Sensitivity Analysis with Functional Inputs

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- **Sensitivity Analysis**: Understanding the "overall impact" of individual inputs or groups of inputs on the output of a computer model.
- **Computer Model**: Focus on deterministic models numerical implementations of explicitly or implicitly defined functions.
- **Today**: Review and propose a few approaches for extending popular sensitivity/uncertainty ideas developed for scalar-valued inputs to:
 - models for which some inputs are themselves functions, and the output of interest is a scalar. (In fact, the output may be a scalar-valued summary of a function.)
 - further, focus on input functions of one variable, e.g. time.

Examples involving time-varying inputs:

- Regional environment models. Boundary conditions may be time-varying functions.
- Chemical reactor models. "Forcing functions" including temperature, concentration, physical mixing rates.
- Groundwater hydrology models. Rainfall rates, pumping rates.
- Injection molding process models. Heat and pressure schedules.

Notation and Restrictions:

- Model inputs: $(x_1...x_m, z_1(t)...z_n(t)) \in \Delta$
- Model output: $y = f(x_1...x_m, z_1(t)...z_n(t))$
- Attention here is focused on scalar $t \in [0, 1]$, where $z_i(t)$ is continuous and "well-behaved"
- Will sometimes substitute a long vector of values over a *t*-grid for the function:

$$z_i(t) \rightarrow \mathbf{z}_i = \begin{pmatrix} z_i(0.00) \\ z_i(0.01) \\ z_i(0.02) \\ \dots \\ z_i(1.00) \end{pmatrix}$$

Three Basic Approaches popular with scalar-input problems, in decreasing order of the number of function evaluations generally required:

- Variance-based sensitivity analysis A multivariate probability distribution is specified for \mathbf{x} over its domain Δ , representing (ideally) situational uncertainty about \mathbf{x} . The goal is to understand how variability propogates to y. (e.g. Saltelli et al., 2000)
- Statistical surrogate-based sensitivity analysis y is assumed to be a relatively "well behaved" function of x that can be formally predicted or estimated via statistical modeling. Sensitivity of y to each x_i is assessed through model parameters (Welch et al., 1992), by computing variance-based indices on the estimate of f, or via a more formal Bayesian approach (Oakley & O'Hagan, 2004).
- Simple approximation-based sensitivity analysis The sensitivity of output to each input is assessed by numerical approximation to $\partial y/\partial x_i$, i=1,2,3,...,m, or to an average of these quantities over Δ or some appropriate subregion (e.g. $\pm 1\%$ about nominal values).

A Toy Function for Examples:

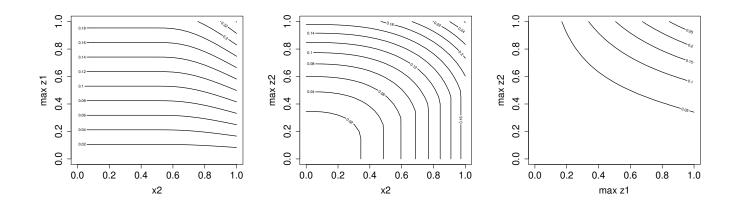
$$y = f(x_1, x_2, z_1, z_2) = \int_{t=0}^{1} \max_{s \in (0, t]} z_1(s) \times \max[(1 - t)x_2, z_2(t)]^2 dt$$

• Note that x_1 does nothing

Some pictures:

$$\bullet \ z_i(t) = \left\{ \begin{array}{ll} 2t \max z_i & t < \frac{1}{2} \\ 2(1-t) \max z_i & t \geq \frac{1}{2} \end{array} \right. \quad i = 1, 2, \quad \max z_i \ \in [0,1]$$

• Unreferenced x or z in each panel $=\frac{1}{2}$

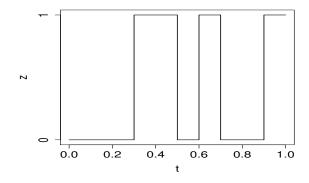


1. Simple approximation-based sensitivity analysis

- Fruth, Roustant, and Kuhnt (2014)
- Restrict attention to input functions z(t) that are:
 - piece-wise constant on intervals defined by a grid on t,

$$G = \{0 = t_0 < t_1 < t_2 < \dots < t_g = 1\}$$

take one of only two values within each interval



• Use a form of sequential bifurcation (Bettonvil, 1995) to progressively refine G. (Important, but I won't consider this aspect here.)

- For a given G, let $\mathbf{z}_i = (z_{i1}, z_{i2}, ... z_{ig})'$.
- Then $y = f(z_1(t), z_2(t), ..., z_n(t)) = f^*(\mathbf{z}_1, \mathbf{z}_2, ..., \mathbf{z}_n)$, i.e. reduction to $g \times n$ two-level scalar-valued inputs ... there is <u>much</u> experimental design literature for this case.
- Define "centered" input values z as $\bar{z}=z-\frac{1}{2}$, so that 0 is the "nominal" value for each input, and $\bar{z}=\pm\frac{1}{2}$.
- ullet The authors use least-squares to fit data from N model runs:

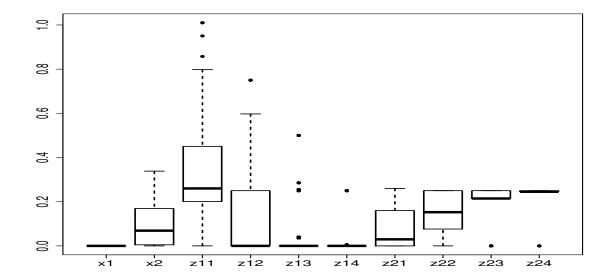
$$(\hat{\alpha}, \hat{\beta}_{ik}, i = 1...n, k = 1...g) = \operatorname{argmin} \sum_{j=1}^{N} [y^j - (\alpha + \sum_{i=1}^{n} \sum_{k=1}^{g} \bar{z}_{ik}^j \beta_{ik})]^2$$

Then use

$$\hat{H}_{ik} = \hat{\beta}_{ik}/(t_k - t_{k-1})$$

as an index of the sensitivity of y to the value of z_i within the kth interval of the t-grid, normalized to be expressed on a per-unit basis of t.

- What should we hope to be estimating here?
- Suppose $G = \{0, \frac{1}{4}, \frac{1}{2}, \frac{3}{4}, 1\}$
- Test function inputs are represented by 2 x's and 8 scalar-valued z's.
- ullet Each input is then associated with $2^{10-1}=512$ "slopes" associated with the edges of a 10-dimensional hypercube ... here they are:



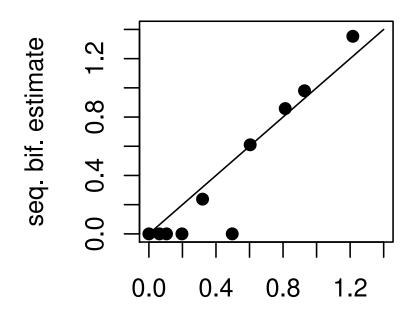
 Basic sequential bifurcation might lead to an accumulated experimental design as follows:

x_1	x_2	z_{11}	z_{12}	z_{13}	z_{14}	z_{21}	z_{22}	z_{23}	z_{24}	y
0	0	0	0	0	0	0	0	0	0	0.0000
1	1	1	1	1	1	1	1	1	1	1.0100
1	1	1	1	1	0	0	0	0	0	0.3384
1	1	1	1	1	1	1	0	0	0	0.3978
1	1	1	1	1	1	1	1	1	0	0.7649
1	1	1	1	1	1	1	1	0	0	0.5504
1	1	0	0	0	0	0	0	0	0	0.0000
1	1	1	1	0	0	0	0	0	0	0.3384
1	1	1	0	0	0	0	0	0	0	0.3384
1	1	1	1	1	1	0	0	0	0	0.3384
1	0	0	0	0	0	0	0	0	0	0.0000

(Note that a different experimental design would have been developed if the inputs had been listed in a different order ...)

• Data collected from this design lead to the following values of \hat{H} (compared to the the "truth" from a full 2^{10} design):

x_1	x_2	z_{12}	z_{12}	z_{13}	z_{14}	z_{21}	z_{22}	z_{23}	z_{24}
0.000	0.000	1.353	0.000	0.000	0.000	0.238	0.610	0.858	0.980
0.000	0.105	1.216	0.497	0.197	0.063	0.320	0.604	0.813	0.928



mean of all from 2¹⁰

- These errors are not realizations of random noise in the data (since there is none), but can be thought of as *bias* in estimators that have no variance.
- If $\mathbf{y} = \mathbf{X}_1 \boldsymbol{\beta}_1$ is used as the basis of analysis, but the data are actually generated by a "true" model: $\mathbf{y} = \mathbf{X}_1 \boldsymbol{\beta}_1 + \mathbf{X}_2 \boldsymbol{\beta}_2$ then the least-squares estimate $\boldsymbol{\beta}_1$ is

$$\hat{\boldsymbol{\beta}}_1 = \boldsymbol{\beta}_1 + (\mathbf{X}_1'\mathbf{X}_1)^{-1}\mathbf{X}_1'\mathbf{X}_2\boldsymbol{\beta}_2 = \boldsymbol{\beta}_1 + \mathbf{A}\boldsymbol{\beta}_2$$

- The experimental design determines X_1 and X_2 , and hence the alias matrix A.
- Mitchell (1974) proposed using $||\mathbf{A}||$ as an index of design quality for estimating main-effects models when second-order terms are present in the data-generating process.
- We modify this idea slightly here to omit the first row of **A** since this corresponds to bias in the model intercept, which is of no real interest to us.

\hat{H} and alias indices for designs of different sizes:

- SB = Sequential Bifurcation (as shown)
- FO SB = Foldover of Sequential Bifurcation design
- PB = minimal Plackett-Burman design
- FO PB = Foldover of Plackett-Burman design
- 2_{III}^{10-5} = Minimum Aberation Regular Fraction of Resolution III
- ullet $2_{IV}^{10-4}=$ Minimum Aberation Regular Fraction of Resolution IV
- $2_{IV}^{10-3} =$ (larger) Minimum Aberation Regular Fraction of Resolution IV
- 2¹⁰ = Full Two-Level Factorial design

design	x1	x_2	z_{11}	z_{12}	z_{13}	z_{14}	z_{21}	z_{22}	z_{23}	z_{24}	N	$ \mathbf{A}_2 $	$ \mathbf{A}_3 $
SB	0.000	0.000	1.353	0.000	0.000	0.000	0.238	0.610	0.858	0.980	11	22.50	22.50
FO SB	0.000	0.000	1.197	0.500	0.500	0.500	0.119	0.305	0.429	0.490	20	0	22.50
РВ	-0.106	0.104	1.239	0.416	-0.049	-0,035	0.598	0.245	0.394	1.206	12	10.00	5.83
FO PB	-0.047	0.066	1.098	0.445	0.079	-0.213	0.340	0.559	0.842	0.879	24	0	5.83
2^{10-6}_{III}	0.017	0.138	1.394	<u>0.674</u>	0.174	0.076	0.325	0.778	0.555	0.817	16	6.00	4.50
2_{IV}^{10-5}	0.014	0.120	1.217	0.496	0.198	0.055	0.173	0.659	0.825	0.938	32	0	2.50
2_{IV}^{10-4}	0.002	0.105	1.216	0.496	0.197	0.063	0.320	0.603	0.751	0.812	64	0	0.50
2^{10}	0.000	0.105	1.216	0.497	0.197	0.063	0.320	0.604	0.813	0.928	1024	0	0.00

(Underlines are errors of more than 0.10)

2. Variance-based sensitivity analysis

- looss and Ribatet (2009), Jacques et al. (2006) advocate a direct extension of the standard approach for scalar inputs, called the *microparameter method*.
- Quick reminder of the popular scalar-input approach

\mathbf{A}	В	\mathbf{A}_1	${f A}_2$	\mathbf{A}_3
$\begin{bmatrix} x_1 & x_2 & x_3 \end{bmatrix}$				
.23 .46 .81	.53 .27 .26	.53 .46 .81	.23 .27 .81	.23 .46 .26
.71 .52 .33	.21 .04 .37	.21 .52 .33	.71 .04 .33	.71 .52 .37
.48 .21 .50	.88 .49 .94	.88 .21 .50	.48 .49 .50	.48 .21 .94

• Then averages of squared differences of outputs corresponding to paired rows form the basis of sensitivity index estimates:

- $\mathbf{A} \& \mathbf{B} \to \widehat{\mathsf{Var}}(y)$, the unconditional variance
- $\mathbf{B}\&\mathbf{A}_1 \to E_{x_1}\widehat{\mathsf{Var}}_{x_2,x_3}[y|x_1]$
 - First-Order Sensitivity: $\hat{S}(x_1) = 1 E_{x_1} \widehat{\mathsf{Var}}_{x_2,x_3}[y|x_1]/\widehat{\mathsf{Var}}(y)$
- $\mathbf{A} \& \mathbf{A}_1 \to \widehat{E_{x_2,x_2} \mathsf{Var}_{x_1}}[y|x_2,x_3]$
 - Total Sensitivity: $\hat{T}(x_1) = E_{x_2,x_2} \widehat{\mathsf{Var}}_{x_1}[y|x_2,x_3]/\mathsf{Var}(y)$
- and similarly for other inputs, using a different A_i but the same A and B in each case.

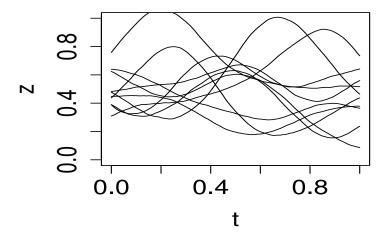
The same approach can be taken when any or all inputs are functional

- Functional inputs (or their vector approximations) are regarded as realizations of stochastic processes (or multivariate distributions)
- For example, a Gaussian process with

$$E(z(t)) = \frac{1}{2}, \ \operatorname{Var}(z(t)) = (\frac{1}{6})^2$$

$$\operatorname{Corr}(z(t_1), z(t_2)) = e^{-\theta|t_1 - t_2|^{1.99}} \ \text{with} \ \theta = 10.$$

Realizations:



• In the examples that follow, I use this process model for both z_1 and z_2 , and represent them as 101-element vectors \mathbf{z}_1 and \mathbf{z}_2 .

• With:

- x_1 and $x_2 \sim U[0,1]$, and each of \mathbf{z}_1 and \mathbf{z}_2 as described above
- 6 input arrays, 100,000 rows per array (600,000 function evaluations)

results for the example model are:

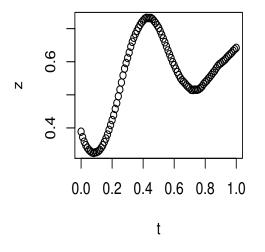
	x_1	x_2	${f z}_1$	${f z}_2$
\hat{S}	0.0092	0.1065	0.2565	0.5937
\hat{T}	0.0000	0.1277	0.2896	0.6382

- This is useful, but it offers little insight into how z_1 and z_2 influence y.
- Proposal: "Factor" the functional input into one or a few scalar-valued summaries and an *independent* functional residual (of hopefully little importance).

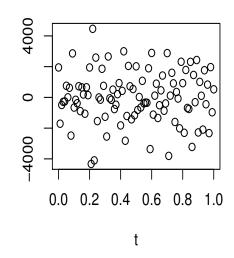
Special case: Gaussian processes: $\mathbf{z} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

- Scalar-valued summaries: $\mathbf{s} = \mathbf{C}'\mathbf{z}$
 - e.g. coefficients of a low-order least-squares polynomial approximation to ${\bf z}$
- ullet A "residual": $\mathbf{r} = (\mathbf{I} \mathbf{C}(\mathbf{C}'\mathbf{C})^{-1}\mathbf{C}')\mathbf{\Sigma}^{-1}\mathbf{z}$
- ullet Both ${f s}$ and ${f r}$ are multivariate normal, and independent, and ${f z}$ can be recovered from ${f s}$ and ${f r}$

- Example:
 - Univariate $s=\bar{z}$
 - ${f z}$ generated as before



$$\rightarrow s = 0.5522 \text{ and}$$



• Hence, our example can be viewed as:

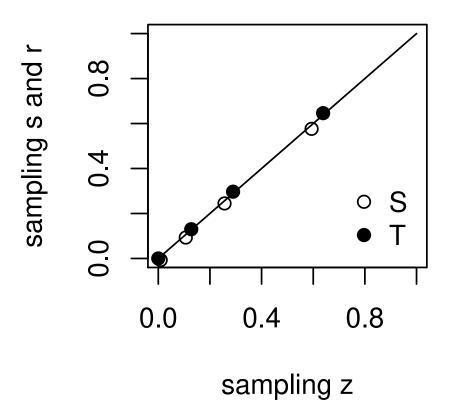
$$y = f(x_1, x_2, s_1, \mathbf{r}_1, s_2, \mathbf{r}_2)$$

- Use $s_1 = \operatorname{ave}(\mathbf{z}_1)$ and $s_2 = \operatorname{ave}(\mathbf{z}_2)$
- Now 8 input arrays, 100,000 rows per array (800,000 function evaluations)

	x_1	x_2	s_1	${f r}_1$	s_2	\mathbf{r}_2
\hat{S}	-0.0068	0.0931	0.1350	0.1096	0.5534	0.0230
\hat{T}	0	0.1300	0.1656	0.1313	0.5989	0.0472

- ullet s_2 is important, while ${f r}_2$ has little impact
- ullet s_1 is more important than ${f r}_1$, which is comparable to x_2





• Comparing $S(\mathbf{z}_1)$ to $S(s_1) + S(\mathbf{r}_1)$, et cetera

3. Statistical surrogate-based sensitivity analysis

- loose and Ribatet (2009) also discussed using a *joint modeling* approach to sensitivity analysis with functional inputs, based on fitting two models to output data.
- The (conditional) mean and variance of the output are modeled as functions of scalar-valued inputs only, i.e.
 - for inputs = $(x_1, ..., x_m, z_1(t), ..., z_n(t))$,
 - estimate models for $E(y|x_1,...,x_m)$ and $Var(y|x_1,...,x_m)$.
- So, for example, $E_{x's} Var_{z's}(y|x's)$ can be estimated by integrating the dispersion model w.r.t. the distribution of x's, et cetera.
- Authors used GLM and GAM in their examples.
- In this form, the approach does not separate the variability associated with different functional inputs.

Here I'll try something related, and refer to it as "semi-modeling":

• Draw F realizations of each input function,

$$z_1^{i_1}(t)...z_n^{i_n}(t), i_1...i_n = 1...F.$$

Model y only for the selected function values, i.e.

$$y = f(x_1...x_m, i_1...i_n)$$

where $i_1...i_n$ are categorical variables, each with values 1...F, indexing associated input function values.

• Given training data, fit a single predictive model of the output:

$$\hat{y} = \hat{f}(x_1...x_m, i_1...i_n)$$

- Then, for example, using a random sample of size R (much larger than F) of each of $x_1...x_m, i_1...i_n$ and $x_1'...x_m', i_1'...i_n'$, $-\widehat{\text{Var}}(y) = \frac{1}{2R} \sum_{r=1}^{R} (\hat{y}(x_1^r...x_m^r, i_1^r...i_n^r) \hat{y}(x_1'^r...x_m'^r, i_1'^r...i_n'^r))^2$
 - $-\hat{T}(x_1) = \frac{1}{2R} \sum_{r=1}^{R} (\hat{y}(x_1^r ... x_m^r, i_1^r ... i_n^r) \hat{y}(x_1'^r ... x_m^r, i_1^r ... i_n^r))^2 / \widehat{\mathsf{Var}}(y)$
 - $\hat{S}(x_1) =$

$$[\widehat{\mathsf{Var}}(y) - \tfrac{1}{2R} \sum_{r=1}^R (\hat{y}(x_1^r...x_m^r, i_1^r...i_n^r) - \hat{y}(x_1^r...x_m'^r, i_1'^r...i_n'^r))^2]/\widehat{\mathsf{Var}}(y)$$

and similarly for other inputs, both scalar and functional.

ullet Here I model y with a stationary Gaussian stochastic process model, where for

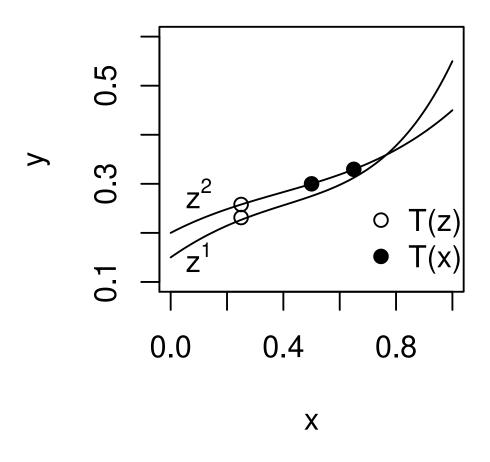
$$y = f(x_1...x_m, i_1, ...i_n), y' = f(x_1'...x_m', i_1', ...i_n'),$$

$$E(y) = E(y') = \mu, \text{Var}(y) = \text{Var}(y') = \sigma^2, \text{Cov}(y, y') = \sigma^2 e^{-\text{dist}},$$

$$\text{dist} = \sum_{i=1}^m \theta_j (x_j - x_j')^2 + \sum_{i=1}^n \phi_j I(i_j \neq i_j'),$$

fitting parameters via maximum likelihood.



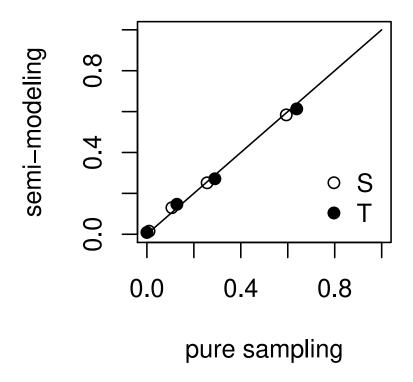


Results for the example model:

- F = 50 realizations of each of x_1 , x_2 , \mathbf{z}_1 and \mathbf{z}_2 , distributed as before.
- Design constructed by repeating each input value 5 times, and forming the N=250-run experimental design via the maximin distance criterion.
- Result provides a predictor of y for any combination of x_1 , x_2 and any of the 50 drawn realizations for each of \mathbf{z}_1 and \mathbf{z}_2 .
- Results (R = 10,000):

	x_1	x_2	${f z}_1$	${f z}_2$
\hat{S}	0.0151	0.1299	0.2517	0.5839
\hat{T}	0.0088	0.1465	0.2715	0.6133

4-Input Comparison

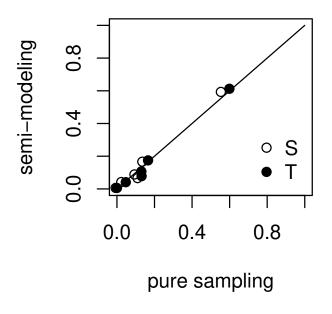


• Results are consistent with those from the pure sampling-based approach, but requiring far fewer function evaluations.

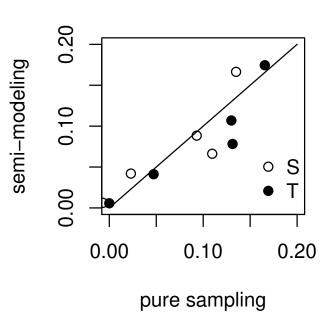
• Replacing \mathbf{z}_1 with s_1 and \mathbf{r}_1 , and \mathbf{z}_2 with s_2 and \mathbf{r}_2 :

	x_1	x_2	s_1	${f r}_1$	s_2	${f r}_2$
\hat{S}	0.0062	0.0884	0.1664	0.0662	0.5931	0.0420
\hat{T}	0.0057	0.1069	0.1745	0.0782	0.6124	0.0412

6-Input Comparison



Smallest 5 of 6



Concluding thoughts:

- Can more bias-resistant alternatives to Sequential Bifurcation be developed for the piecewise constant inputs case (that doesn't require too many runs)?
- Traditional variance-based sensitivity analysis may be most effective if functional inputs can be decomposed into independent (1.) important low-dimensional, and (2.) less important higher-dimensional components.
- Meta-models that are accurate approximations for a moderate sample of functional inputs may improve the efficiency of variance-based sensitivity analysis.

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