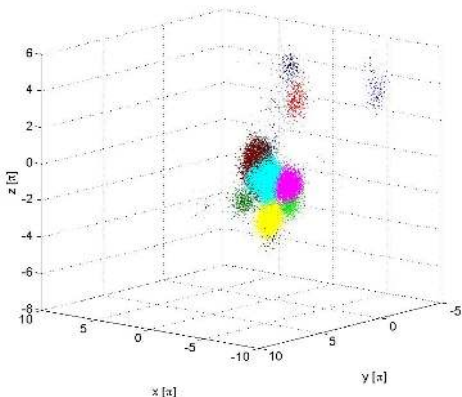
Structure of DNA Oligonucleotide

The trajectory of a 15-AT B-DNA oligonucleotide is simulated by AMBER with explicit water in $d = 84$ dieder angles contains $T = 1 \cdot 10^5$ time steps with each time step of $100fs$ length and covers $1ns$ at $T = 300K$.



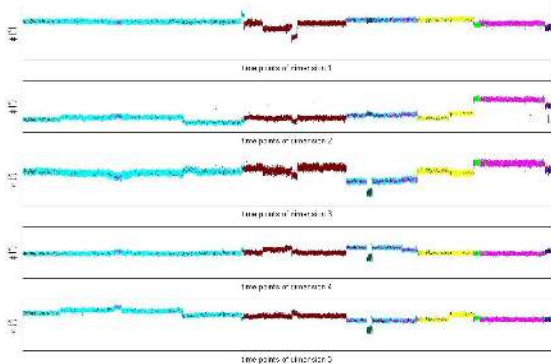
Metastability of reduced B-DNA: 7 states

SDNGCA returns a $9d$ target space with $5d$ multimodal subspace. For illustration we show only a 3 dimensional subspace of the target space with 7 metastable states.



Reduced Gaussian target space of 12-alanine

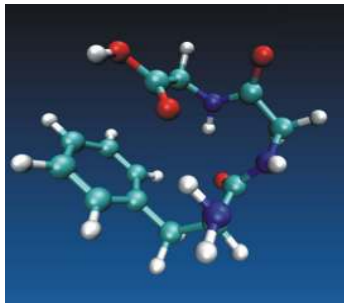
Time Series of reduced B-DNA: 7 states



First five most multimodal components from the target space

Structure of Phenylalanyl-Glycyl-Glycine Tripeptide

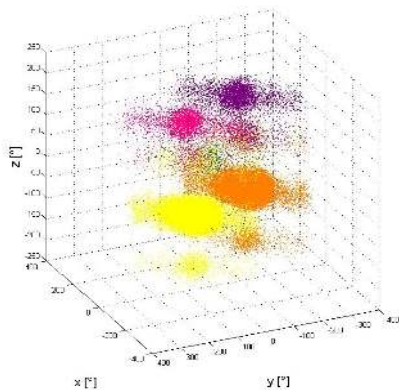
The trajectory simulated by AMBER with implicit water in $d = 11$ dieder angles contains $T = 2 \cdot 10^4$ time steps with each time step of $50fs$ length and covers $0.5ns$ at $T = 300K$.



Structure of Phenylalanyl-Glycyl-Glycine Tripeptide

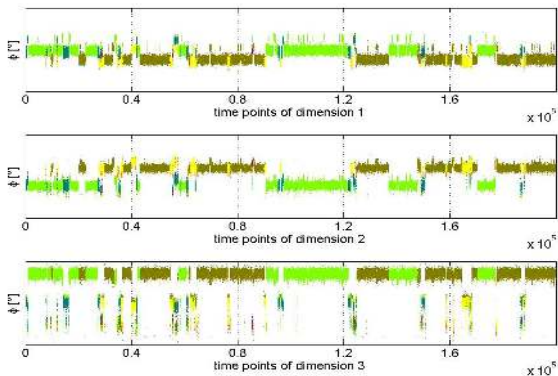
Metastability of reduced PGGT data: 9 states

SDNGCA returns a $4d$ target space with $3d$ multimodal subspace containing 9 metastable states.



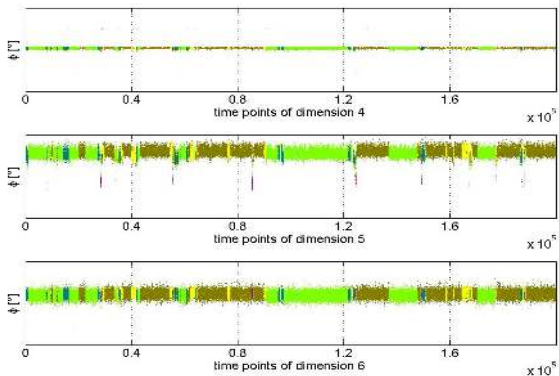
Reduced Gaussian target space of 12-alanine

Time Series of reduced PGGT data: 9 states



first 3 most multimodal components from the target space

Time Series of remaining component



Projection of the data onto the components 4-6.

2 Semi-parametric framework

- Model
- The key idea

Data $X_1, \dots, X_n \in \mathbb{R}^d$ i.i.d., d large. For simplicity $\mathbb{E}X_i = 0$.

Goal: structural analysis.

Basic observation: high dimensional data tends to be normal:

a random projection $X^\top \omega$ is approximately normal for the most of directions ω .

Gaussian component of the data is usually **uninformative** (noise).

Approach: project the data on the **Non-Gaussian component**.

Let X_i i.i.d. with a density $\rho(\cdot)$. Suppose

$$\rho(x) = \phi_{0,\Sigma}(x)q(Tx) \quad (1)$$

- $\phi_{\mu,\Sigma}$, the normal density with parameter (μ, Σ)
- $T : \mathbb{R}^d \rightarrow \mathbb{R}^m$ is a linear operator with $\mathcal{J} = \text{Ker}(T)^\perp$.
- $q : \mathbb{R}^m \rightarrow \mathbb{R}$, $m \leq d$, a smooth nonlinear function.

\mathcal{J} is the target **non-Gaussian subspace**, m is the **non-Gaussian dimension**.

Interpretation: $X = Z + \varepsilon$ where ε is an independent Gaussian noise, Z , a signal.

(1) links **pure Gaussian(PCA)** and **pure non-Gaussian (ICA)** modeling.

Aim: recover \mathcal{J} and possibly m .

Lemma

Assume that $\rho(x)$ is the structured density according to (1). If $\psi(x) \in \mathcal{C}^1(\mathbb{R}^d, \mathbb{R})$ fulfills

$$\mathbb{E}[X\psi(X)] = 0$$

then

$$\beta(\psi) \stackrel{\text{def}}{=} \mathbb{E}[\nabla\psi(X)] \in \mathcal{J}.$$

Moreover, if $\mathbb{E}[X\psi(X)] = \Delta \neq 0$, then there exists $\beta \in \mathcal{J}$ s.t.

$$\|\beta - \beta(\psi)\|_2 \leq \|\Sigma^{-1}\Delta\|_2.$$

Two big steps:

- **Sampling** Build some functions ψ_1, \dots, ψ_M such that

$$\mathbb{E}_n \{ X \psi_j(X) \} = n^{-1} \sum_{i=1}^n X_i \psi_j(X_i) = 0.$$

Then every vector

$$\hat{\beta}_j = \mathbb{E}_n \nabla \psi_j(X) = n^{-1} \sum_{i=1}^n \nabla \psi_j(X_i)$$

belongs \mathcal{J} up to the empirical errors $(\mathbb{E} - \mathbb{E}_n) \nabla \psi_j(X)$ and $\Sigma^{-1}(\mathbb{E} - \mathbb{E}_n) \{ X \psi_j(X) \}$.

- **Reduced Rank Regression** problem: Utilize $\{ \hat{\beta}_j \}$ for recovering the target m -dimensional non-Gaussian subspace.

3 NGCA procedures

- NGCA 1G: Linear projection + PCA
- NGCA 2G: Convex projection
- NGCA 3G: SD Relaxation

Blanchard, Kawanabe, Sugiyama, Sp, K.-R. Müller, [In search of non-Gaussian components of a high-dimensional distribution](#), [J. Mach. Learn. Res.](#), 7 (2006) pp. 247–282.

(1): take any $h(\cdot)$ and consider $\psi(x) = h(x) - \alpha^\top x$.

Select α s.t. $\mathbb{E}_n\{X\psi(X)\} = \mathbb{E}_n\{Xh(X)\} - \mathbb{E}_nXX^\top\alpha = 0$.

Problem: requires to compute and study the inverse of the empirical covariance matrix.

(2): use PCA to recover the non-Gaussian subspace from the $\hat{\beta}_j$'s.

Problem: most of vectors $\hat{\beta}_j$ are uninformative, PCA often fails in dimensions over 10.

Diederichs, Juditsky, Sp, Schütte (2010). [Sparse Non-Gaussian Component Analysis](#). to appear [IEEE of Inf. Theory](#), 2010.

Given functions h_1, \dots, h_L compute

$$\begin{aligned}\widehat{\gamma}_\ell &\stackrel{\text{def}}{=} \mathbb{E}_n[Xh_\ell(X)] \approx \gamma_\ell \stackrel{\text{def}}{=} \mathbb{E}[Xh_\ell(X)] \\ \widehat{\eta}_\ell &\stackrel{\text{def}}{=} \mathbb{E}_n[\nabla h_\ell(X)] \approx \eta_\ell \stackrel{\text{def}}{=} \mathbb{E}[\nabla h_\ell(X)].\end{aligned}$$

Convex projection: given a direction $\xi \in \mathbb{R}^d$, solve

$$\widehat{c} = \underset{c \in \mathbb{R}^L}{\operatorname{argmin}} \left\| \xi - \sum_\ell c_\ell \widehat{\eta}_\ell \right\|_2 \quad \text{subject to} \quad \|c\|_1 \stackrel{\text{def}}{=} \sum_\ell |c_\ell| \leq 1, \quad \sum_\ell c_\ell \widehat{\gamma}_\ell = 0,$$

Define

$$\widehat{\beta} = \widehat{\beta}(\widehat{c}) = \sum_\ell \widehat{c}_\ell \widehat{\eta}_\ell.$$

Consider the functions of the form

$$h_{\omega}(x) \stackrel{\text{def}}{=} h(\omega^{\top} x) e^{-\lambda \|x\|^2/2}$$

with a given function h and a vector $\omega \in \mathcal{B}_d$.

- Choose randomly a set of directions $\{\xi_j\}$, $j = 1, \dots, M$ and for every j a family of directions $\{\omega_{j\ell}\}$, $\ell = 1, \dots, L$.
- compute $\hat{\gamma}_{\ell,j} = \mathbb{E}_n X h_{\omega_{\ell,j}}(X)$ and $\hat{\eta}_{\ell,j} = \mathbb{E}_n [\nabla h_{\omega_{\ell,j}}(X)]$.
- Solve for every $j = 1, \dots, M$

$$\{\hat{c}_{\ell,j}\} = \underset{c \in \mathbb{R}^L}{\operatorname{argmin}} \left\| \xi_j - \sum_{\ell} c_{\ell} \hat{\eta}_{\ell,j} \right\|_2, \quad \text{subject to} \quad \sum_{\ell=1}^L c_{\ell} \hat{\gamma}_{\ell,j} = 0, \|c\|_1 \leq 1$$

leading to

$$\hat{\beta}_j = \sum_{\ell=1}^L \hat{c}_{\ell,j} \hat{\eta}_{\omega_{\ell,j}}$$

Lemma

Let $h(\cdot)$ be bounded and continuously differentiable. For a fixed constant $C = C(h)$, it holds

$$\mathbb{E} \max_{\ell} |\hat{\gamma}_{\ell} - \gamma_{\ell}|^2 + |\hat{\eta}_{\ell} - \eta_{\ell}|^2 \leq C(h)n^{-1} \min\{d, \log L\} =: \varepsilon^2.$$

Suppose to be given the vectors $\hat{\beta}_1, \dots, \hat{\beta}_M$ such that

$$\|(I - \Pi_{\mathcal{J}})\hat{\beta}_j\| \leq \varepsilon$$

where $\Pi_{\mathcal{J}}$ is a projector on a m -dimensional subspace.

Reduced Rank Regression problem: given m , recover \mathcal{J} (or $\Pi_{\mathcal{J}}$) from $\hat{\beta}_1, \dots, \hat{\beta}_M$.

More challenging: recover m and \mathcal{J} .

PCA solution:

$$\hat{\mathcal{J}} = \underset{\dim(\mathcal{J})=m}{\operatorname{argmin}} \sum_j \|(I - \Pi_{\mathcal{J}})\hat{\beta}_j\|^2 = \langle \text{first } m \text{ eigenvectors of } \sum_j \hat{\beta}_j \hat{\beta}_j^\top \rangle.$$

Requires that $\lambda_m(\sum_j \beta_j \beta_j^\top) \geq M\epsilon^2$. Works poorly if most of the $\hat{\beta}_j$'s are **non-informative**.

Rounding ellipsoid approach: (see Yu.Nesterov, 2004) Define the set

$$\mathcal{A} \stackrel{\text{def}}{=} \{\hat{\beta}_1, -\hat{\beta}_1, \hat{\beta}_2, -\hat{\beta}_2, \dots\}.$$

and a centered ellipsoid of minimum volume that encloses \mathcal{A} . Recover \mathcal{J} from \mathcal{E} .

Leads to the accuracy $\|\Pi_{\mathcal{J}} - \Pi_{\hat{\mathcal{J}}}\|$ of order $d^{1/2}\epsilon$.

Structural adaptation idea (Hristache, Juditsky, Polzehl and Sp., 2003):

use the estimated ellipsoid \mathcal{E}_{k-1} as a prior information to improve the quality of estimation in the next step.

Leads to **sequential procedure**: alternate two steps

- estimate the model (vectors β_j) using the given structure
- estimate the structure (ellipsoid \mathcal{E})

Method: sample the directions ξ_j and the vectors $\omega_{\ell,j}$ due to length of semiaxis of \mathcal{E}_{k-1} .

This ensures that a certain fraction of ξ_j , $\hat{\gamma}_{\ell,j}$ and $\hat{\eta}_{\ell,j}$ is informative and hence, the corresponding solutions $\hat{\beta}_j$ are informative as well.

Pros:

- Convex projection helps preserve the individual estimation error;
- Rounding ellipsoid approach is more robust than PCA.

Open questions: choice of informative ξ , estimation of m .

Drawbacks:

- computation of $\widehat{\beta}_j$ using randomly chosen directions $\{\xi_j\}$ is expensive
- computation of Fritz-John ellipsoid of the set $\mathcal{S} := \{\widehat{\beta}_1, -\widehat{\beta}_1, \widehat{\beta}_2, -\widehat{\beta}_2, \dots\}$ always requires use of the inverse covariance matrix $\widehat{\Sigma}^{-1}$.
- Structural adaptation does not work in high dimensions.

Aims at a **direct estimation** of the projector Π on the target space \mathcal{J} from the data by solving a single semidefinite optimization problem.

Problem: given: $\omega_1, \dots, \omega_L \in \mathcal{S}_d$

► suppress the noise via the constraint

$$\sum_{\ell=1} \hat{c}_\ell \hat{\gamma}_\ell = \sum_{\ell=1} \hat{c}_\ell \mathbb{E}_n [Xh(\omega_\ell^\top X)] = 0.$$

► access \mathcal{J} via $\sum_{\ell} c_\ell \hat{\eta}_\ell = \sum_{\ell} c_\ell \mathbb{E}_n \nabla h(\omega_\ell^\top X)$.

Notation: $\hat{U} \stackrel{\text{def}}{=} [\hat{\eta}_1, \dots, \hat{\eta}_L] \in \mathbb{R}^{d \times L}$, $\hat{G} \stackrel{\text{def}}{=} [\hat{\gamma}_1, \dots, \hat{\gamma}_L] \in \mathbb{R}^{d \times L}$.

Minimax Approach (cf. Dalalyan, Juditsky, Sp 2009, JMLR): Solve the problem

$$\hat{\Pi} = \underset{\Pi}{\operatorname{argmin}} \max_c \left\{ \left\| (I - \Pi) \hat{U} c \right\|_2^2 \quad \left. \begin{array}{l} \Pi \text{ is a projector on a} \\ m\text{-dimensional subspace of } \mathbb{R}^d \\ c \in \mathbb{R}^L, \hat{G} c = 0, \|c\|_1 = 1 \end{array} \right\} \quad (2)$$

where Π is a Euclidean projector in \mathbb{R}^d .

- ▶ **Advantage:** shortcut of point estimation and target space reconstruction.
- ▶ **Problem:** (2) is a non-convex, non-smooth, hard optimization problem.

Aim: reduce the original problem to an approximate, convex-concave and smooth problem with an acceptable complexity.

Idea: drop non-convex constraints and solve an approximating semidefinite problem.

Joint with A. Nemirovsky:

i. Use positive semidefinite matrix $X = cc^\top$ as "new variable":

$$\|(I - \Pi)\widehat{U}c\|_2^2 = \text{tr}[\widehat{U}(I - \Pi)\widehat{U}X].$$

ii. Relax $\text{rank} X = 1$ to $|X|_1 \stackrel{\text{def}}{=} \sum |X_{ij}| \leq 1$.

iii. Relax $\widehat{G}c = 0$ to $\text{tr}[\widehat{G}X\widehat{G}] \leq \rho^2$.

iv. Relax $\text{rank} \Pi = m$ to $\text{tr} \Pi = m$, $0 \preceq \Pi \preceq I$.

Leads to a relaxed saddle point convex-concave problem:

$$\min_P \max_X \left\{ \text{tr} \left[\widehat{U}(I - P)\widehat{U}X \right] \mid \begin{array}{l} 0 \preceq P \preceq I, \text{tr}[P] = m, \\ X \succeq 0, |X|_1 \leq 1, \text{tr}[\widehat{G}X\widehat{G}] \leq \rho^2 \end{array} \right\}.$$

Solving the relaxed convex-concave SD problem:

- ▶ For large $L > 10^3$, interior point methods are too expensive;
- ▶ Adopt a subgradient descent-ascent method, e.g. [dual extrapolation method](#) (*Nesterov 2007*);
- ▶ complexity of one step $\mathcal{O}(d \log d)$;
- ▶ precision $\mathcal{O}(\frac{1}{k})$, where k is the number of steps.

Theorem

Let \hat{P} be an *optimal solution of the relaxed SDP* and assume that

- i. Π^* on \mathcal{J} is a convex combination of rank-one matrices $Uc c^\top U^\top$
- ii. c satisfies $Gc = 0$ and $\|c\|_1 \leq 1$.

Then it holds of $\hat{\Pi}$, spanned by m eigenvectors of \hat{P} :

$$\begin{aligned}\|(I - \hat{\Pi})Uc\|_2 &\leq C_1 \sqrt{m+1}(\rho + \lambda_{\min}^{-1}(\Sigma) + \varepsilon) \\ \|\hat{\Pi} - \Pi^*\|_2^2 &\leq C_2(m+1)[(\rho + \varepsilon)\lambda_{\min}^{-1}(\Sigma)]^2.\end{aligned}$$

Aim: improve the estimation error of Π .

Approach:

- i. directional sampling:** choose L directions ω_ℓ uniform from \mathcal{S}_d to compute $\widehat{U} = [\widehat{\eta}_1, \dots, \widehat{\eta}_L] \in \mathbb{R}^{d \times L}$ and $\widehat{G} = [\widehat{\gamma}_1, \dots, \widehat{\gamma}_L] \in \mathbb{R}^{d \times L}$

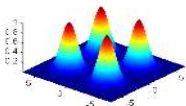
- ii. use result \widehat{P}_k** to get a "better" initial guess for the directions ω_ℓ in iteration $k+1$.

Definition of final projector $\widehat{\Pi}: \widehat{P}_{k+1} := [h_1, \dots, h_d]^T \Lambda [h_1, \dots, h_d]$ and $\widehat{\Pi} := [h_1, \dots, h_m]$.

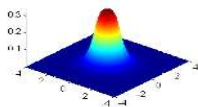
4 Numerical Experiments

- Artificial Distributions

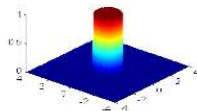
independent Gaussian mixture



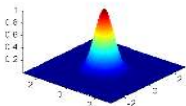
isotropic sub-Gaussian



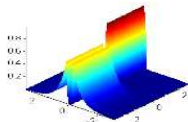
isotropic uniform



isotropic super-Gaussian



dependent Laplacian and uniform

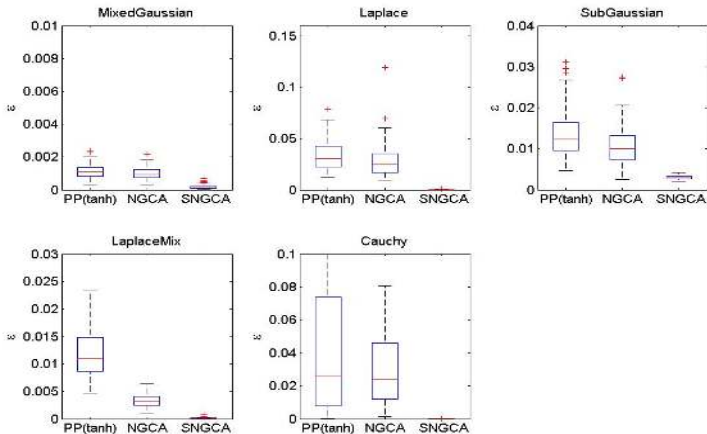


Densities of the non-Gaussian components

The closeness of \mathcal{J} and its estimate $\hat{\mathcal{J}}$ measured by

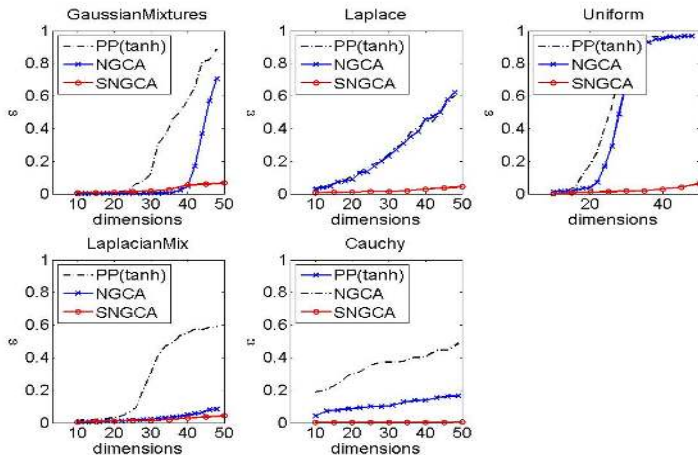
$$\mathcal{E}(\hat{\mathcal{J}}, \mathcal{J}) \stackrel{\text{def}}{=} \frac{1}{2m} \|\Pi_{\mathcal{J}} - \Pi_{\hat{\mathcal{J}}}\|_{Frob}^2 = \frac{1}{m} \sum_{i=1}^m \|(\mathbf{1}_d - \Pi_{\hat{\mathcal{J}}})h_i\|^2 \quad (3)$$

where $\Pi_{\mathcal{J}}$ denotes the orthogonal projection onto \mathcal{J} , $\|\cdot\|_{Frob}$ is the Frobenius norm, $\{h_i\}_{i=1}^m$ is an orthonormal basis of $\hat{\mathcal{J}}$ and $\mathbf{1}_d$ denotes the identity matrix.



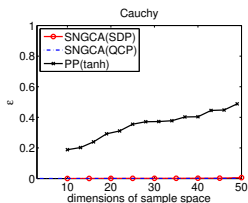
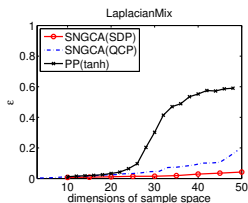
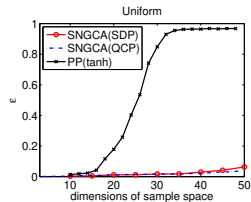
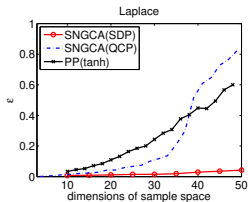
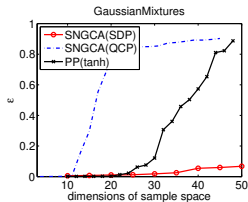
Comparison of PP, NGCA and SDNGCA by estimation error in 10 dimensions.

Comparison of Methods Cont'd: Increase of Dimension I



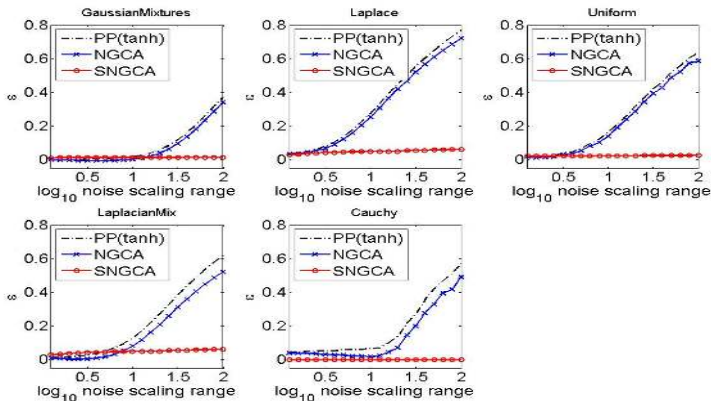
Comparison of PP, NGCA and SNGCA by estimation error for increasing dimensionality .

Comparison of Methods Cont'd: Increase of Dimension II



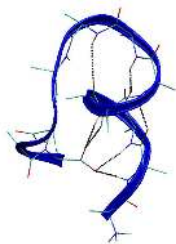
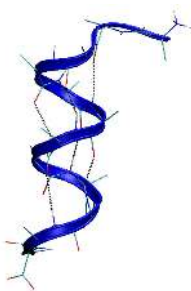
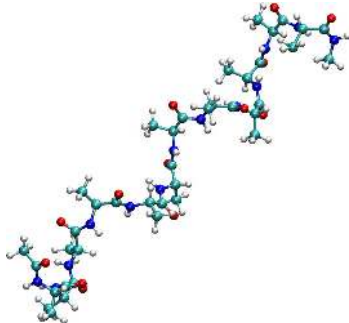
Comparison of PP, SNGCA, SDNGCA by estimation error for increasing dimensionality.

Effect of Numerical Condition on Σ



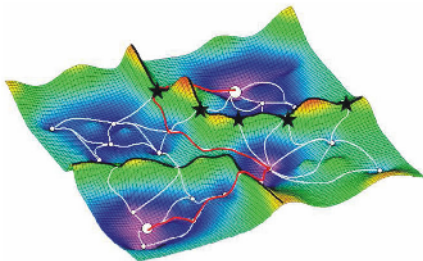
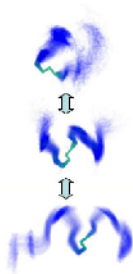
Comparison of PP, NGCA, SDNGCA by estimation error for increasing numerical condition of Σ^{-1} .

Folding states of 12-alanine:



Most probable large scale shapes of 12-alanine, α -helix and β -sheet

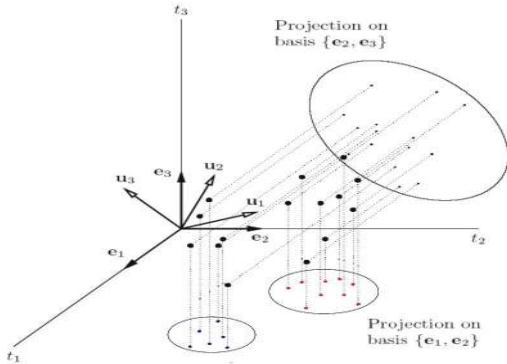
- a. **small and fast** variations around stable geometric mean due to random perturbations of the molecule from the solvent
- b. rare flipping between **long-living** geometric mean configurations of a molecule, called **conformations**



conformational changes of 12-alanine as transition in the landscape of potential energy

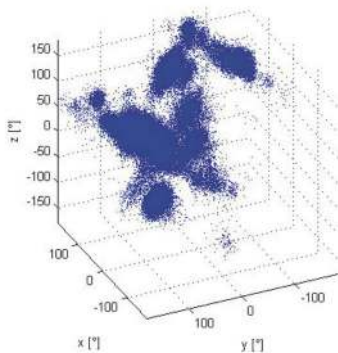
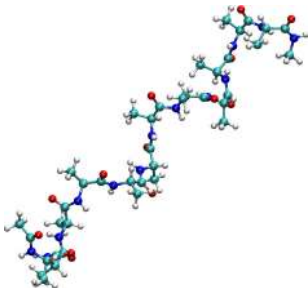
General Picture of Dimension Reduction for Biomolecules

Observation: In conformational dynamics the detection of rare folding events coincides with structural data analysis.



Aim: find a linear combination of rotational angles (dieder angles) spanning a low dimensional conformational subspace.

Example: cluster structure in reduced data (12-alanine)



multimodal component of 12-alanine

Drawbacks of standard methods to detect the cluster structure:

1. Conformational changes are realized by small variations s.t. PCA fails to detect them.
2. Direct Perron Cluster Cluster Analysis is unreliable for $10 \leq d$.
3. Fitting of HMM via EM-algorithm is computationally very expensive for $35 \leq d$ and the EM-algorithm has only local convergence.

- i. Let $\hat{P} := [h_1, \dots, h_d]^T \Lambda [h_1, \dots, h_d]$ and $\hat{P}_{\mathcal{J}} := [h_1, \dots, h_m]$, where \hat{P} the solution of the relaxed SDP.

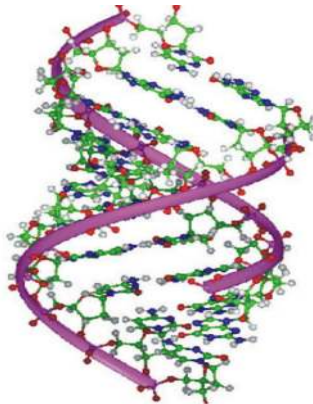
- ii. Project the data X on $[h_1, \dots, h_m]$.

- iii. Compute the well-known dip-index, that is significant to multimodality of every projected data $h_i^\top X$.

- iv. Take the subspace $\mathcal{J}_{multi} \subseteq \mathcal{J}$ as final target space where the projected data with highest dip-index is located.

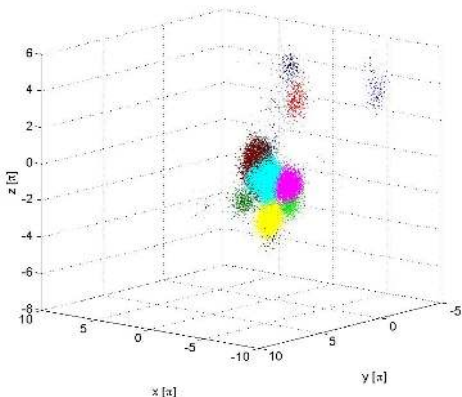
Structure of DNA Oligonucleotide

The trajectory of a 15-AT B-DNA oligonucleotide is simulated by AMBER with explicit water in $d = 84$ dieder angles contains $T = 1 \cdot 10^5$ time steps with each time step of $100fs$ length and covers $1ns$ at $T = 300K$.



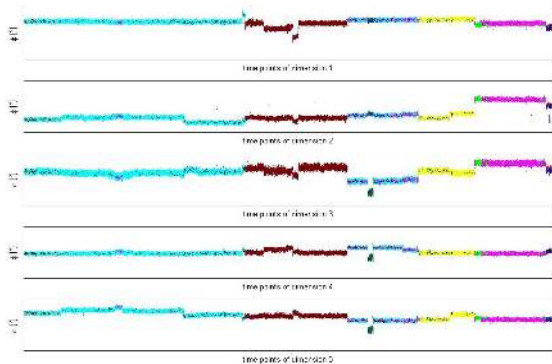
Metastability of reduced B-DNA: 7 states

SDNGCA returns a $9d$ target space with $5d$ multimodal subspace. For illustration we show only a 3 dimensional subspace of the target space with 7 metastable states.



Reduced Gaussian target space of 12-alanine

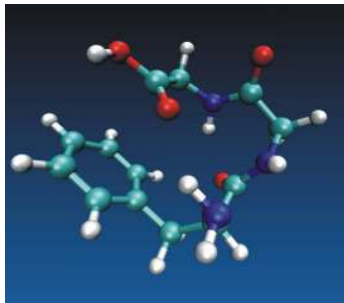
Time Series of reduced B-DNA: 7 states



First five most multimodal components from the target space

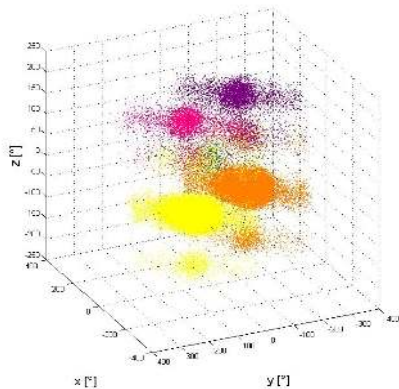
Structure of Phenylalanyl-Glycyl-Glycine Tripeptide

The trajectory simulated by AMBER with implicit water in $d = 11$ dieder angles contains $T = 2 \cdot 10^4$ time steps with each time step of $50fs$ length and covers $0.5ns$ at $T = 300K$.



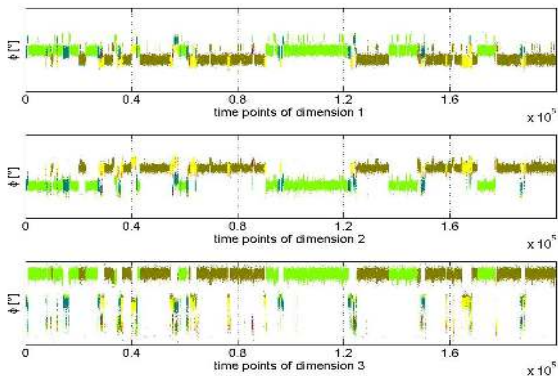
Structure of Phenylalanyl-Glycyl-Glycine Tripeptide

SDNGCA returns a $4d$ target space with $3d$ multimodal subspace containing 9 metastable states.



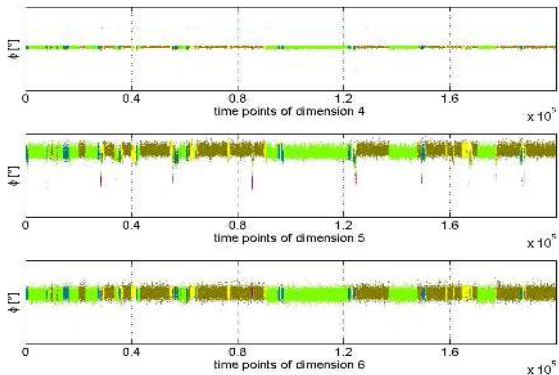
Reduced Gaussian target space of 12-alanine

Time Series of reduced PGGT data: 9 states



first 3 most multimodal components from the target space

Time Series of remaining component



Projection of the data onto the components 4-6.

1. Structural data analysis based on the non-Gaussian vs. Gaussian distinction is effective and computational not too expansive.
2. The Algorithm is independent from any use of $\hat{\Sigma}$.
3. Semidefinite relaxation leads to a statistically more sensitive and structural analysis with not too large complexity $\mathcal{O}(kn^2 + L \log L)$.
4. Convergence rate of the estimation error: $\mathcal{O}((m+1)[\rho \sqrt{\frac{d}{N}} \lambda_{\min}^{-1}(\Sigma)]^2)$.
5. The stochastic reduction of dimensionality works also with stochastic dynamical systems like large biomolecules.

1. Estimation of the reduced dimension m inside of the SDP-approach.
2. Development of criterion to check the new approach in the setting of biomolecules.
3. Development of code with very high performance.