

Repulsiveness for a better integration (not my social program)

WORKSHOP ON KERNEL AND SAMPLING METHODS FOR DESIGN AND
QUANTIZATION, GDR MASCOTT-NUM

JEAN-FRANÇOIS COEURJOLLY

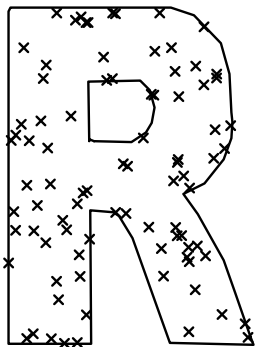
JOINT WORKS WITH A MAZOYER (IMT, TOULOUSE), PO AMBLARD (CNRS, U GRENOBLE ALPES)

November 2021



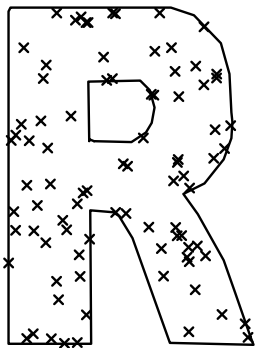
Are you sure you you're not in favor of repulsion ?

- Here are 100 points drawn randomly in ...

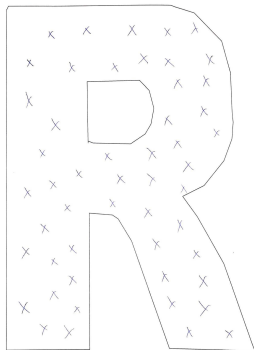


Are you sure you you're not in favor of repulsion ?

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- Then, I asked my daughter to draw a few points



Outline (no interest)

- 1 Introduction
- 2 Spatial point processes and DPPs
- 3 Main result
- 4 Simulation study
- 5 Conclusion



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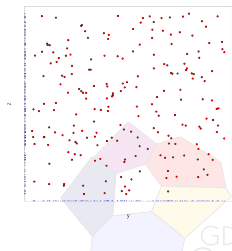
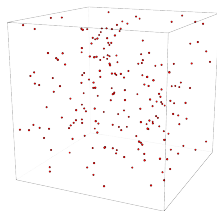
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Space-filling design

Computer experiments

- (1) The design must cover “nicely” $B = [0, 1]^d$;
 - Additional objective : study influence of a subset of inputs on output **without constructing a new design**;
- (2) Cover “nicely” the projection $B_I = [0, 1]^\iota$ where $I \subseteq \{1, \dots, d\}$ with $|I| = \iota < d$.



Spatial point processes for computer experiments

- (1) Use of *repulsive* point processes \mathbf{X} to produce the design;
- (2) Find a repulsive point process \mathbf{X} such that its projection \mathbf{X}_I onto B_I remains a repulsive point process.

Integral estimation problem “reformulation”

How to estimate $\mu(f_I) = \int_{[0,1]^I} f_I(u) du$ for any $f_I : [0,1]^I \rightarrow \mathbb{R}$, $I \subseteq \{1, \dots, d\}$ and $\iota = |I|$, using quadrature points defined in dimension d ?

That is let $U_1, \dots, U_N \in [0,1]^d$ (to be chosen). We want to estimate $\mu(f_I)$ by

$$\hat{\mu}(f_I) = \frac{1}{N} \sum_{j=1}^N f_I((U_j)_I)^a$$

where for any $u \in \mathbb{R}^d$, $u_I = (u_i)_{i \in I}$.

a. eventually weighed

- The d -dimensional design is defined once for all, and used to estimate any ι -dimensional integral.
- In this problem, the design is therefore independent of the integrand (hence a “homogeneous” design makes sense).



The simplest solution

Standard Monte-Carlo

- Let $U_1, \dots, U_N \sim \mathcal{U}([0, 1]^d)$ be iid random variables
- Estimate $\mu(f_I)$ by

$$\widehat{\mu}(f_I) = \frac{1}{N} \sum_{j=1}^N f_I((U_j)_I)$$

- Since $(U_j)_I$ are still iid uniformly distributed on $[0, 1]^t \dots$ for any $I \subseteq \{1, \dots, d\}$

$$\mathbb{E} \widehat{\mu}(f_I) = \mu(f_I) \quad \text{and} \quad \text{Var} \widehat{\mu}(f_I) = N^{-1} \left\{ \mu(f_I^2) - \mu(f_I)^2 \right\}$$

and a central limit theorem holds (and actually much more ...)



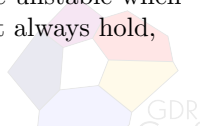
More than 70 years of research in one slide

Improvements of standard MC : huge literature (already in the situation $\iota = d$)

- Stratified Monte-Carlo methods
- MCMC, importance sampling
- Bayesian quadrature
- Quasi Monte-Carlo, Randomized Quasi Monte-Carlo methods
- + control variates, antithetic methods, variance-reduction methods
- ...

General remarks : Either methods are

- quite stable when $d \gg 1$, computationally efficient, but CLT with variance decreasing as N^{-1} ;
- or have MSE much faster ($N^{-1-2/d}$, $N^{-3-2/d+\varepsilon}$) but more unstable when $d \gg 1$, not straightforward to implement, CLT does not always hold, require strong assumptions on f ($f \in C^1$ or C^d).



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Objective

- provide a faster estimator than the standard MC one, for any f_I , under minimal assumptions on f_I (no differentiability)
- using a stochastic model ...in particular a class of repulsive spatial point processes ...and in particular a specific Determinantal Point Process.

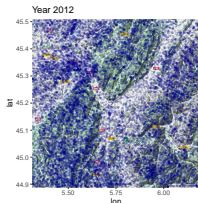
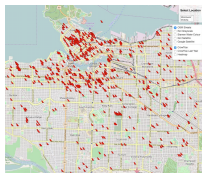
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(Continuous) SPP on \mathbb{R}^d

- Let \mathbf{X} be a spatial point process defined on $[0, 1]^d$, viewed as a locally finite random measure :
 $\mathbf{x} = \{x_1, \dots, x_m\}$, $x_j \in [0, 1]^d$



Intensity functions (the first two ones ...and informally

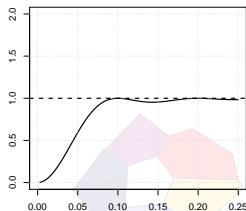
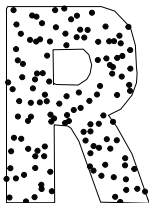
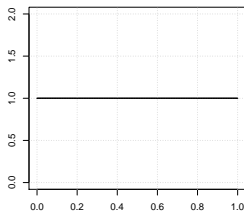
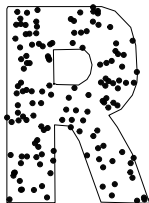
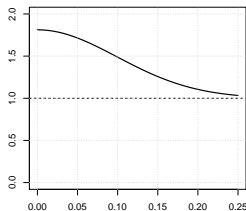
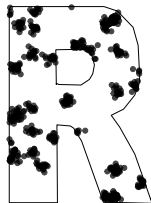
$$\rho(u) = \lim_{|du| \rightarrow 0} \frac{\mathbb{E}\{N(du)\}}{|du||dv|} \quad \text{and} \quad \rho^{(2)}(u, v) = \lim_{|du|, |dv| \rightarrow 0} \frac{\mathbb{E}\{N(du)N(dv)\}}{|du||dv|}$$

- $\rho(u)du \approx$ Prob. to observe a point in $B(u, du)$.
- $\rho^{(2)}(u, v)dudv \approx$ Prob. to observe two distinct points in $B(u, du)$ and $B(v, dv)$.
- if $\rho(\cdot) = \rho$, \mathbf{X} is said to be homogeneous ; ρ =mean number of point per unit volume.

Repulsiveness and its statistical interest

- Pair correlation function (assume $\rho(\cdot) > 0$) :

$$g(u, v) = \frac{\rho^{(2)}(u, v)}{\rho(u)\rho(v)} = \frac{\rho^{(2)}(u, v)}{\rho^2} \quad (\text{homog.}) = g_0(\|v - u\|) \quad (\text{isotr.})$$



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SPP for MC integration ($\iota = d$ for now)

Let \mathbf{X} be a SPP on $[0, 1]^d$ with intensity parameter N and pcf g

$$\widehat{\mu}(f) = N^{-1} \sum_{u \in \mathbf{X}} f(u) \quad \text{is such that } E \widehat{\mu}(f) = \mu(f)$$

and

$$\text{Var } \widehat{\mu}(f) = N^{-1} \int_{[0,1]^d} f(u)^2 du + \int_{[0,1]^d} \int_{[0,1]^d} \{g(u, v) - 1\} f(u)f(v) dudv.$$

(Continuous) Determinantal point processes

$\mathbf{X} \sim \text{DPP}_{B^d}(K)$ for some kernel $K \dots$

...if its k th order intensity ($k \geq 1$) writes

$$\rho^{(k)}(u_1, \dots, u_k) = \det \left[\left(K(u_i, u_j) \right)_{i,j=1,\dots,k} \right].$$

where K admits the Mercer decomposition $K(u, v) = \sum_{j \in \mathcal{N}_d} \lambda_j \phi_j(u) \bar{\phi}_j(v)$, where $\{\phi_j\}_j$ forms an orthonormal basis of $L^2(B^d)$; $\lambda_j \in [0, 1]$ = eigenvalues.

- Introduced by O. Macchi to model fermions ;
- Appear in the study of eigenvalues of certain random matrices, zeroes of Gaussian Analytic Functions (e.g. Permantle, Peres, Hough, Johanson, Soshnikov, ...)
- Very tractable class of models of repulsive point processes. Assume $K(u, u) = N$, then $\rho(u) = N$, and

$$g(u, v) = N^{-2} \det \begin{pmatrix} N & K(u, v) \\ K(u, v) & N \end{pmatrix} = 1 - \frac{|K(u, v)|^2}{N^2} < 1 !!$$

Why DPPs are interesting ?

Assume $K(u, u) = N$ (also valid for inhomogeneous)

$$\text{Var } \widehat{\mu}(f) = N^{-1} \int_{[0,1]^d} f(u)^2 du - \underbrace{N^{-2} \sum_{j,k \in \mathcal{N}_d} \lambda_j \lambda_k \left| \int_{[0,1]^d} f(u) du \phi_j(u) \bar{\phi}_k(u) \right|^2}_{\leq 0 \forall \text{ sign}(f)}$$

- Remark : Projection DPP, $\lambda_j \in \{0, 1\}$, in which case N is an integer.



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Bardenet and Hardy'19

- $\iota = d$, build an ad-hoc OPE (Leg. polyn.)
- **Pros** : $\text{Var } \widehat{\mu}(f) \propto N^{-1-1/d}$; CLT for $\mu(f)$;
- **Cons** : $f \in C_b^1([0, 1]^d)$ and compactly supported (due to inhomog. kernel); $\iota < d$ cannot be considered; proofs very long.

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Our approach : DPP wrt Fourier basis

- $N = n_1 \times \dots \times n_d$ and $E_N = \{1, \dots, n_1\} \times \dots \times \{1, \dots, n_d\}$ ($\#E_N = N$)
- $\phi_j(u) = e^{2i\pi j^\top u}$ for $u \in [0, 1]^d$
- then we consider the (**homogeneous**) kernel K defined by

$$K(u, v) = \sum_{j \in E_N} e^{2i\pi j^\top (u-v)} = \prod_{\ell=1}^d K_\ell(u_\ell, v_\ell) \quad \text{where} \quad K_\ell(u_\ell, v_\ell) = \underbrace{\sum_{j=1}^{n_\ell} e^{2i\pi j(u_\ell - v_\ell)}}_{\propto \text{Dirichlet kernel in Sign. Proc.}}$$



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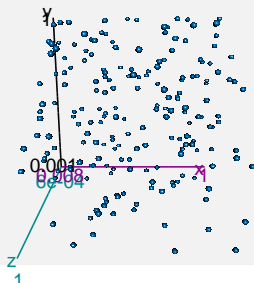
Theorem (MCA'19) : Why taking E_N as a rectangular set ?

Let $\mathbf{X} \sim \text{DPP}_{[0,1]^d}(K)$ and let $I \subseteq \{1, \dots, d\}$, then

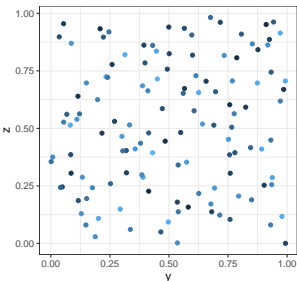
$$\mathbf{X}_I \sim (-1/N_{I^c}) - \text{DPP}_{B^I}(N_{I^c} K_I)$$

where $N = N_I \times N_{I^c}$ and $K_I = \prod_{\ell \in I} K_\ell$. In particular, $g_{\mathbf{X}_I} < 1, \forall I!!$

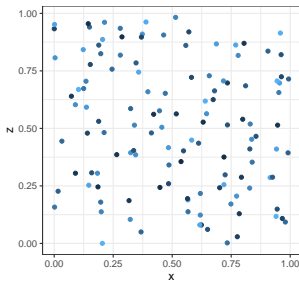
An example of a 3d-realization



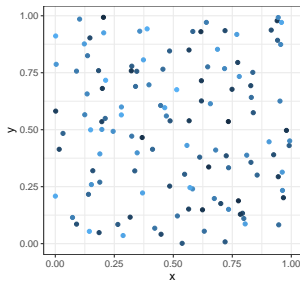
Projection on (y,z)



Projection on (x,z)



Projection on (x,y)



Our main result

- Fourier coefficient : $\hat{f}_I(k) = \langle f_I, \phi_k \rangle$
- Assume $f_I \in \mathcal{H}^s([0, 1]^d) = \{f_I \in L^2 : \sum_{k \in \mathbb{Z}^d} (1 + \|k\|_\infty)^{2s} |\hat{f}(k)|^2 < \infty\}$.

Theorem (CMA'19+)

Let $d > 1$ and assume $N = n_1 \times \dots \times n_d$ s.t. $n_\ell N^{-1/d} \rightarrow 1$

(i) If $s > 1/2$ and $\|f_I\|_\infty < \infty$ then

$$\text{Var}(\widehat{\mu}(f_I)) \sim N^{-1-1/d} \sigma^2(f_I) \quad \text{with} \quad \sigma^2(f_I) = \sum_{k \in \mathbb{Z}^d} \|k\|_1 |\hat{f}_I(k)|^2.$$

$$\sqrt{N^{1+1/d}} (\widehat{\mu}(f_I) - \mu(f_I)) \xrightarrow{d} N(0, \sigma^2(f_I)).$$

(ii) If $0 < s < 1/2$, then

$$\text{Var}(\widehat{\mu}(f_I)) = \mathcal{O}(N^{-1-\frac{2s}{d}}) = o(N^{-1})$$

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Setting

Estimate $\mu(f)$ for different d

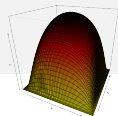
- ① For $N = 100, \dots, 1000$ and $d = 1, \dots, 6$: 2500 replications of a Dirichlet DPP producing N points in $[0, 1]^d$;
- ② For each d and some given f :
 - ① Estimate $\mu(f)$ for each N using the design constructed in dimension d ;
 - ② Linear regression of empirical variance vs. $\log(N)$ and estimate the slope.

Estimate $\mu(f_I)$ for different ι

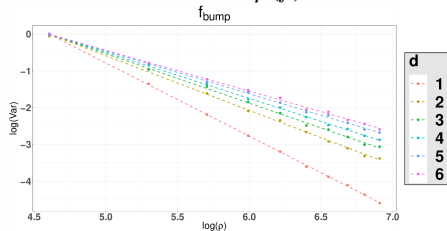
- ① For $N = 100, \dots, 1000$: 2500 replications of a Dirichlet DPP producing N points in $[0, 1]^6$;
- ② For $\iota \in \{6, \dots, 1\}$:
 - ① For a ι -dimensional function f_I , estimate $\mu(f_I)$ for each N using the design constructed in dimension 6 ;
 - ② Linear regression of empirical vs. $\log(N)$ and estimate the slope.

Bump : $f \in \mathcal{H}^s$, for any $s > 0$

$$f(x) \propto \prod_{i=1}^d \exp\left(-\frac{0.1}{.25-(x_i-.5)^2}\right)$$



Estimation of $\mu(f)$ via \mathbf{X}

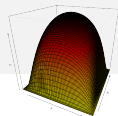


d	Conf. interval	Estimation	-1-1/d
1	[-2.016 ; -1.966]	-1.991	-2.000
2	[-1.536 ; -1.438]	-1.487	-1.500
3	[-1.391 ; -1.302]	-1.347	-1.333
4	[-1.282 ; -1.225]	-1.254	-1.250
5	[-1.2 ; -1.129]	-1.164	-1.200
6	[-1.153 ; -1.073]	-1.113	-1.167

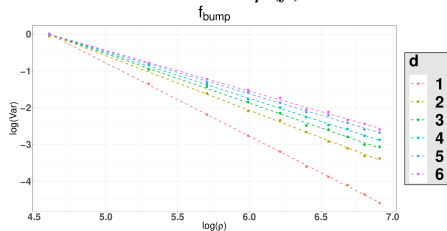


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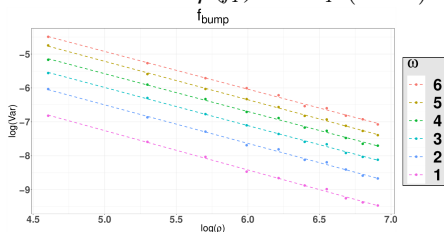


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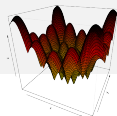
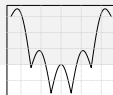
Estimation of $\mu(f_I)$ via \mathbf{X}_I ($d = 6$)



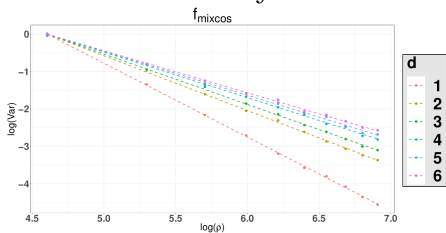
ω	Conf. interval	Estimation	$-1-1/d$
6	[-1.153 ; -1.073]	-1.113	-1.167
5	[-1.162 ; -1.104]	-1.133	-1.167
4	[-1.157 ; -1.077]	-1.117	-1.167
3	[-1.16 ; -1.09]	-1.125	-1.167
2	[-1.186 ; -1.087]	-1.136	-1.167
1	[-1.204 ; -1.116]	-1.160	-1.167

Mix-cos : $f \in \mathcal{H}^{3/2}$

$$f(x) \propto \sum_{i=1}^d \{0.1|\cos(5\pi(x_i - 1/2))| + (x_i - 1/2)^2\}$$

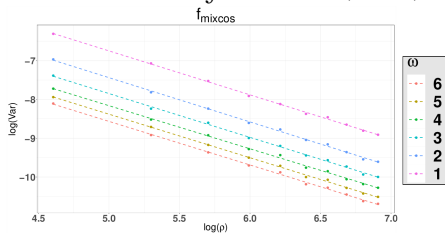


Estimation of $\int f_d$ via \mathbf{X}



d	Conf. interval	Estimation	-1-1/d
1	[-2.003 ; -1.942]	-1.972	-2.000
2	[-1.5 ; -1.429]	-1.464	-1.500
3	[-1.384 ; -1.31]	-1.347	-1.333
4	[-1.253 ; -1.183]	-1.218	-1.250
5	[-1.219 ; -1.123]	-1.171	-1.200
6	[-1.159 ; -1.089]	-1.124	-1.167

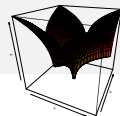
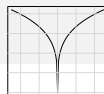
Estimation of $\int f_I$ via \mathbf{X}_I ($d = 6$)



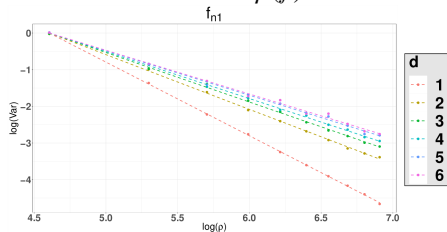
ω	Conf. interval	Estimation	-1-1/d
6	[-1.159 ; -1.089]	-1.124	-1.167
5	[-1.154 ; -1.091]	-1.123	-1.167
4	[-1.151 ; -1.072]	-1.111	-1.167
3	[-1.168 ; -1.092]	-1.130	-1.167
2	[-1.18 ; -1.102]	-1.141	-1.167
1	[-1.163 ; -1.104]	-1.134	-1.167

“ L^{γ} -norm” : $f(x) \propto \sum_{i=1}^d |u_i - 1/2|^{1/4}$,

$f \in H^s$ for any $s < 3/4$

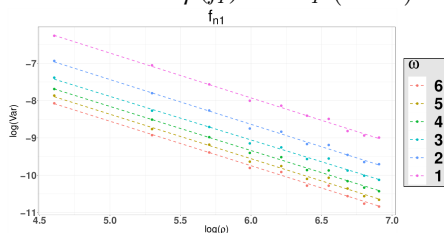


Estimation of $\mu(f)$ via \mathbf{X}



d	Conf. interval	Estimation	-1-1/d
1	[-2.033 ; -1.986]	-2.009	-2.000
2	[-1.529 ; -1.458]	-1.493	-1.500
3	[-1.393 ; -1.306]	-1.349	-1.333
4	[-1.318 ; -1.246]	-1.282	-1.250
5	[-1.272 ; -1.161]	-1.217	-1.200
6	[-1.262 ; -1.123]	-1.192	-1.167

Estimation $\mu(f_I)$ via \mathbf{X}_I ($d = 6$)



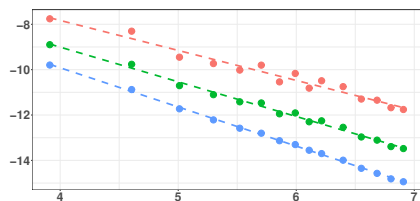
ω	Conf. interval	Estimation	-1-1/d
6	[-1.262 ; -1.123]	-1.192	-1.167
5	[-1.269 ; -1.114]	-1.191	-1.167
4	[-1.248 ; -1.118]	-1.183	-1.167
3	[-1.243 ; -1.094]	-1.168	-1.167
2	[-1.272 ; -1.129]	-1.200	-1.167
1	[-1.263 ; -1.14]	-1.201	-1.167

A last example in the situation \mathcal{H}^s with $s < 1/2$

- Particular case $d = 1$, $B = [0, 1]$,

$$h_\gamma(x) = \sum_{j \geq 1} \frac{\cos(2\pi jx)}{2\pi j^\gamma}, \quad \gamma > 1/2.$$

Then $h_\gamma \in \mathcal{H}^s$ for any $0 < s < \gamma - 1/2$



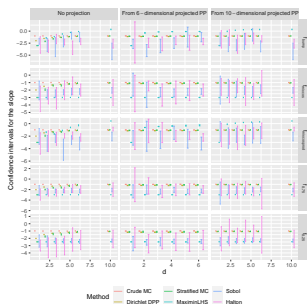
γ — 0.625 — 0.75 — 0.875

γ	s	$-1-2s/d$	Estimation	Conf. interval
0.625	0.125	-1.25	-1.32	[-1.48 ; -1.17]
0.750	0.250	-1.50	-1.53	[-1.61 ; -1.45]
0.875	0.375	-1.75	-1.73	[-1.76 ; -1.69]



More simulations ...

- Simulations were made for $d = 6$ or 10 and any $\iota = 1, \dots, d$ for additional functions, and for our DirDPP design and for : BH-DPP [BH'19], crudeMC, stratified MC, Sobol, MaximinLHS, Halton designs.



Just impossible to sum up >20 figures (please refer to [EJS'21](#)) ...but ...

- for a specific function, sample size, dimension : there were **as expected** better methods (and actually much faster)
- DirDPP was the only one to remain **stable** in terms of behaviour for all considered functions, sample size, dimension d and for any $\iota = 1, \dots, d$.

Outline (no interest)

- 1 Introduction
- 2 Spatial point processes and DPPs
- 3 Main result
- 4 Simulation study
- 5 Conclusion**



Summary :

- A simple MC estimator based on a homogeneous projection DPP.
- Universal result : unbiased est. and CLT with rate $\sqrt{N^{1+1/d}}$
- Results available
 - for $f \in \mathcal{H}^{1/2}$, i.e. for some non-differentiable functions
 - on lower dimensional spaces



Summary :

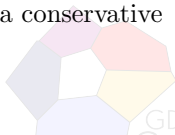
- A simple MC estimator based on a homogeneous projection DPP.
- Universal result : unbiased est. and CLT with rate $\sqrt{N^{1+1/d}}$
- Results available
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 - on lower dimensional spaces

What I didn't say :

- Asymptotic normality “checked” (qqplots and Shapiro-Wilk test).
- Asympt. CI for $\mu(f_I)$? Interesting fact when $f_I \in C_b^1([0, 1]^t)$

$$\sigma^2(f_I) = \sum_{j \in Z^t} \|j\|_1 |\hat{f}(j)|^2 \leq 4\pi^2 \int_{[0,1]^t} \|\nabla f_I(u)\|^2 du =: \mathcal{J}_I$$

$\Rightarrow \mathcal{J}_I$ is easily estimated (using again \mathbf{X}_I !!), leading to a conservative asympt. CI.



Where to go ? (just a few perspectives)



- for the same problem : other projection kernels, infinite integrals
- for a given f : inhomogeneous kernels to exploit f
- lots of possible statistical applications (density estimation, regression,...) when a random design is involved

- d -dimensional sphere, torus, ...?
- simulations= $\mathcal{O}(dN^3)$ (for now), not sequential but code and replications are available :
<https://github.com/AdriMaz>



Tricky Fourier basis

Just to give an idea, let's consider $d = 1$

$$\begin{aligned}
 \text{Var} \widehat{I}_1(f_1) &= \frac{1}{N} \sum_{j \in \mathbb{Z}} |\langle f, \phi_j \rangle|^2 - \frac{1}{N^2} \sum_{j,k=1}^N |\langle f, \phi_j \bar{\phi}_k \rangle|^2 \\
 &= \frac{1}{N} \sum_{j \in \mathbb{Z}} |\langle f, \phi_j \rangle|^2 - \frac{1}{N^2} \sum_{j,k=1}^N |\langle f, \phi_{j-k} \rangle|^2 \\
 &= \frac{1}{N} \sum_{j \in \mathbb{Z}} |\hat{f}(j)|^2 - \frac{1}{N^2} \sum_{|j| \leq N} (N - |j|) |\hat{f}(j)|^2 \\
 &= \frac{1}{N} \sum_{|j| \geq N} |\hat{f}(j)|^2 + \frac{1}{N^2} \sum_{|j| \leq N} |j| |\hat{f}(j)|^2 \\
 &\sim \frac{1}{N^2} \sum_{j \in \mathbb{Z}} |j| |\hat{f}(j)|^2
 \end{aligned}$$

