Sequential Design of Computer Experiments for Numerical Dosimetry

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Motivation

Sequential Design of Computer Experiments ...

Estimation of the α -quantile q_{α} of the distribution of $Y = f(\mathbf{X})$, for a given α in (0, 1),

$$q_{\alpha} = \inf \left\{ q : \mathbb{P}(Y \leqslant q) > \alpha \right\} \;.$$

- f is an unknown, expensive-to-evaluate real-valued function
- X is a random vector having a known distribution on a compact subset A ⊆ ℝ^d.

We aim at estimating q_{α} by using as few evaluations of *f* as possible

... for Numerical Dosimetry

At wich level are fetuses exposed to Radio Frequency Electromagnetic Fields ?



Background: Gaussian Process Modelling

Assume that *f* is a sample of a zero-mean Gaussian process (GP) having a covariance function k: GP(0, k(.,.))

Conditionally to $\mathbf{y}_t = (y_1, \dots, y_t)'$, the mean $\mu_t(u)$ and covariance $k_t(u, v)$ are given by

$$\mu_t(\boldsymbol{u}) = \mathbf{k}_t(\boldsymbol{u})'\mathbf{K}_t^{-1}\mathbf{y}_t ,$$

$$k_t(\boldsymbol{u}, \boldsymbol{v}) = k(\boldsymbol{u}, \boldsymbol{v}) - \mathbf{k}_t(\boldsymbol{u})'\mathbf{K}_t^{-1}\mathbf{k}_t(\boldsymbol{v}) ,$$

where $\mathbf{k}_t(u) = [k(x_1, u) \dots k(x_t, u)]'$, ' denotes the matrix transposition, $\mathbf{K}_t = [k(x_i, x_j)]_{1 \le i, j \le t}$, *u* and *v* and the x_i 's are in \mathcal{A} .

Covariance function

Since the SAR is supposed to be smooth, we shall use the square exponential covariance function

$$k_{\mathrm{SE}}(u,v) = \exp\left(-rac{\|u-v\|^2}{2\ell^2}
ight) \;, u,v \in \mathcal{A} \;, \ell > 0 \;,$$

where ||u|| denotes the euclidean norm of u in \mathbb{R}^d ,

Background: Methodologies

Sequential strategies

Bayesian optimization: find the maximum of f, optimizing an acquisition function

- Expected Improvement [Vazquez et al., 2010]
- Confidence Bound Criteria (GP-UCB [Srinivas et al., 2010], Branch and Bounds [De Freitas et al., 2012])

EI has been adapted for

- Contour estimation [Ranjan et al., 2009]
- Estimation of P(Y ≥ s) where s is a given threshold (SUR) [Bect et al., 2012]

Quantile estimation

- Non sequential approach [Oakley, 2004]
- Extension of the SUR criterion [Arnaud et al., 2010]

We really need a sequential strategy, but improvement based criteria demand Monte Carlo samplings of the GP and the conditional GPs, which made them difficult to use for d > 2

Quantile estimation

We shall compare the quantile estimators with $\tilde{q}_{\alpha,m}$ defined by

$$\tilde{q}_{\alpha,m} = \inf\left\{q: \frac{1}{m}\sum_{i=1}^m \mathbb{1}_{\{f(\mathbf{x}_i)\leqslant q\}} > \alpha\right\}$$

where x_1, \ldots, x_m are *m* fixed points in \mathcal{A} . Let $A = \{x_1, \ldots, x_m\} \subset \mathcal{A}$.

Pure exploration criterion

- Minimizes the global uncertainty on the estimation of f
- New point x_{t+1} to add to the set of *t* observations:

$$x_{t+1} \in \operatorname*{arg\,max}_{x \in \mathcal{A}} \sigma_t(x)$$
.

 Propose methodologies more adapted to our quantile estimation issue to realize the exploration-exploitation trade-off

GPS

- ► Let $\mu_t^U(x) = \mu_t(x) + \sqrt{\beta_t}\sigma_t(x)$ and $\mu_t^L(x) = \mu_t(x) \sqrt{\beta_t}\sigma_t(x)$ with $\beta_t = 2\ln\left(\frac{\pi^2 t^2}{6}\right) + 2\ln\left(\frac{m}{\delta}\right)$ where *m* is the cardinal of A
- Let $\hat{q}^U_{\alpha,t}$ and $\hat{q}^L_{\alpha,t}$ be the estimators of the α -quantile of μ^U_t and μ^L_t

$$\hat{\boldsymbol{q}}_{\alpha,t}^{\boldsymbol{U}} = \inf\left\{\boldsymbol{q}: \frac{1}{m}\sum_{i=1}^{m} \mathbb{1}_{\{\boldsymbol{\mu}_{t}^{\boldsymbol{U}}(\mathbf{x}_{i}) \leqslant \boldsymbol{q}\}} > \alpha\right\}$$
$$\hat{\boldsymbol{q}}_{\alpha,t}^{\boldsymbol{L}} = \inf\left\{\boldsymbol{q}: \frac{1}{m}\sum_{i=1}^{m} \mathbb{1}_{\{\boldsymbol{\mu}_{t}^{\boldsymbol{L}}(\mathbf{x}_{i}) \leqslant \boldsymbol{q}\}} > \alpha\right\}$$

Proposition

For all δ in (0, 1), for all $t \ge 1$, with probability greater than $(1 - \delta)$,

$$\tilde{q}_{\alpha,m} \in [\hat{q}_{\alpha,t}^L, \hat{q}_{\alpha,t}^U]$$
.

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GPS (cont.)

Let $U_{\alpha,t}$ and $L_{\alpha,t}$ be the following sets :

$$U_{\alpha,t} = \left\{ \boldsymbol{x} \in \mathbf{A} : \mu_t^{\boldsymbol{U}}(\boldsymbol{x}) \ge \hat{\boldsymbol{q}}_{\alpha,t}^{\boldsymbol{L}} \right\} \text{ and } \boldsymbol{L}_{\alpha,t} = \left\{ \boldsymbol{x} \in \mathbf{A} : \mu_t^{\boldsymbol{L}}(\boldsymbol{x}) \le \hat{\boldsymbol{q}}_{\alpha,t}^{\boldsymbol{U}} \right\} , \ t \ge 1.$$

Proposition

With probability greater than $(1 - \delta)$, for all $t \ge 1$,

$$|\hat{\boldsymbol{q}}_{lpha,t} - \tilde{\boldsymbol{q}}_{lpha,m}| \leqslant \sqrt{eta_t} \sup_{\mathbf{x} \in U_{lpha,t}} \sigma_t(\mathbf{x}) \; .$$

Criterion

 x_{t+1} to add to the set of *t* observations is such that:

 $x_{t+1} \in \underset{x \in U_{\alpha,t}}{\operatorname{arg\,max}} \sigma_t(x)$.

Illustration : 1D Gaussian Process sample path





$$t = 4$$











t = 8

0.4 0.6

0.8

0.2

t = 10



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GPS+

- Let $S_{\alpha,t} \subseteq A$ be the compact subset such that $S_{\alpha,t} = \prod_{i=1}^{d} [x_{\min,t}^{(i)}, x_{\max,t}^{(i)}] \times \cdots \times [x_{\min}^{(d)}, x_{\max}^{(d)}]$. Here $x_{\min,t}^{(i)}$ and $x_{\max,t}^{(i)}$ denote the smallest (resp. the largest) *i*th component of the points in $\bar{U}_{\alpha,t}$
- $\blacktriangleright \text{ where } \bar{U}_{\alpha,t} = \left\{ x \in \mathbf{S}_{\alpha,t-1} : \mu_t^U(x) \ge \hat{q}_{\alpha,t}^L \right\}$
- where $S_{\alpha,t} = \{x_{t,1}, \dots, x_{t,m_t}\} \cup \overline{U}_{\alpha,t}$,
- where $\{x_{t,1}, \ldots, x_{t,m_t}\}$ are m_t points randomly chosen in $S_{\alpha,t}$
- By convention $S_{\alpha,0} = A$.

Criterion

 x_{t+1} to add to the set of *t* observations is such that:

$$x_{t+1} \in \operatorname*{arg\,max}_{x \in \overline{U}_{\alpha,t}} \sigma_t(x)$$
.

Note: Since the size of the grid varies at each iteration of the process, we use $\beta_t = 2 \ln \left(\frac{\pi^2 t^2}{6}\right) + 2 \ln \left(\frac{|S_{\alpha,t-1}|}{\delta}\right)$.

Illustration : 2D Gaussian Process sample path



t = 2



t = 62



t = 102



Numerical Dosimetry ?

In general

Virtually expose human 3D-models to one source of EMF in order to evaluate the Specific Absorption Rate (the SAR, in $W.kg^{-1}$) SAR computation in our case is done through Finite Difference in Time Domain (FDTD) method The SAR depends on

- the geometry of the models
- the dielectric properties of the tissues
- the type and position of the EMF source

Fetus exposure

- Very few models are available
- The simulations are expensive in terms of computational load
- The preparation of the simulations is very complex

We focus on the fetal brain exposure

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Application I: GPS, Japanese model and plane wave

- Plane wave exposure: far field sources (base stations antennas, WiFi boxes)
- 900 MHz vertically polarized electromagnetic plane waves with a 1 Volt per meter amplitude
- Start by performing 5 randomly chosen evaluations of the SAR in order to have an estimation of *I*



Application I: GPS, Japanese model and plane wave (cont.)









Contour plot and observations



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Application II: GPS+, Victoria and Samsung Galaxy Tab

- Model Victoria is sitting working on her Samsung Galaxy Tab at 3G frequency (1940 MHz)
- 3 parameters: height, nearness and slope of the tablet
- Start by performing 20 evaluations of the SAR from a LHS in order to have an estimation of /



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Application II: GPS+, Victoria and Samsung Galaxy Tab (cont.)



Conclusion

- We propose two novel sequential approaches for quantile estimation
- Successfully applied to real data coming from numerical dosimetry