# Sensitivity Analysis with Functional Inputs 

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- Sensitivity Analysis: Understanding the "overall impact" of individual inputs or groups of inputs on the output of a computer model.
- Computer Model: Focus on deterministic models - numerical implementations of explicitly or implicitly defined functions.
- Today: Review and propose a few approaches for extending popular sensitivity/uncertainty ideas developed for scalar-valued inputs to:
- models for which some inputs are themselves functions, and the output of interest is a scalar. (In fact, the output may be a scalar-valued summary of a function.)
- further, focus on input functions of one variable, e.g. time.


## Examples involving time-varying inputs:

- Regional environment models. Boundary conditions may be time-varying functions.
- Chemical reactor models. "Forcing functions" including temperature, concentration, physical mixing rates.
- Groundwater hydrology models. Rainfall rates, pumping rates.
- Injection molding process models. Heat and pressure schedules.


## Notation and Restrictions:

- Model inputs: $\left(x_{1} \ldots x_{m}, z_{1}(t) \ldots z_{n}(t)\right) \in \Delta$
- Model output: $y=f\left(x_{1} \ldots x_{m}, z_{1}(t) \ldots z_{n}(t)\right)$
- Attention here is focused on scalar $t \in[0,1]$, where $z_{i}(t)$ is continuous and "well-behaved"
- Will sometimes substitute a long vector of values over a $t$-grid for the function:

$$
z_{i}(t) \rightarrow \mathbf{z}_{i}=\left(\begin{array}{c}
z_{i}(0.00) \\
z_{i}(0.01) \\
z_{i}(0.02) \\
\ldots \\
z_{i}(1.00)
\end{array}\right)
$$

Three Basic Approaches popular with scalar-input problems, in decreasing order of the number of function evaluations generally required:

- Variance-based sensitivity analysis - A multivariate probability distribution is specified for x over its domain $\Delta$, representing (ideally) situational uncertainty about $\mathbf{x}$. The goal is to understand how variability propogates to $y$. (e.g. Saltelli et al., 2000)
- Statistical surrogate-based sensitivity analysis - $y$ is assumed to be a relatively "well behaved" function of $\mathbf{x}$ that can be formally predicted or estimated via statistical modeling. Sensitivity of $y$ to each $x_{i}$ is assessed through model parameters (Welch et al., 1992), by computing variance-based indices on the estimate of $f$, or via a more formal Bayesian approach (Oakley \& O'Hagan, 2004).
- Simple approximation-based sensitivity analysis - The sensitivity of output to each input is assessed by numerical approximation to $\partial y / \partial x_{i}$, $i=1,2,3, \ldots, m$, or to an average of these quantities over $\Delta$ or some appropriate subregion (e.g. $\pm 1 \%$ about nominal values).


## A Toy Function for Examples:

$$
y=f\left(x_{1}, x_{2}, z_{1}, z_{2}\right)=\int_{t=0}^{1} \max _{s \in(0, t]} z_{1}(s) \times \max \left[(1-t) x_{2}, z_{2}(t)\right]^{2} d t
$$

- Note that $x_{1}$ does nothing

Some pictures:

$$
\text { - } z_{i}(t)=\left\{\begin{array}{ll}
2 t \max z_{i} & t<\frac{1}{2} \\
2(1-t) \max z_{i} & t \geq \frac{1}{2}
\end{array} \quad i=1,2, \quad \max z_{i} \in[0,1]\right.
$$

- Unreferenced $x$ or $z$ in each panel $=\frac{1}{2}$



## 1. Simple approximation-based sensitivity analysis

- Fruth, Roustant, and Kuhnt (2014)
- Restrict attention to input functions $z(t)$ that are:
- piece-wise constant on intervals defined by a grid on $t$,

$$
G=\left\{0=t_{0}<t_{1}<t_{2}<\ldots<t_{g}=1\right\}
$$

- take one of only two values within each interval

- Use a form of sequential bifurcation (Bettonvil, 1995) to progressively refine $G$. (Important, but I won't consider this aspect here.)
- For a given $G$, let $\mathbf{z}_{i}=\left(z_{i 1}, z_{i 2}, \ldots z_{i g}\right)^{\prime}$.
- Then $y=f\left(z_{1}(t), z_{2}(t), \ldots, z_{n}(t)\right)=f^{*}\left(\mathbf{z}_{1}, \mathbf{z}_{2}, \ldots, \mathbf{z}_{n}\right)$, i.e. reduction to $g \times n$ two-level scalar-valued inputs ... there is much experimental design literature for this case.
- Define "centered" input values $z$ as $\bar{z}=z-\frac{1}{2}$, so that 0 is the "nominal" value for each input, and $\bar{z}= \pm \frac{1}{2}$.
- The authors use least-squares to fit data from $N$ model runs:

$$
\left(\hat{\alpha}, \hat{\beta}_{i k}, i=1 \ldots n, k=1 \ldots g\right)=\operatorname{argmin} \sum_{j=1}^{N}\left[y^{j}-\left(\alpha+\sum_{i=1}^{n} \sum_{k=1}^{g} \bar{z}_{i k}^{j} \beta_{i k}\right)\right]^{2}
$$

- Then use

$$
\hat{H}_{i k}=\hat{\beta}_{i k} /\left(t_{k}-t_{k-1}\right)
$$

as an index of the sensitivity of $y$ to the value of $z_{i}$ within the $k$ th interval of the $t$-grid, normalized to be expressed on a per-unit basis of $t$.

- What should we hope to be estimating here?
- Suppose $G=\left\{0, \frac{1}{4}, \frac{1}{2}, \frac{3}{4}, 1\right\}$
- Test function inputs are represented by $2 x$ 's and 8 scalar-valued $z$ 's.
- Each input is then associated with $2^{10-1}=512$ "slopes" associated with the edges of a 10 -dimensional hypercube ... here they are:

- Basic sequential bifurcation might lead to an accumulated experimental design as follows:

| $x_{1}$ | $x_{2}$ | $z_{11}$ | $z_{12}$ | $z_{13}$ | $z_{14}$ | $z_{21}$ | $z_{22}$ | $z_{23}$ | $z_{24}$ | $y$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.0000 |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1.0100 |
| 1 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0.3384 |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0.3978 |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 0 | 0.7649 |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 0 | 0 | 0.5504 |
| 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.0000 |
| 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.3384 |
| 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.3384 |
| 1 | 1 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0.3384 |
| 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.0000 |

(Note that a different experimental design would have been developed if the inputs had been listed in a different order ...)

- Data collected from this design lead to the following values of $\hat{H}$ (compared to the the "truth" from a full $2^{10}$ design):

| $x_{1}$ | $x_{2}$ | $z_{12}$ | $z_{12}$ | $z_{13}$ | $z_{14}$ | $z_{21}$ | $z_{22}$ | $z_{23}$ | $z_{24}$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 0.000 | 0.000 | 1.353 | $\underline{0.000}$ | $\underline{0.000}$ | 0.000 | 0.238 | 0.610 | 0.858 | 0.980 |
| 0.000 | 0.105 | 1.216 | 0.497 | 0.197 | 0.063 | 0.320 | 0.604 | 0.813 | 0.928 |



- These errors are not realizations of random noise in the data (since there is none), but can be thought of as bias in estimators that have no variance.
- If $\mathbf{y}=\mathbf{X}_{1} \boldsymbol{\beta}_{1}$ is used as the basis of analysis, but the data are actually generated by a "true" model: $\mathbf{y}=\mathbf{X}_{1} \boldsymbol{\beta}_{1}+\mathbf{X}_{2} \boldsymbol{\beta}_{2}$ then the least-squares estimate $\boldsymbol{\beta}_{1}$ is

$$
\hat{\boldsymbol{\beta}}_{1}=\boldsymbol{\beta}_{1}+\left(\mathbf{X}_{1}^{\prime} \mathbf{X}_{1}\right)^{-1} \mathbf{X}_{1}^{\prime} \mathbf{X}_{2} \boldsymbol{\beta}_{2}=\boldsymbol{\beta}_{1}+\mathbf{A} \boldsymbol{\beta}_{2}
$$

- The experimental design determines $\mathbf{X}_{1}$ and $\mathbf{X}_{2}$, and hence the alias matrix $\mathbf{A}$.
- Mitchell (1974) proposed using $\|\mathbf{A}\|$ as an index of design quality for estimating main-effects models when second-order terms are present in the data-generating process.
- We modify this idea slightly here to omit the first row of $\mathbf{A}$ since this corresponds to bias in the model intercept, which is of no real interest to us.
$\hat{H}$ and alias indices for designs of different sizes:
- $\mathrm{SB}=$ Sequential Bifurcation (as shown)
- $\mathrm{FO} S B=$ Foldover of Sequential Bifurcation design
- $\mathrm{PB}=$ minimal Plackett-Burman design
- FO PB $=$ Foldover of Plackett-Burman design
- $2_{I I I}^{10-5}=$ Minimum Aberation Regular Fraction of Resolution III
- $2_{I V}^{10-4}=$ Minimum Aberation Regular Fraction of Resolution IV
- $2_{I V}^{10-3}=$ (larger) Minimum Aberation Regular Fraction of Resolution IV
- $2^{10}=$ Full Two-Level Factorial design

(Underlines are errors of more than 0.10 )


## 2. Variance-based sensitivity analysis

- looss and Ribatet (2009), Jacques et al. (2006) advocate a direct extension of the standard approach for scalar inputs, called the microparameter method.
- Quick reminder of the popular scalar-input approach

| A | B | $\mathbf{A}_{1}$ | $\mathbf{A}_{2}$ | $\mathbf{A}_{3}$ |
| :---: | :---: | :---: | :---: | :---: |
| $x_{1}$ $x_{2}$ $x_{3}$ | $x_{1}$ $x_{2}$ $x_{3}$ | $x_{1}$ $x_{2}$ $x_{3}$ | $x_{1}$ $x_{2}$ $x_{3}$ | $x_{1} x_{2} \quad x_{3}$ |
| . 23.46 .81 | . 53.27 .26 | . 53.46 .81 | . 23.27 .81 | . 23.46 .26 |
| . 71.52 .33 | . 21.04 .37 | . 21.52 .33 | . 71.04 .33 | . 71.52 .37 |
|  | . .. | ... ... | ..... | ... ... |
| . 48.21 .50 | . 88.49 .94 | . 88.21 .50 | . 48.49 .50 | . 48.21 .94 |

- Then averages of squared differences of outputs corresponding to paired rows form the basis of sensitivity index estimates:
- $\mathbf{A} \& \mathbf{B} \rightarrow \widehat{\operatorname{Var}}(y)$, the unconditional variance
- $\mathbf{B} \& \mathbf{A}_{1} \rightarrow E_{x_{1}} \widehat{\operatorname{Var}}_{x_{2}, x_{3}}\left[y \mid x_{1}\right]$
- First-Order Sensitivity: $\hat{S}\left(x_{1}\right)=1-E_{x_{1}} \widehat{\operatorname{Var}}_{x_{2}, x_{3}}\left[y \mid x_{1}\right] / \widehat{\operatorname{Var}}(y)$
- $\mathbf{A \&} \mathbf{A}_{1} \rightarrow E_{x_{2}, x_{2}} \widehat{V a r}_{x_{1}}\left[y \mid x_{2}, x_{3}\right]$
- Total Sensitivity: $\hat{T}\left(x_{1}\right)=E_{x_{2}, x_{2}} \widehat{\operatorname{Var}}_{x_{1}}\left[y \mid x_{2}, x_{3}\right] / \operatorname{Var}(y)$
- and similarly for other inputs, using a different $\mathbf{A}_{i}$ but the same $\mathbf{A}$ and $\mathbf{B}$ in each case.

The same approach can be taken when any or all inputs are functional

- Functional inputs (or their vector approximations) are regarded as realizations of stochastic processes (or multivariate distributions)
- For example, a Gaussian process with

$$
\begin{gathered}
E(z(t))=\frac{1}{2}, \operatorname{Var}(z(t))=\left(\frac{1}{6}\right)^{2} \\
\operatorname{Corr}\left(z\left(t_{1}\right), z\left(t_{2}\right)\right)=e^{-\theta\left|t_{1}-t_{2}\right|^{1.99}} \text { with } \theta=10
\end{gathered}
$$

- Realizations:

- In the examples that follow, I use this process model for both $z_{1}$ and $z_{2}$, and represent them as 101-element vectors $\mathbf{z}_{1}$ and $\mathbf{z}_{2}$.
- With:
$-x_{1}$ and $x_{2} \sim U[0,1]$, and each of $\mathbf{z}_{1}$ and $\mathbf{z}_{2}$ as described above
- 6 input arrays, 100,000 rows per array (600,000 function evaluations)
results for the example model are:

|  | $x_{1}$ | $x_{2}$ | $\mathbf{z}_{1}$ | $\mathbf{z}_{2}$ |
| ---: | ---: | ---: | ---: | ---: |
| $\hat{S}$ | 0.0092 | 0.1065 | 0.2565 | 0.5937 |
| $\hat{T}$ | 0.0000 | 0.1277 | 0.2896 | 0.6382 |

- This is useful, but it offers little insight into how $z_{1}$ and $z_{2}$ influence $y$.
- Proposal: "Factor" the functional input into one or a few scalar-valued summaries and an independent functional residual (of hopefully little importance).

Special case: Gaussian processes: $\mathbf{z} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

- Scalar-valued summaries: $\mathrm{s}=\mathbf{C}^{\prime} \mathbf{z}$
- e.g. coefficients of a low-order least-squares polynomial approximation to $\mathbf{z}$
- A "residual" : $\mathbf{r}=\left(\mathbf{I}-\mathbf{C}\left(\mathbf{C}^{\prime} \mathbf{C}\right)^{-1} \mathbf{C}^{\prime}\right) \boldsymbol{\Sigma}^{-1} \mathbf{z}$
- Both $\mathbf{s}$ and $\mathbf{r}$ are multivariate normal, and independent, and $\mathbf{z}$ can be recovered from $\mathbf{s}$ and $\mathbf{r}$
- Example:
- Univariate $s=\bar{z}$
- z generated as before

N

$\rightarrow s=0.5522$ and
t


- Hence, our example can be viewed as:

$$
y=f\left(x_{1}, x_{2}, s_{1}, \mathbf{r}_{1}, s_{2}, \mathbf{r}_{2}\right)
$$

- Use $s_{1}=\operatorname{ave}\left(\mathbf{z}_{1}\right)$ and $s_{2}=\operatorname{ave}\left(\mathbf{z}_{2}\right)$
- Now 8 input arrays, 100,000 rows per array (800,000 function evaluations)

|  | $x_{1}$ | $x_{2}$ | $s_{1}$ | $\mathbf{r}_{1}$ | $s_{2}$ | $\mathbf{r}_{2}$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $\hat{S}$ | -0.0068 | 0.0931 | 0.1350 | 0.1096 | 0.5534 | 0.0230 |
| $\hat{T}$ | 0 | 0.1300 | 0.1656 | 0.1313 | 0.5989 | 0.0472 |

- $s_{2}$ is important, while $\mathbf{r}_{2}$ has little impact
- $s_{1}$ is more important than $\mathbf{r}_{1}$, which is comparable to $x_{2}$


## Comparison



- Comparing $S\left(\mathbf{z}_{1}\right)$ to $S\left(s_{1}\right)+S\left(\mathbf{r}_{1}\right)$, et cetera


## 3. Statistical surrogate-based sensitivity analysis

- loose and Ribatet (2009) also discussed using a joint modeling approach to sensitivity analysis with functional inputs, based on fitting two models to output data.
- The (conditional) mean and variance of the output are modeled as functions of scalar-valued inputs only, i.e.
- for inputs $=\left(x_{1}, \ldots, x_{m}, z_{1}(t), \ldots, z_{n}(t)\right)$,
- estimate models for $E\left(y \mid x_{1}, \ldots, x_{m}\right)$ and $\operatorname{Var}\left(y \mid x_{1}, \ldots, x_{m}\right)$.
- So, for example, $E_{x^{\prime} s} \operatorname{Var}_{z^{\prime} s}\left(y \mid x^{\prime} s\right)$ can be estimated by integrating the dispersion model w.r.t. the distribution of $x^{\prime} s$, et cetera.
- Authors used GLM and GAM in their examples.
- In this form, the approach does not separate the variability associated with different functional inputs.

Here I'll try something related, and refer to it as "semi-modeling":

- Draw $F$ realizations of each input function,

$$
z_{1}^{i_{1}}(t) \ldots z_{n}^{i_{n}}(t), i_{1} \ldots i_{n}=1 \ldots F .
$$

- Model $y$ only for the selected function values, i.e.

$$
y=f\left(x_{1} \ldots x_{m}, i_{1} \ldots i_{n}\right)
$$

where $i_{1} \ldots i_{n}$ are categorical variables, each with values $1 \ldots F$, indexing associated input function values.

- Given training data, fit a single predictive model of the output:

$$
\hat{y}=\hat{f}\left(x_{1} \ldots x_{m}, i_{1} \ldots i_{n}\right)
$$

- Then, for example, using a random sample of size $R$ (much larger than $F$ ) of each of $x_{1} \ldots x_{m}, i_{1} \ldots i_{n}$ and $x_{1}^{\prime} \ldots x_{m}^{\prime}, i_{1}^{\prime} \ldots i_{n}^{\prime}$,

$$
\begin{aligned}
& -\widehat{\operatorname{Var}}(y)=\frac{1}{2 R} \sum_{r=1}^{R}\left(\hat{y}\left(x_{1}^{r} \ldots x_{m}^{r}, i_{1}^{r} \ldots i_{n}^{r}\right)-\hat{y}\left(x_{1}^{\prime r} \ldots x_{m}^{\prime r}, i_{1}^{\prime r} \ldots i_{n}^{\prime r}\right)\right)^{2} \\
& -\hat{T}\left(x_{1}\right)=\frac{1}{2 R} \sum_{r=1}^{R}\left(\hat{y}\left(x_{1}^{r} \ldots x_{m}^{r}, i_{1}^{r} \ldots i_{n}^{r}\right)-\hat{y}\left(x_{1}^{\prime r} \ldots x_{m}^{r}, i_{1}^{r} \ldots i_{n}^{r}\right)\right)^{2} / \widehat{\operatorname{Var}}(y) \\
& -\hat{S}\left(x_{1}\right)= \\
& {\left[\widehat{\operatorname{Var}}(y)-\frac{1}{2 R} \sum_{r=1}^{R}\left(\hat{y}\left(x_{1}^{r} \ldots x_{m}^{r}, i_{1}^{r} \ldots i_{n}^{r}\right)-\hat{y}\left(x_{1}^{r} \ldots x_{m}^{\prime r}, i_{1}^{\prime r} \ldots i_{n}^{\prime r}\right)\right)^{2}\right] / \widehat{\operatorname{Var}}(y)}
\end{aligned}
$$

and similarly for other inputs, both scalar and functional.

- Here I model $y$ with a stationary Gaussian stochastic process model, where for

$$
\begin{gathered}
y=f\left(x_{1} \ldots x_{m}, i_{1}, \ldots i_{n}\right), y^{\prime}=f\left(x_{1}^{\prime} \ldots x_{m}^{\prime}, i_{1}^{\prime}, \ldots i_{n}^{\prime}\right) \\
E(y)=E\left(y^{\prime}\right)=\mu, \operatorname{Var}(y)=\operatorname{Var}\left(y^{\prime}\right)=\sigma^{2}, \operatorname{Cov}\left(y, y^{\prime}\right)=\sigma^{2} e^{-\mathrm{dist}} \\
\operatorname{dist}=\sum_{j=1}^{m} \theta_{j}\left(x_{j}-x_{j}^{\prime}\right)^{2}+\sum_{j=1}^{n} \phi_{j} I\left(i_{j} \neq i_{j}^{\prime}\right)
\end{gathered}
$$

fitting parameters via maximum likelihood.

## Semi-Modeling



Results for the example model:

- $F=50$ realizations of each of $x_{1}, x_{2}, \mathbf{z}_{1}$ and $\mathbf{z}_{2}$, distributed as before.
- Design constructed by repeating each input value 5 times, and forming the $N=250$-run experimental design via the maximin distance criterion.
- Result provides a predictor of $y$ for any combination of $x_{1}, x_{2}$ and any of the 50 drawn realizations for each of $\mathbf{z}_{1}$ and $\mathbf{z}_{2}$.
- Results ( $R=10,000$ ):

|  | $x_{1}$ | $x_{2}$ | $\mathbf{z}_{1}$ | $\mathbf{z}_{2}$ |
| :--- | ---: | ---: | ---: | ---: |
| $\hat{S}$ | 0.0151 | 0.1299 | 0.2517 | 0.5839 |
| $\hat{T}$ | 0.0088 | 0.1465 | 0.2715 | 0.6133 |

## 4-Input Comparison



- Results are consistent with those from the pure sampling-based approach, but requiring far fewer function evaluations.
- Replacing $\mathbf{z}_{1}$ with $s_{1}$ and $\mathbf{r}_{1}$, and $\mathbf{z}_{2}$ with $s_{2}$ and $\mathbf{r}_{2}$ :

|  | $x_{1}$ | $x_{2}$ | $s_{1}$ | $\mathbf{r}_{1}$ | $s_{2}$ | $\mathbf{r}_{2}$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $\hat{S}$ | 0.0062 | 0.0884 | 0.1664 | 0.0662 | 0.5931 | 0.0420 |
| $\hat{T}$ | 0.0057 | 0.1069 | 0.1745 | 0.0782 | 0.6124 | 0.0412 |



## Concluding thoughts:

- Can more bias-resistant alternatives to Sequential Bifurcation be developed for the piecewise constant inputs case (that doesn't require too many runs)?
- Traditional variance-based sensitivity analysis may be most effective if functional inputs can be decomposed into independent (1.) important low-dimensional, and (2.) less important higher-dimensional components.
- Meta-models that are accurate approximations for a moderate sample of functional inputs may improve the efficiency of variance-based sensitivity analysis.


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