## Uncertainty quantification and visualization for functional random variables

## MascotNum Workshop 2014

S. Nanty ${ }^{1,3}$<br>C. Helbert ${ }^{2}$<br>A. Marrel ${ }^{1}$<br>$N$. Pérot ${ }^{1}$<br>C. Prieur ${ }^{3}$<br>${ }^{1}$ CEA, DEN/DER/SESI/LSMR, F-13108, Saint-Paul-lez-Durance, France<br>${ }^{2}$ Ecole Centrale de Lyon, France<br>${ }^{3}$ Universite Joseph Fourier and INRIA, France

April 23rd 2014


## Introduction

- Identify/characterize the statistical properties of functional random variables.
- The variables are dependent and linked to a scalar (or vectorial) covariate.
- Propose a methodology of uncertainty characterization in order to:
- get an estimate of the joint probability density function of the variables,
- simulate new samples according to the estimated distribution,
- adapt visualization tools to identify uncertainty characteristics of dependent functional variables.


## Problem description

- Let $f_{1}, \ldots, f_{m}: I \times \Omega \rightarrow \mathbb{R}$ be dependent functional random variables.
- Let $Y$ be a random variable, called covariate.
- Let $\mathcal{M}$ be a computer code/simulator such that

$$
Y=\mathcal{M}\left(f_{1}, \ldots, f_{m}\right)
$$

- Let $f_{j}^{i}$ be the $i^{\text {th }}$ realization of the $j^{\text {th }}$ functional random variable, for $1 \leq i \leq n, 1 \leq j \leq m$.


## Proposed methodology

- Two main steps:

1. Decomposition on a reduced functional basis, taking into account the covariate
2. Modeling of the probability density function of the decomposition coefficients


## Table of Contents

Dimension reduction by functional decomposition

Estimation of coefficient probability density function

Illustration on an analytical example

Application on a nuclear safety test case

Associated uncertainty visualization tool

Conclusion

## Decomposition on a functional basis

## Definition

Let $f: I \rightarrow \mathbb{R}, x \in I$.

$$
f(x)=\sum_{k=1}^{+\infty} \alpha_{k} \phi_{k}(x)
$$

- $\alpha_{k}$ coefficients,
- $\phi_{k}$ basis functions


## Decomposition on a functional basis

## Definition

Let $f: I \rightarrow \mathbb{R}, x \in I$.

$$
\hat{f}(x)=\sum_{k=1}^{d} \alpha_{k} \phi_{k}(x)
$$

- $\alpha_{k}$ coefficients,
- $\phi_{k}$ basis functions,
- $d$ basis size


## Partial Least Squares regression

- Let $X(n \times p)$ and $Y(n \times q)$ data matrices of respectively observable and predicted variables.
- $X$ and $Y$ are centered and standardized.
- Principle: linear regression between the projections of $X$ and $Y$ in a new space, called latent variables, whose correlation is maximal.


## Partial Least Squares regression

- Let $X(n \times p)$ and $Y(n \times q)$ data matrices of respectively observable and predicted variables.
- $X$ and $Y$ are centered and standardized.
- Principle: linear regression between the projections of $X$ and $Y$ in a new space, called latent variables, whose correlation is maximal.


## Algorithm of PLS regression [Wold, 1975]

- Initialization: $X_{0}=X, Y_{0}=Y$
- At each step $h$, we are seeking for the latent variables $\alpha_{h}=X_{h-1} u_{h}$ and $\omega_{h}=Y_{h-1} v_{h}$ solutions of

$$
\max _{\left\|u_{h}\right\|=\left\|v_{h}\right\|=1} \operatorname{cov}\left(X_{h-1} u_{h}, Y_{h-1} v_{h}\right)
$$

- "Deflation": $X_{h}=X_{h-1}-\alpha_{h} \phi_{h}^{T}$, with $\phi_{h}=X_{h-1} \alpha_{h} /\left(\alpha_{h}^{T} \alpha_{h}\right)$


## Partial Least Squares decomposition

- It can be deduced from the deflation step that $X$ can be written as follows:

$$
X=A \Phi^{T}+\epsilon
$$

where the column vectors of $A$ and $\Phi$ are respectively $\alpha_{h}$ and $\phi_{h}$ and $\epsilon$ are the residuals.

## Partial Least Squares decomposition

- It can be deduced from the deflation step that $X$ can be written as follows:

$$
X=A \Phi^{T}+\epsilon
$$

where the column vectors of $A$ and $\Phi$ are respectively $\alpha_{h}$ and $\phi_{h}$ and $\epsilon$ are the residuals.

- Let the column vectors of $X$ be functions discretized on $p$ points and $Y$ be the covariate.
$\Rightarrow A$ is the matrix of coefficients of the decomposition.
$\Rightarrow \Phi$ is the matrix of basis functions.


## Partial Least Squares decomposition

- It can be deduced from the deflation step that $X$ can be written as follows:

$$
X=A \Phi^{T}+\epsilon
$$

where the column vectors of $A$ and $\Phi$ are respectively $\alpha_{h}$ and $\phi_{h}$ and $\epsilon$ are the residuals.

- Let the column vectors of $X$ be functions discretized on $p$ points and $Y$ be the covariate.
$\Rightarrow A$ is the matrix of coefficients of the decomposition.
$\Rightarrow \Phi$ is the matrix of basis functions.
- Basis functions are fitted to data, and
- adjusted to maximize the correlation between the functions and the covariate.


## Simultaneous PLS decomposition

- Objective: extend PLS decomposition to deal with multiple functional dependent variables simultaneously
- We suppose that functions $f_{1} \ldots f_{m}$ are correlated and have common reduction directions.
- Let $t_{1}<\cdots<t_{p} \in I$
- Let $\mathbf{f}_{i}=\left(f_{i}\left(t_{1}\right), \ldots, f_{i}\left(t_{p}\right)\right)$ be the discretized version of $f_{i}$, $i=1, \ldots, m$.
- Let each column vector of $X$ be:

$$
\left(\mathbf{f}_{1}, \ldots, \mathbf{f}_{m}\right) \in \mathbb{R}^{d m}
$$

- Simultaneous PLS decomposition consists in applying the PLS decomposition to the previously defined matrix $X$.
$\rightarrow$ SPLS decomposition


## Objectives

- Estimate the probability density function (pdf) of $d$ coefficients from SPLS decomposition
- High dimension: $d>10$
$\Rightarrow$ kernel density estimation not adapted
$\rightarrow$ Solution: Gaussian mixture model



## Gaussian Mixture

- Probability density function of a Gaussian mixture:

$$
g\left(\alpha \mid \mu_{1}, \Sigma_{1}, \ldots, \mu_{G}, \Sigma_{G}\right)=\sum_{k=1}^{G} \tau_{k} \phi\left(\alpha \mid \mu_{k}, \Sigma_{k}\right), \forall \alpha \in \mathbb{R}^{d}
$$

- $G$ clusters
- $n$ sample points
- $\phi$ : Gaussian probability density function
- $\tau_{k}, \mu_{k}, \Sigma_{k}$ : proportion, mean and covariance matrix of cluster $k$


## Gaussian Mixture

- Probability density function of a Gaussian mixture:

$$
g\left(\alpha \mid \mu_{1}, \Sigma_{1}, \ldots, \mu_{G}, \Sigma_{G}\right)=\sum_{k=1}^{G} \tau_{k} \phi\left(\alpha \mid \mu_{k}, \Sigma_{k}\right), \forall \alpha \in \mathbb{R}^{d}
$$

- $G$ clusters
- $n$ sample points
- $\phi$ : Gaussian probability density function
- $\tau_{k}, \mu_{k}, \Sigma_{k}$ : proportion, mean and covariance matrix of cluster $k$
- Advantages / drawbacks
+ Fast algorithm for parameter estimation
+ Very fast simulation of a new realization
+ Can be used in dimension $d>10$
- parametric model: modeling hypothesis
- Number of clusters to be determined
- Expectation-Maximization algorithm (EM) [Dempster et al., 1977] estimates the parameters of the Gaussian mixture model.
- Let us introduce $z_{i k}$, the probability of the $i^{\text {th }}$ point to be in the $k^{\text {th }}$ cluster.


## Expectation Minimization algorithm:

1. Initialize parameters $\tau_{k}^{(0)}, \mu_{k}^{(0)}$ et $\Sigma_{k}^{(0)}$
2. Expectation Step: Compute $z_{i k}^{(j)}$
3. Maximization Step: Compute $\tau_{k}^{(j+1)}, \mu_{k}^{(j+1)}, \Sigma_{k}^{(j+1)}$
4. Repeat steps $2-3$ until convergence

## Number of parameters reduction

- Total number of Gaussian mixture parameters:

$$
N_{T}=G-1+G d+G \frac{d(d+1)}{2}
$$

- $G$ : number of clusters in the model
- $N_{T}$ increases quickly with the dimension $d$
$\rightarrow$ Solution: sparse covariance matrices estimation


## Two methods

- sEM method: penalizing the inverses of covariance matrices [Krishnamurthy, 2011]
- sEM2 method: penalizing the covariance matrices


## sEM method

Penalizing the inverses of covariance matrices

- A lasso penalization on the inverses of the covariance matrices is added in the maximization step:

$$
\hat{\Sigma}_{k}=\operatorname{argmax}_{\Sigma_{k}}(\ell) \quad-\quad \hat{\Sigma}_{k}=\operatorname{argmax}_{\Sigma_{k}}\left(\ell-\lambda\left\|\Sigma_{k}^{-1}\right\|_{1}\right)
$$

- $\|M\|_{1}=\sum_{i, j=1}^{p} M_{i, j}$.
- The penalization parameter $\lambda$ is chosen bycross-validation.
- The penalized maximization is solved by [Friedman et al., 2008] coordinate descent-based algorithm.


## sEM method

Penalizing the inverses of covariance matrices sEM algorithm [Krishnamurthy, 2011]

1. Initialize parameters $\tau_{k}^{(0)}, \mu_{k}^{(0)}$ et $\Sigma_{k}^{(0)}$
2. Expectation Step: Compute $z_{i k}^{(j)}$
3. Maximization Step: Compute $\tau_{k}^{(j+1)}, \mu_{k}^{(j+1)}$
4. $\Sigma_{k}^{(j+1)} \leftarrow \operatorname{argmax}_{\Sigma}\left(\ell-\lambda\left\|\Sigma^{-1}\right\|_{1}\right)$
5. Repeat steps $2-4$ until convergence

## sEM2 method

Penalizing the covariance matrices

- A lasso penalization on the covariance matrices is added in the maximization step:

$$
\hat{\Sigma}_{k}=\operatorname{argmax}_{\Sigma_{k}}\left(\ell-\lambda\left\|P * \Sigma_{k}\right\|_{1}\right)
$$

-     * stands for Hadamard product.
- $P$ : penalization matrix.
- The penalization parameter $\lambda$ is chosen by cross-validation.
- The penalized maximization is solved by [Wang, 2013] coordinate descent-based algorithm.


## sEM2 method

## Penalizing the covariance matrices

- Several proposed matrices $P$ :
- sEM2.1: Equal weights to all matrix elements. All elements are penalized in the same way.

$$
P_{i j}=1
$$

## sEM2 method

## Penalizing the covariance matrices

- Several proposed matrices $P$ :
- sEM2.1: Equal weights to all matrix elements. All elements are penalized in the same way.

$$
P_{i j}=1
$$

- sEM2.2: Diagonal elements are not penalized. All others are penalized equally.

$$
P_{i j}= \begin{cases}1 & \text { if } i \neq j \\ 0 & \text { if } i=j\end{cases}
$$

## sEM2 method

Penalizing the covariance matrices

- Several proposed matrices $P$ :
- sEM2.1: Equal weights to all matrix elements. All elements are penalized in the same way.

$$
P_{i j}=1
$$

- sEM2.2: Diagonal elements are not penalized. All others are penalized equally.

$$
P_{i j}= \begin{cases}1 & \text { if } i \neq j \\ 0 & \text { if } i=j\end{cases}
$$

- sEM2.3: The lower the off-diagonal element, the more penalized. Diagonal elements are not penalized.

$$
P_{i j}= \begin{cases}\frac{1}{\Sigma_{i j}} & \text { if } i \neq j \\ 0 & \text { if } i=j\end{cases}
$$

## sEM2 method

## Penalizing the covariance matrices

## sEM2 algorithm

1. Initialize parameters $\tau_{k}^{(0)}, \mu_{k}^{(0)}$ et $\Sigma_{k}^{(0)}$
2. Expectation Step: Compute $z_{i k}^{(j)}$
3. Maximization Step: Compute $\tau_{k}^{(j+1)}, \mu_{k}^{(j+1)}$
4. $\Sigma_{k}^{(j+1)} \leftarrow \operatorname{argmax}_{\Sigma}\left(\ell-\lambda\|P * \Sigma\|_{1}\right)$
5. Repeat steps $2-4$ until convergence

## Illustration on an analytical example



- 2 temporal functional random variables $f_{1}$ et $f_{2}$ depending on three random variables $a_{1}, a_{2}, a_{3}$.
- $a_{1}, a_{2}, a_{3}$ have uniform distributions.
- Let define the covariate

$$
Y\left(a_{1}, a_{2}, a_{3}\right)=\int_{0}^{1}\left(f_{1}\left(t, a_{1}, a_{2}, a_{3}\right)+f_{2}\left(t, a_{1}, a_{2}, a_{3}\right)\right) d t
$$

## Illustration on an analytical example

- Hypothesis
- Learning dataset: 600 curves
- Test dataset: 1000 curves
- SPLS decomposition + Gaussian mixture model
- Optimal number $G^{*}$ of clusters chosen with Bayesian Information Criterion (BIC)
- Proposed criteria to select the basis size $d$ and assess the quality of the characterization method:
- Criterion C1: Goodness-of-fit of estimated coefficients pdf and the real pdf with1 [Fromont et al., 2012] test.
- Criterion C2: Goodness-of-fit of estimated covariate pdf and the pdf computed with known covariates with Kolmogorov-Smirnov (KS) test.
- Criterion C3: Relative mean square between correlation on estimated functions and realizations of the variables.
$\rightarrow$ First step: use of EM algorithm


## Illustration on an analytical example

## Criterion C1: Comparison of coefficients densities



- Maximal median at $d=8$ components.
- After $d=8$, model quality decreases.


## Illustration on an analytical example

## Criterion C2: Comparison of covariates densities



- Maximal median at $d=6,8,10,18$ components.
- Low variance for $d=8$.
- Very close acceptance rates for all basis sizes.


## Illustration on an analytical example

## Criterion C3: Comparison of correlations between variables



- The correlation decrease is very even.
- Relative error is about $40 \%$ for $d=8$.


## Illustration on an analytical example: conclusions

- Based on the three criteria $\rightarrow$ optimal basis size $d^{*}=8$ :
- Good acceptance rates are obtained with EM algorithm.
- The relative errors on correlation are still quite high.
- The same criteria have been computed for other estimation algorithms $\rightarrow$ similar results obtained (same $d^{*}$ and criteria values).
- For the analytical example, the EM algorithm seems to be the best choice (efficient, easy and fast): as the number of parameters is quite low in this example ( $n=89$ for $d=8$ ), the use of sparse algorithms does not improve the estimation.
- In practice, if no test basis is available, criteria C1, C2 and C3 are computed by cross-validation.


## A nuclear safety test case (1)



## A nuclear safety test case (2)

Dataset:

- 3 functional random variables depending on time
- Scalar covariate: a safety criterion
- Learning sample: 400 samples
- Logarithmic transformation of the sample (positivity constraint)

Methodology:

- SPLS decomposition + Gaussian mixture model + EM algorithm
- Optimal $G^{*}$ determined by BIC
- Criteria C1, C2 and C3 computed by cross-validation
- Optimal $d^{*}$ chosen by the analysis of the three criteria


## A nuclear safety test case (3)

- Criterion 1:
- optimal $d^{*}=4$
- acceptance rates under $80 \%$ for $d>8$
- fast decrease of acceptance rates for $d \geqslant 10$

- Criterion 2: low acceptance rates for all basis sizes

A nuclear safety test case (4)

- Criterion 3: quite good approximation of functional variable correlations for $d^{*}=4$.





## Visualization: High Density Region boxplot (HDR)

- Proposed by [Hyndman and Shang, 2010] and based on
- Principal Component Analysis
- First two basis functions selected
- Kernel density estimation

- Application on the analytical example:
- Black curve: functional median
- Colored curves: outliers
- Dark (resp. light) gray zone: 50\% (resp. 95\%) highest density region


## Visualization: Modified HDR boxplot

- Combination of the HDR boxplot and our proposed characterization methodology (SPLS + Gaussian mixture model)
$\Rightarrow$ Simultaneous visualization of multiple functions
$\Rightarrow$ Taking into account a covariate
$\Rightarrow$ Decomposition on higher basis
- Illustration on the analytical example:




## Conclusion and perspectives

- Development of a global methodology to simultaneously characterize dependent functional random variables linked to a covariate.
$\rightarrow$ Simultaneous PLS decomposition + Gaussian mixture with sparse covariance matrices
$\Rightarrow$ Estimation of probabilities for the variables to exceed a threshold.
$\Rightarrow$ Simulation according to the estimated pdf.
$\Rightarrow$ Visualization of the uncertainty of the variables.
- Different proposed criteria to assess the methodology efficiency:
- Application on an analytical example: good results
- Application on a nuclear safety test case: functions and correlations quite well reproduced but the covariate pdf not well fitted

Perspectives:

- Computing probabilities and quantiles to exceed a threshold.
- Using this methodology to run uncertainty propagation and sensitivity analysis studies.


## References

$\square$ Dempster, A. P., Laird, N. M., and Rubin, D. B. (1977).
Maximum likelihood from incomplete data via the EM algorithm.
Journal of the Royal Statistical Society. Series B (Methodological), pages 1-38.


Friedman, J., Hastie, T., and Tibshirani, R. (2008).
Sparse inverse covariance estimation with the graphical lasso.
Biostatistics, 9(3):432-441.
Fromont, M., Laurent, B., Lerasle, M., and Reynaud-Bouret, P. (2012).
Kernels based tests with non-asymptotic bootstrap approaches for two-sample problem.
In 25th Annual Conference on Learning Theory, volume 23, pages 23.1-23.22.


Hyndman, R. J. and Shang, H. L. (2010).
Rainbow plots, bagplots, and boxplots for functional data.
Journal of Computational and Graphical Statistics, 19(1):29-45.
Krishnamurthy, A. (2011).
High-dimensional clustering with sparse gaussian mixture models.
Wang, H. (2013).
Coordinate descent algorithm for covariance graphical lasso.
Statistics and Computing, pages 1-9.
Wold, H. (1975).
Path models with latent variables: The nipals approach.

## Appendix

Analytical example definition

$$
\begin{aligned}
f_{1}\left(t, a_{1}, a_{2}, a_{3}\right) & =0.8 a_{2} B B(t)+a_{1}+c_{1}(t)+h(t) \\
f_{2}\left(t, a_{1}, a_{2}, a_{3}\right) & =a_{2} B B(t)+a_{1}+c_{2}(t)
\end{aligned}
$$

with

$$
\begin{aligned}
& a_{1} \sim \mathcal{U}(0,0.05) ; a_{2} \sim \mathcal{U}(0.05,0.2) ; a_{3} \sim \mathcal{U}(2,3) \\
& c_{1}(t)= \begin{cases}t-1 & \text { if } t<\frac{70}{512} \\
\frac{372}{512}-t & \text { otherwise }\end{cases} \\
& c_{2}(t)= \begin{cases}1-t & \text { if } t<0.5 \\
\frac{64}{5 a_{3}}-0.5 t & \text { if } 0.5<t<0.5+\frac{10 a_{3}}{512} \\
0.5-t & \text { otherwise }\end{cases} \\
& h(t)=0.15\left(1-\left|\frac{t-100 a_{3}}{60}\right|\right)
\end{aligned}
$$

