

The complexity of optimizing noisy functions on graphs

Mascotnum Workshop on "stochastic simulators"

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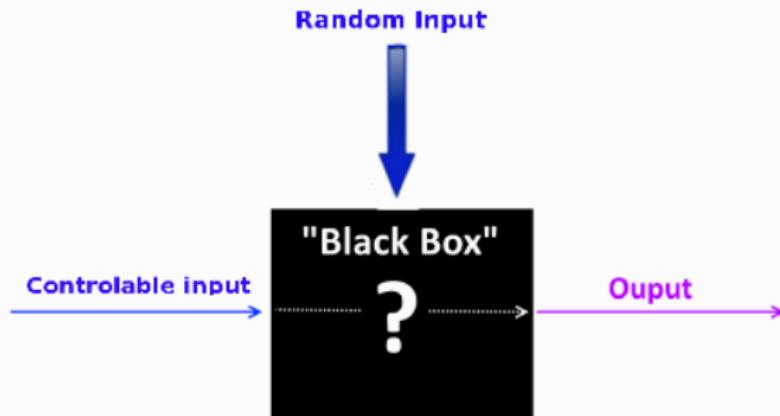
Warm-up: 1.5 arms

Discrete, Non-smooth Functions – The Vanilla Bandit Model

Graph Structure and Smooth Functions

PAC Optimization

Random Experiment



Random black box

$$F : \mathcal{X} \times \Omega \rightarrow \mathbb{R}$$

where \mathcal{X} = space \mathcal{X} of controlled variables and Ω = random space

Target function $f : \mathcal{X} \rightarrow \mathbb{R}$ $f(x) = \mathbb{E}[F(x, \cdot)]$ or some other functionnal of $F(x, \cdot)$

Typically (in the sequel): $F(x, \omega) = f(x) + \epsilon(\omega)$ where $\epsilon \sim \mathcal{N}(0, \sigma^2)$.

Black Box Optimization

Black-box interaction model:

- choose $X_1 = \phi_1(U_1)$, observe $Y_1 = F(X_1, \omega_1)$
- choose $X_2 = \phi_2(X_1, Y_1, U_2)$ observe $Y_2 = F(X_2, \omega_2)$,
- etc...

Target = optimize f : $f^* = \max_{\mathcal{X}} f$ (or min, or find level set, etc.)

Strategy: sampling rule $(\phi_t)_{t \geq 1}$, stopping time τ

PAC setting: for a risk δ and a tolerance ϵ , return $X_{\tau+1}$ such that

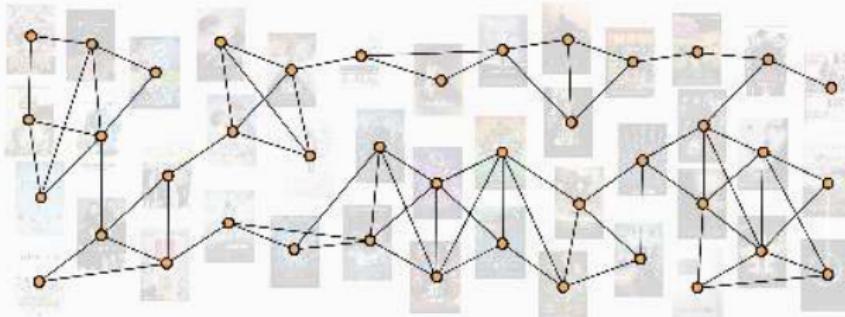
$$\mathbb{P}_f(f(X_{\tau+1}) < f^* - \epsilon) \leq \delta .$$

The set of correct answers is $\mathcal{X}_\epsilon(f) = \{x \in X : f(x) \geq f^* - \epsilon\}$.

Goal: find a strategy minimizing $\mathbb{E}_f[\tau]$ for all possible functions f
(no sub-optimal multiplicative constant!)

Optimizing on a Graph

Movie similarity graph:



Warm-up: 1.5 arms

A Simplistic and yet Interesting Example

$\mathcal{X} = \{0, 1\}$, $f = (0, \mu)$ ie. $F(0, \cdot) = 0, F(1, \cdot) \sim \mathcal{N}(\mu, \sigma^2)$

Here, $X_t = 1$ for $t \leq \tau$ and $Y_t \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2)$.

Equivalent to testing the overlapping hypotheses: $\mu \leq \epsilon$ vs $\mu \geq -\epsilon$.

Theorem

For every (ϵ, δ) -PAC strategy,

$$\liminf_{\delta \rightarrow 0} \frac{\mathbb{E}_\mu[\tau_\delta]}{\log(1/\delta)} \geq \frac{2\sigma^2}{(|\mu| + \epsilon)^2}.$$

Besides, the choice

$$\tau_\delta = \inf \left\{ t \in \mathbb{N} : \frac{t(|\hat{\mu}_t| + \epsilon)^2}{2\sigma^2} > 3 \ln(\ln(t) + 1) + \mathcal{T}(\ln(1/\delta)) \right\},$$

where $\mathcal{T}(x) \simeq x + 4 \ln(1 + x + \sqrt{2x})$ and $X_{\tau+1} = \mathbb{1}\{\hat{\mu}_{\tau_\delta} > 0\}$ is such that $\forall \mu \in \mathbb{R}$

$$\limsup_{\delta \rightarrow 0} \frac{\mathbb{E}_\mu[\tau_\delta]}{\ln(1/\delta)} \leq \frac{2\sigma^2}{(|\mu| + \epsilon)^2}.$$

Lower Bounds: Information Inequalities

Idea: Cannot stop as long as an alternative hypothesis is likely to produce similar data.

Alternative hypotheses: $\text{Alt}(f) = \{g : \mathcal{X} \rightarrow \mathbb{R} \text{ such that } \mathcal{X}_\epsilon(f) \not\subset \mathcal{X}_\epsilon(g)\}$
and not $\mathcal{X}_\epsilon(f) \cap \mathcal{X}_\epsilon(g) = \emptyset$: just "one of the correct answers for f is incorrect for g ".

Here,

- if $\mu > \epsilon$, $\text{Alt}(\mu) = (-\infty, -\epsilon)$
- if $\mu < -\epsilon$, $\text{Alt}(\mu) = (\epsilon, +\infty)$
- if $-\epsilon \leq \mu \leq \epsilon$, $\text{Alt}(\mu) = (-\infty, -\epsilon) \cup (\epsilon, +\infty)$

Change-of-Measure Lemma

For a given strategy, likelihood of the observation under function f :
 $\ell(Y_1, \dots, Y_t; f)$.

$$\text{Likelihood ratio } L_t(f, g) := \log \frac{\ell(Y_1, \dots, Y_t; f)}{\ell(Y_1, \dots, Y_t; g)}.$$

Lemma

Let f and g be two functions.

1. **Low-level form:** for all $x \in \mathbb{R}$, $n \in \mathbb{N}^*$, for every event $C \in \mathcal{F}_n$,

$$\mathbb{P}_g(C) \geq e^{-x} [\mathbb{P}_f(C) - \mathbb{P}_f(L_n(f, g) \geq x)].$$

2. **High-level form:** for any stopping time τ and any event $C \in \mathcal{F}_\tau$,

$$\mathbb{E}_f [L_\tau(f, g)] = \text{KL} \left(\mathbb{P}_f^{X_1, \dots, X_\tau}, \mathbb{P}_g^{X_1, \dots, X_\tau} \right) \geq \text{kl}(\mathbb{P}_f(C), \mathbb{P}_g(C)),$$

where $\text{kl}(x, y) = x \ln(x/y) + (1-x) \ln((1-x)/(1-y))$ is the binary relative entropy.

Proofs: Change-of-Measure Lemma

- Low-level form:

$$\begin{aligned}\mathbb{P}_g(C) &= \mathbb{E}_f \left[\mathbb{1}_C \exp(-L_t(f, g)) \right] \geq \mathbb{E}_f \left[\mathbb{1}_C \mathbb{1}_{(L_t(f, g) < x)} e^{-L_t(f, g)} \right] \\ &\geq e^{-x} \mathbb{P}_f \left(C \cap (L_t(f, g) < x) \right) \\ &\geq e^{-x} [\mathbb{P}_f(C) - \mathbb{P}_f(L_n(f, g) \geq x)].\end{aligned}$$

- High-level form: Information Theory (data-processing inequality)

$$\begin{aligned}\mathbb{E}_\mu [L_\tau(f, g)] &= \text{KL} \left(\mathbb{P}_f^{X_1, \dots, X_\tau}, \mathbb{P}_g^{X_1, \dots, X_\tau} \right) \\ &\geq \text{KL} \left(\mathbb{P}_f^{\mathbb{1}_C}, \mathbb{P}_g^{\mathbb{1}_C} \right) = \text{kl}(\mathbb{P}_f(C), \mathbb{P}_g(C)).\end{aligned}$$

1.5 arm: Lower Bound for $|\mu| > \epsilon$

$\mathbb{E}_\mu[\tau] < \infty$ and Wald's lemma yield:

$$\mathbb{E}_\mu[L_\tau(\mu, \lambda)] = \mathbb{E}_\mu[\tau] \text{KL}(\mathcal{N}(\mu, \sigma^2), \mathcal{N}(\lambda, \sigma^2)) = \mathbb{E}_\mu[\tau] \frac{(\mu - \lambda)^2}{2\sigma^2} .$$

Hence, by the high-level form of the lemma:

$$\mathbb{E}_\mu[\tau] \frac{(\mu - \lambda)^2}{2\sigma^2} \geq \text{kl}(\mathbb{P}_\mu(C), \mathbb{P}_\lambda(C)) .$$

For $\mu < -\epsilon$, choosing $\lambda = \epsilon$ and $C = \{X_{\tau+1} = 1\}$, which is such that $\mathbb{P}_\mu(C) \leq \delta$ and $\mathbb{P}_\lambda(C) \geq 1 - \delta$, directly yields

$$\mathbb{E}_\mu[\tau] \frac{(|\mu| + \epsilon)^2}{2\sigma^2} \geq \text{kl}(\delta, 1 - \delta) \approx \log \frac{1}{\delta} .$$

Similarly, for $\mu > \epsilon$, we use $\lambda = -\epsilon$ and $C = (X_{\tau+1} = 0)$ so as to obtain the same inequality.

⇒ non-asymptotic lower bound

1.5 arm: lower bound for $-\epsilon \leq \mu \leq \epsilon$

For a fixed $\eta > 0$. Introducing

$$n_\delta := \left\lceil \frac{2\sigma^2(1-\eta)}{(|\mu| + \epsilon)^2} \ln \frac{1}{\delta} \right\rceil ,$$

and the event $C_\delta = \{\tau_\delta \leq n_\delta\}$, we prove that

$$\mathbb{P}_\mu (\tau_\delta \leq n_\delta) = \mathbb{P}_\mu (C_\delta, X_{\tau+1} = 0) + \mathbb{P}_\mu (C_\delta, X_{\tau+1} = 1) \rightarrow 0 \text{ when } \delta \rightarrow 0 .$$

Choosing $\lambda = \epsilon$ yields that $\mathbb{P}_\lambda (C_\delta, X_{\tau+1} = 0) \leq \mathbb{P}_\lambda (X_{\tau+1} = 0) \leq \delta$. The event $C_\delta \cap \{X_{\tau+1} = 0\}$ belongs to F_{n_δ} . Hence, for all $x \in \mathbb{R}$,

$$\delta \geq e^{-x} [\mathbb{P}_\mu (C_\delta, X_{\tau+1} = 0) - \mathbb{P}_\mu (L_{n_\delta}(\mu, \epsilon) \geq x)] ,$$

which can be rewritten as

$$\mathbb{P}_\mu (C_\delta, X_{\tau+1} = 0) \leq \delta e^x + \mathbb{P}_\mu (L_{n_\delta}(\mu, \epsilon) \geq x) .$$

The choice $x = (1 - \eta/2) \ln(1/\delta)$ yields

$$\mathbb{P}_\mu (C_\delta, X_{\tau+1} = 0) \leq \delta^{\frac{\eta}{2}} + \mathbb{P}_\mu \left(\frac{L_{n_\delta}(\mu, \epsilon)}{n_\delta} \geq \frac{1 - \eta/2}{1 - \eta} \frac{(|\mu| + \epsilon)^2}{2\sigma^2} \right) .$$

By the law of large numbers and the fact that $n_\delta \rightarrow +\infty$ when $\delta \rightarrow 0$,

$$\frac{L_{n_\delta}(\mu, \epsilon)}{n_\delta} \xrightarrow{\delta \rightarrow 0} \mathbb{E}_\mu \left[\ln \frac{\ell(X_1; \mu)}{\ell(X_1; \epsilon)} \right] = \text{KL} \left(\mathbb{P}_\mu^{X_1}, \mathbb{P}_\epsilon^{X_1} \right) = \frac{(\mu - \epsilon)^2}{2\sigma^2} < \frac{1 - \eta/2}{1 - \eta} \frac{(|\mu| + \epsilon)^2}{2\sigma^2} .$$

Strategy: GLRT

Idea: cannot decide as long as incompatible hypotheses remain δ -likely.

Generalized Likelihood Ratio Test:

$$\begin{aligned} \ln \frac{\max_{\mu \in \mathcal{R}} \ell(X_1, \dots, X_t; \mu)}{\max_{\lambda \in \text{Alt}(\hat{f}_t)} \ell(X_1, \dots, X_t; \lambda)} &= \ln \inf_{\lambda \in \text{Alt}(\hat{\mu}_t)} \frac{\ell(X_1, \dots, X_t; \hat{\mu}(t))}{\ell(X_1, \dots, X_t; \lambda)} \\ &= \inf_{\lambda \in \text{Alt}(\hat{\mu}_t)} \frac{t(\hat{\mu}_t - \lambda)^2}{2\sigma^2} \end{aligned}$$

where $\hat{\mu}(t) = \frac{1}{s} \sum_{i=1}^t X_i$. Here:

$$\begin{aligned} \tau_\delta &= \inf \left\{ t \in \mathbb{N} : \lambda \in \max_{\text{Alt}(\hat{\mu}_t)} \ln \frac{\ell(X_1, \dots, X_t; \hat{\mu}(t))}{\ell(X_1, \dots, X_t; \lambda)} > \beta(t, \delta) \right\} \\ &= \inf \left\{ t \in \mathbb{N} : \frac{t(|\hat{\mu}_t| + \epsilon)^2}{2\sigma^2} > \beta(t, \delta) \right\} \end{aligned}$$

and

$$X_{\tau+1} = 1\{\hat{\mu}_{\tau_\delta} > 0\}$$

Efficiency of the Sequential GLRT Procedure

Fix $\mu \in \mathbb{R}$ and let $\alpha \in [0, \epsilon)$.

$$\begin{aligned} \mathbb{E}[\tau_\delta] &\leq \sum_{t=1}^{\infty} \mathbb{P}\left(t(|\hat{\mu}_t| + \epsilon)^2 \leq 2\sigma^2 \beta(t, \delta)\right) \\ &\leq \sum_{t=1}^{\infty} \mathbb{P}(|\hat{\mu}_t - \mu| > \alpha) + \sum_{t=1}^{\infty} \mathbb{P}\left(t(|\hat{\mu}_t| + \epsilon)^2 \leq 2\sigma^2 \beta(t, \delta), |\hat{\mu}_t - \mu| \leq \alpha\right) \\ &\leq \sum_{t=1}^{\infty} \mathbb{P}(|\hat{\mu}_t - \mu| > \alpha) + \sum_{t=1}^{\infty} \mathbb{P}\left(t(|\mu| - \alpha + \epsilon)^2 \leq 2\sigma^2 \beta(t, \delta), |\hat{\mu}_t - \mu| \leq \alpha\right). \end{aligned}$$

The first term is upper bounded by a constant (independent of δ), while the second is upper bounded by

$$T_0(\delta) = \inf \left\{ T \in \mathbb{N}^* : \forall t \geq T, t(|\mu| - \alpha + \epsilon)^2 \leq 2\sigma^2 \beta(t, \delta) \right\}.$$

For $\beta(t, \delta) = 3 \ln(\ln(t) + 1) + \mathcal{T}(\ln(1/\delta))$, since for every $\alpha \geq 0$, $\gamma \geq 1 + \alpha$ and $t > 0$,

$$t \geq \gamma + 2\alpha \ln(\gamma) \Rightarrow t \geq \gamma + \alpha \ln(t),$$

we have

$$T_0(\delta) = \frac{2\sigma^2}{(|\mu| - \alpha + \epsilon)^2} \ln \frac{1}{\delta} + o_{\delta \rightarrow 0} \left(\ln \frac{1}{\delta} \right).$$

Letting α go to zero, one obtains, for all $\mu \in \mathbb{R}$,

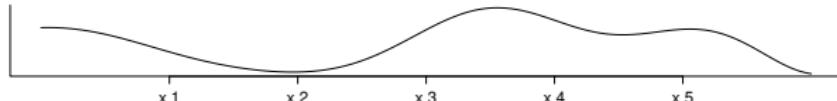
$$\limsup_{\delta \rightarrow 0} \frac{\mathbb{E}_\mu[\tau_\delta]}{\ln(1/\delta)} \leq \frac{2\sigma^2}{(|\mu| + \epsilon)^2}.$$

Discrete, Non-smooth Functions

– The Vanilla Bandit Model

Best-Arm Identification with Fixed Confidence

K options = probability distributions $\nu = (\nu_a)_{1 \leq a \leq K}$
 $\nu_a \in \mathcal{F}$ exponential family parameterized by its expectation μ_a



At round t , you may:

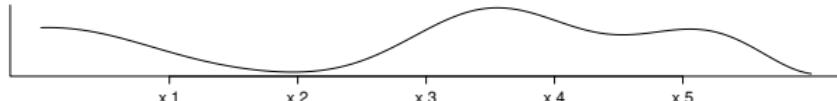
- choose an option $A_t = \phi_t(A_1, X_1, \dots, A_{t-1}, X_{t-1}) \in \{1, \dots, K\}$
- observe a new independent sample $X_t \sim \nu_{A_t}$

so as to identify the best option $a^* = \operatorname{argmax}_a \mu_a$ and $\mu^* = \max_a \mu_a$
as fast as possible: stopping time τ .

Fixed-budget setting	Fixed-confidence setting
given $\tau = T$ minimize $\mathbb{P}(\hat{a}_\tau \neq a^*)$	minimize $\mathbb{E}[\tau]$ under constraint $\mathbb{P}(\hat{a}_\tau \neq a^*) \leq \delta$

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as fast as possible: stopping time τ_δ .

Fixed-budget setting	Fixed-confidence setting
given $\tau = T$ minimize $\mathbb{P}(\hat{a}_\tau \neq a^*)$	minimize $\mathbb{E}[\tau_\delta]$ under constraint $\mathbb{P}(\hat{a}_\tau \neq a^*) \leq \delta$

Racing Strategy for $\epsilon = 0$, see [Kaufmann & Kalyanakrishnan '13]

$\mathcal{R} := \{1, \dots, K\}$ set of remaining arms.

$r := 0$ current round

while $|\mathcal{R}| > 1$

- $r := r + 1$
- draw each $x \in \mathcal{R}$, compute $\hat{f}_r(x)$, the empirical mean of the r samples observed so far
- compute the **empirical best** and **empirical worst** arms:

$$b_r = \operatorname{argmax}_{x \in \mathcal{R}} \hat{f}_r(x) \quad w_r = \operatorname{argmin}_{x \in \mathcal{R}} \hat{f}_r(x)$$

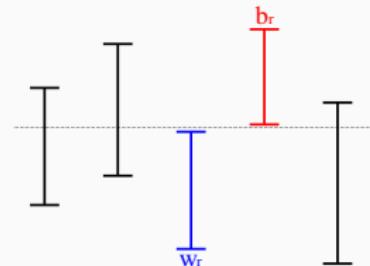
as well as an ucb for w_r and an lcb for b_r .

- Elimination step: if

$$\ell_{b_r}(r) > u_{w_r}(r),$$

then eliminate w_r : $\mathcal{R} := \mathcal{R} \setminus \{w_r\}$

end



Output: \hat{x} the single element remaining in \mathcal{R} .

The case $\epsilon = 0$

Theorem [G. and Kaufmann, 2016]

For any δ -correct algorithm,

$$\mathbb{E}_{\mu}[\tau] \geq T^*(\mu) \text{kl}(\delta, 1 - \delta),$$

where

$$\begin{aligned} T^*(\mu)^{-1} &= \sup_{w \in \Sigma_K} \inf_{g \in \text{Alt}(f)} \sum_{x \in \mathcal{X}} w(x) \text{kl}(f(x), g(x)) \\ &= \max_{w \in \Sigma_K} \min_{y \neq x^*(f)} \inf_{f(y) \leq \lambda \leq f^*} w(x^*(f)) \text{kl}(f^*, \lambda) + w(y) \text{kl}(f(y), \lambda). \end{aligned}$$

- A kind of **game** : you choose the proportions of draws $w \in \Sigma_K$, the opponent chooses the alternative confusing function.
- the **optimal proportions of arm draws** are

$$w^{*,f} = \operatorname{argmax}_{w \in \Sigma_K} \inf_{g \in \text{Alt}(f)} \sum_{x \in \mathcal{X}} w(x) \text{kl}(f(x), g(x)).$$

Entropic Lower Bound (high-level form)

Assume that f and g have a different maximum $x^*(f) \neq x^*(g)$. Then

$$\begin{aligned} \sum_{x \in \mathcal{X}} \mathbb{E}_f[N_x(\tau)] \text{kl}(f(x), g(x)) &= \text{KL}\left(\mathbb{P}_f^{(X_1, \dots, X_\tau)}, \mathbb{P}_g^{(X_1, \dots, X_\tau)}\right) \\ &\geq \text{KL}\left(\mathbb{P}_f^{\mathbb{1}\{\hat{x}_\tau = x^*(f)\}}, \mathbb{P}_g^{\mathbb{1}\{\hat{x}_\tau = x^*(f)\}}\right) \\ &\geq \text{kl}\left(\mathbb{P}_f(\hat{x}_\tau = x^*(f)), \mathbb{P}_g(\hat{x}_\tau = x^*(f))\right) \\ &\geq \text{kl}(1 - \delta, \delta) . \end{aligned}$$

Entropic Lower Bound

[Kaufmann, Cappé, G.'15],[G., Ménard, Stoltz '16]

For every δ -correct procedure, if $x^*(f) \neq x^*(g)$ then

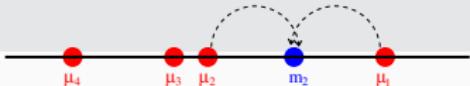
$$\sum_{x \in \mathcal{X}} \mathbb{E}_f[N_x(\tau)] \text{kl}(f(x), g(x)) \geq \text{kl}(1 - \delta, \delta) .$$

Combining the Entropic Lower Bounds

Entropic Lower Bound

If $x^*(f) \neq x^*(g)$, any δ -correct algorithm satisfies

$$\sum_{x \in \mathcal{X}} \mathbb{E}_f[N_x(\tau)] \text{kl}(f(x), g(x)) \geq \text{kl}(\delta, 1 - \delta).$$



Let $\text{Alt}(f) = \{g : x^*(g) \neq x^*(f)\}$.

$$\inf_{g \in \text{Alt}(f)} \sum_{x \in \mathcal{X}} \mathbb{E}_\mu[N_x(\tau)] \text{kl}(f(x), g(x)) \geq \text{kl}(\delta, 1 - \delta)$$

$$\mathbb{E}_f[\tau] \times \inf_{g \in \text{Alt}(f)} \sum_{x \in \mathcal{X}} \frac{\mathbb{E}_f[N_x(\tau)]}{\mathbb{E}_f[\tau]} \text{kl}(f(x), g(x)) \geq \text{kl}(\delta, 1 - \delta)$$

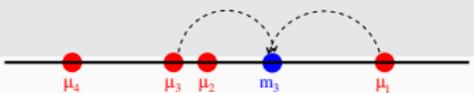
$$\mathbb{E}_f[\tau] \times \left(\sup_{w \in \Sigma_K} \inf_{g \in \text{Alt}(f)} \sum_{x \in \mathcal{X}} w(x) \text{kl}(f(x), g(x)) \right) \geq \text{kl}(\delta, 1 - \delta)$$

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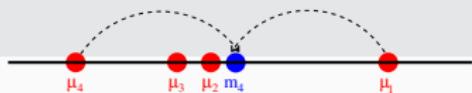
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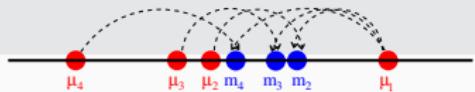
$$\mathbb{E}_f[\tau] \times \left(\sup_{w \in \Sigma_K} \inf_{g \in \text{Alt}(f)} \sum_{x \in \mathcal{X}} w(x) \text{kl}(f(x), g(x)) \right) \geq \text{kl}(\delta, 1 - \delta)$$

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Information Complexity of PAC Optimization

Theorem

For any ϵ -PAC family of converging strategies,

$$\liminf_{\delta \rightarrow 0} \frac{\mathbb{E}_f[\tau_\delta]}{\ln(1/\delta)} \geq T_\epsilon^*(f) = \sup_{w \in \Sigma_K} \max_{x \in \mathcal{X}_\epsilon(f)} \min_{y \neq x} \inf_{\substack{(\lambda_x, \lambda_y): \\ \lambda_x \leq \lambda_y - \epsilon}} w(x) \frac{f(x) - g(x)}{2} + w(y) \frac{(f(y) - g(y))}{2},$$

where $\Sigma_K = \{w \in [0, 1]^{\mathcal{X}} : \sum_x w(x) = 1\}$.

- Two armed-case: $T_\epsilon^*(f) = \frac{8\sigma^2}{(\text{osc}(f)+\epsilon)^2}$.
- Approximation:

$$\sum_{y \neq x^*(f)}^K \frac{2\sigma^2}{(f^* + \epsilon - f(y))^2} \leq T_\epsilon^*(f) \leq 2 \times \sum_{y \neq x^*(f)}^K \frac{2\sigma^2}{(f^* + \epsilon - f(y))^2}.$$

- Computation: Denoting $f^{x,\epsilon}$ the bandit instance such that $f^{x,\epsilon}(y) = f(y)$ for all $y \neq x$ and $f^{x,\epsilon}(x) = f(x) + \epsilon$,

$$T_\epsilon^*(f) = \min_{x \in \mathcal{X}_\epsilon(f)} T_0(f^{x,\epsilon}) = T_0(f^{x^*(f),\epsilon}),$$

where $x^*(f) \in \mathcal{X}^*(f)$ is an optimal arm \implies simple and efficient algorithm.

GLRT stopping rule

$$\tau_\delta = \inf \left\{ t \in \mathbb{N} : \max_{x \in \mathcal{X}} \inf_{g \in \mathcal{R} \setminus \mathcal{R}_x} \sum_{y \in \mathcal{X}} N_y(t) d(\hat{f}_t(y), g(y)) > \beta(t, \delta) \right\},$$

$$X_{\tau+1} = \operatorname{argmax}_{x=1, \dots, M} \inf_{g \in \mathcal{R} \setminus \mathcal{R}_x} \sum_{y \in \mathcal{X}} N_y(\tau_\delta) d(\hat{f}_{\tau_\delta}(y), \lambda_j) = \operatorname{argmax}_{x \in \mathcal{X}} \hat{f}_\tau(x).$$

Lemma

For any sampling rule, the parallel GLRT test $(\tau_\delta, X_{\tau+1})$ using the threshold function

$$\beta(t, \delta) = 3K \ln(1 + \ln t) + K\mathcal{T}\left(\frac{\ln(1/\delta)}{K}\right)$$

is δ -correct: for all $f \in \mathcal{R}$, $\mathbb{P}_f \left(\tau_\delta < \infty, \hat{X}_{\tau+1} \notin \mathcal{X}^*(f) \right) \leq \delta$.

Tracking the Optimal Proportions

Sampling rule:

$$X_{t+1} \in \begin{cases} \operatorname{argmin}_{x \in U_t} N_x(t) & \text{if } U_t \neq \emptyset \quad (\text{forced exploration}), \text{ or otherwise} \\ \operatorname{argmax}_{x \in \mathcal{X}} t \times w_{\epsilon}^{*, \hat{f}_t(x)}(x) - N_x(t) & (\text{tracking the plug-in estimate}). \end{cases}$$

An instance (f, ϵ) is **regular** if the set

$$\mathcal{W}_{\epsilon}^*(f) = \operatorname{argmax}_{w \in \Sigma_K} \max_{x \in \mathcal{X}_{\epsilon}} \min_{y \neq y} \inf_{\substack{(\lambda_x, \lambda_y): \\ \lambda_x \leq \lambda_y - \epsilon}} w(x) \frac{(f(x) - g(x))^2}{2} + w(y) \frac{(f(y) - g(y))^2}{2}$$

is of cardinality one. For a regular instance, $\mathcal{W}_{\epsilon}^*(f) = \{w_{\epsilon}^*(f)\}$.

Lemma

Let (f, ϵ) be a regular instance of almost optimal best arm identification. Then, under the ϵ -Tracking sampling rule,

$$\mathbb{P}_f \left(\forall x \in \mathcal{X}, \lim_{t \rightarrow \infty} \frac{N_x(t)}{t} = w_{\epsilon}^{*, f}(x) \right) = 1.$$

Optimality of the Track-and-Stop Algorithm

Theorem

For every $\delta \in (0, 1]$, the ϵ -TaS(δ) algorithm is (ϵ, δ) -PAC. Moreover, for every instance (f, ϵ) , if τ_δ denotes the stopping rule of ϵ -TaS(δ),

$$\limsup_{\delta \rightarrow 0} \frac{\mathbb{E}_f[\tau_\delta]}{\ln(1/\delta)} \leq T^*_\epsilon(f) .$$

Experiments

Two regular instances

$$\begin{aligned}f_1 &= [0.2 \ 0.4 \ 0.5 \ 0.55 \ 0.7] & \epsilon_1 &= 0.1 \\f_2 &= [0.4 \ 0.5 \ 0.6 \ 0.7 \ 0.75 \ 0.8] & \epsilon_2 &= 0.15,\end{aligned}$$

and one non-regular instance

$$f_3 = [0.2 \ 0.3 \ 0.45 \ 0.55 \ 0.6 \ 0.6] \quad \epsilon_3 = 0.1.$$

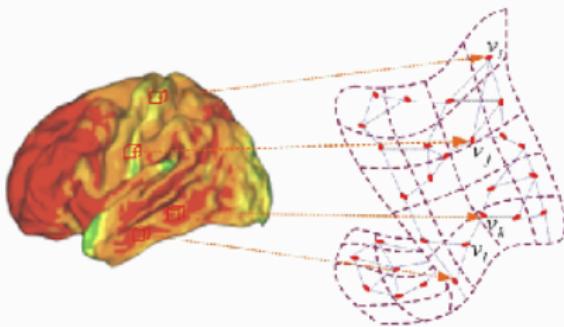
	$T^*(f) \ln(1/\delta)$	ϵ -TaS	KL-LUCB	UGapE	KL-Racing
$f_1, \epsilon_1 = 0.1$	97	171 (104)	322 (137)	324 (143)	372 (159)
$f_2, \epsilon_2 = 0.15$	108	162 (83)	345 (135)	344 (141)	402 (146)
$f_3, \epsilon_3 = 0.1$	531	501 (261)	1236 (403)	1199 (414)	1348 (436)

Table 1: Estimated values of $\mathbb{E}_{\mu_i}[\tau_\delta]$ based on $N = 1000$ repetitions for different instances and algorithms (standard deviation indicated in parenthesis).

Graph Structure and Smooth Functions

Graph Smoothness

See [Spectral bandits for smooth graph functions, by M. Valko, R. Munos, B. Kveton, T. Kocák]



Src: A Spectral Graph Regression Model for Learning Brain Connectivity of Alzheimer's Disease ,by Hu, Cheng, Sepulcre and Li

Weighted graph structure with adjacency matrix $\mathbf{W} = (w_{x,y})_{x,y \in \mathcal{X}}$, and

$$S_G(f) \triangleq \sum_{x,y \in \mathcal{X}} w_{x,y} \frac{(f(x) - f(y))^2}{2} = f^T \mathcal{L} f = \|f\|_{\mathcal{L}}^2 \leq R$$

for some (known) smoothness parameter R , where \mathcal{L} is the **graph**

Laplacian: $\mathcal{L}_{x,y} = -w_{x,y}$ for $x \neq y$ and $\mathcal{L}_{x,x} = \sum_{y \neq x} w_{x,y}$.

\implies The values of f at two points $x, y \in \mathcal{X}$ are close if $w_{x,y}$ is large.

Complexity of Graph BAI

The set of considered signals is

$$\mathcal{M}_R = \{f \in \mathbb{R}^K : f^\top \mathcal{L} f \leq R\}$$

Proposition

For any δ -correct strategy and any R -smooth function f ,

$$\mathbb{E}_f[T_\delta] \geq T_R^*(f) \text{ kl}(\delta, 1 - \delta)$$

where

$$T_R^*(f)^{-1} \triangleq \sup_{\omega \in \Delta_K} \inf_{g \in \mathcal{A}_R(f)} \sum_{x \in \mathcal{X}} \omega_x \frac{(f(x) - g(x))^2}{2} \quad (1)$$

for the set

$$\mathcal{A}_R(f) \triangleq \{g \in \mathcal{M}_R : \exists x \in \mathcal{X} \setminus \{\mathcal{X}_*(f)\}, g(x) \geq g(\mathcal{X}_*(\mu))\}$$

Computing the Complexity

Lemma

Let $\mathcal{D} \subseteq \mathbb{R}^K$ be a compact set. Then function $f : \Delta_K \rightarrow \mathbb{R}$ defined as $f(\omega) = \inf_{d \in \mathcal{D}} \omega^\top d$ is a concave function and $d^*(\omega) = \arg \min_{d \in \mathcal{D}} \omega^\top d$ is a supergradient of f at ω .

Lemma

Let $h : \Delta_K \rightarrow \mathbb{R}$ be a function such that

$$h(w) = \inf_{g \in \text{Alt}_R(f)} \sum_{x \in \mathcal{X}} w(x) \frac{(f(x) - g(x))^2}{2}$$

Then function h is L -Lipschitz with respect to $\|\cdot\|_1$ for any

$$L \geq \max_{x,y \in \mathcal{X}} \frac{(f(x) - f(y))^2}{2}.$$

Mirror Ascent Algorithm

One can compute the best response g_w^* and therefore, the supergradient of concave function $h(w) = \sum_{x \in \mathcal{X}} \omega_x k(f(x), g_w^*(x))$ at w .

Proposition

For $(\Delta_K, \|\cdot\|_1)$, $w_1 = (\frac{1}{K}, \dots, \frac{1}{K})^\top$, $Q = \sqrt{\log K}$, for any $L \geq \max_{x,y \in \mathcal{X}} \frac{(f(x) - f(y))^2}{2}$, mirror ascent with generalized negative entropy as the mirror map

$$\Phi(w) = \sum_{x \in \mathcal{X}} w(x) \log(w(x)) - w(x)$$

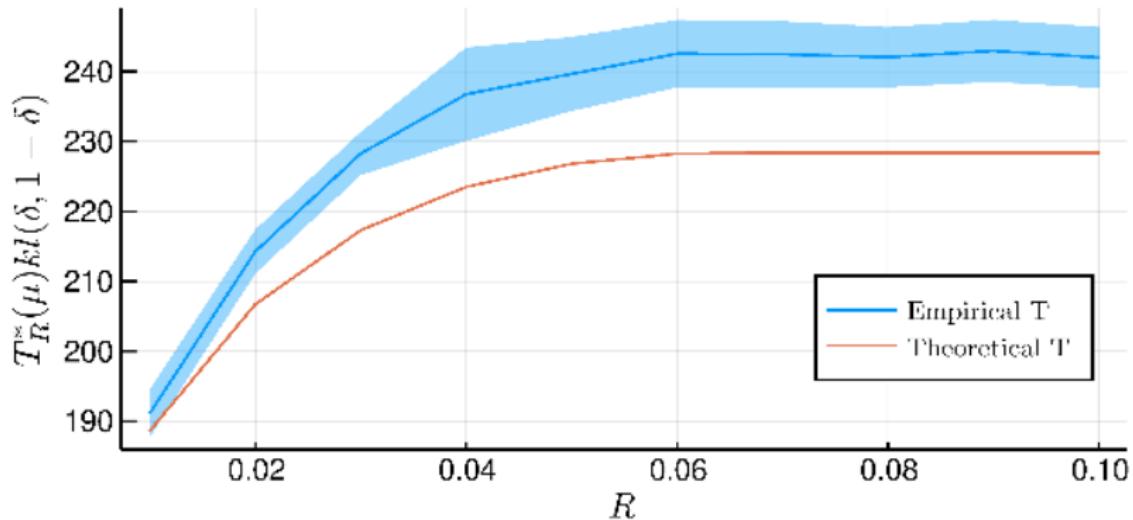
and $\eta = \frac{1}{L} \sqrt{\frac{2 \log K}{t}}$ satisfies:

$$h(w^*) - h\left(\frac{1}{t} \sum_{s=1}^t w_s\right) \leq L \sqrt{\frac{2 \log K}{t}}.$$

- 1: **Input and initialization:**
- 2: \mathcal{L} : graph Laplacian
- 3: δ : confidence parameter
- 4: R : upper bound on the smoothness of f
- 5: Play each $x \in \mathcal{X}$ once and observe rewards $F(x, \omega)$
- 6: \hat{f}_t = empirical estimate of f
- 7: **while** Stopping Rule not satisfied **do**
- 8: Compute $\omega^*(\hat{f}_t)$ by mirror ascent
- 9: Choose X_t according to Tracking Sampling Rule
- 10: Obtain reward $F(X_t, \omega_t)$
- 11: Update \hat{f}_t accordingly
- 12: **end while**
- 13: Output arm $\hat{x} = \operatorname{argmax}_{x \in \mathcal{X}} \hat{f}(x)$

Experiment

Empirical vs theoretical stopping time



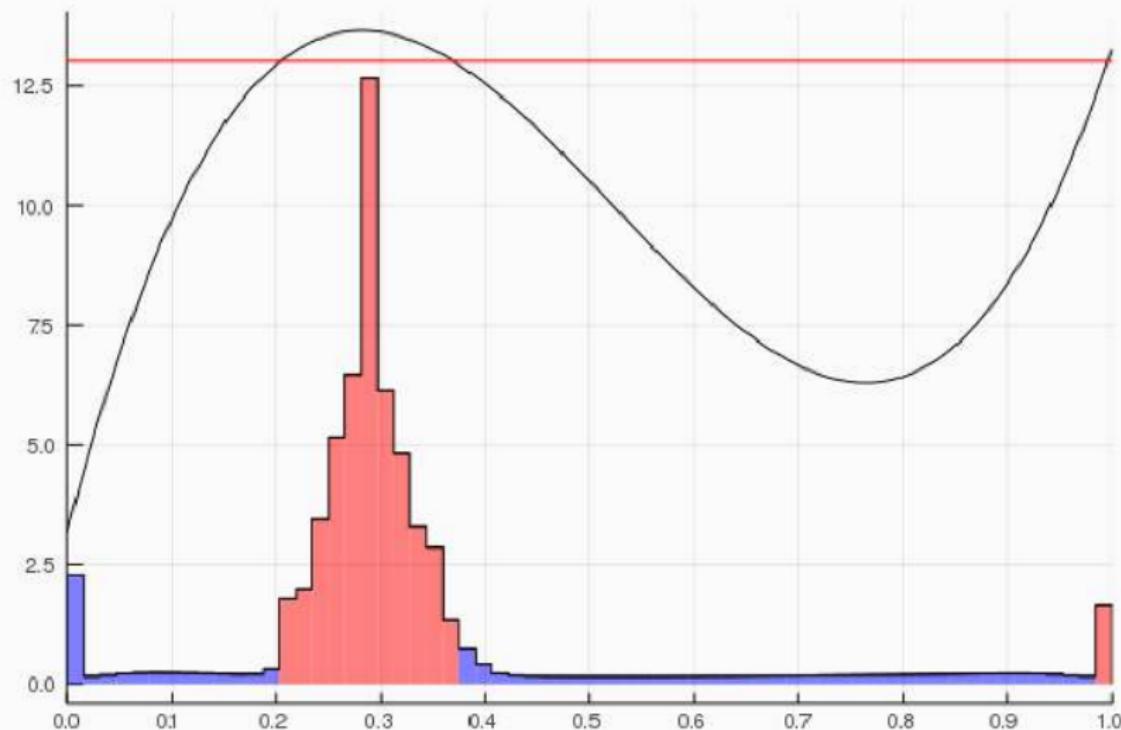
Optimizing a Continuous Function

- Need a way to encode the *regularity* of the function.
 - Discretization: gaph = grid
-
- Idea 1: Lipschitz \implies neighbors means differ by at most L
 - Idea 2: Graph Laplacian regularity

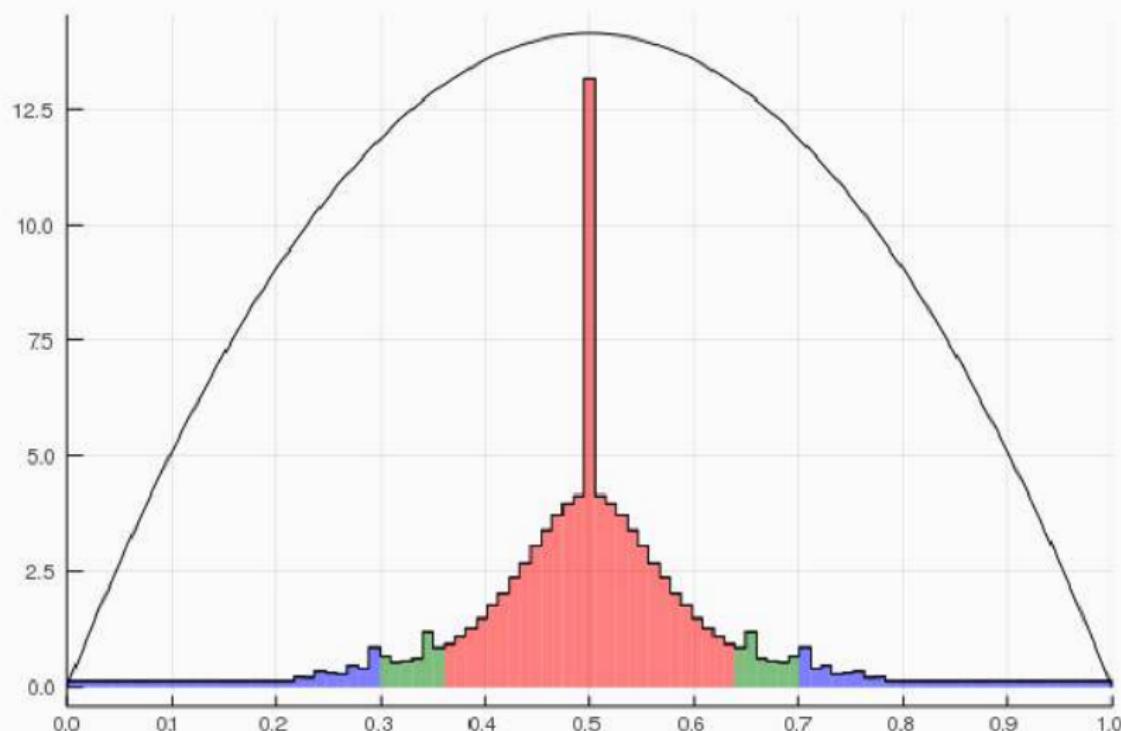
$$\implies \int_{R^d} \|\nabla f(x)\|^2 dx \leq C .$$

- Question: what happens when the grid refines? Is there a *limit*?
 - Limiting Complexity of ϵ, δ -PAC optimization?
 - Limiting density of optimal arm draws?
 - How do known methods compare?

Example



Example (large ϵ)



Example (small ϵ)

