

# Active Learning of (small) Quantile Sets

Romain Ait Abdelmalek-Lomenech<sup>1</sup>

*joint work with Julien Bect<sup>1</sup>, Vincent Chabridon<sup>2</sup> & Emmanuel Vazquez<sup>1</sup>*

<sup>1</sup>Université Paris-Saclay, CNRS, CentraleSupélec, L2S  
<sup>2</sup> EDF R&D, PRISME Team

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December 11, Paris



Quantile Set Inversion

SUR methods for Quantile Set Inversion

Estimation of small Quantile Sets

Numerical experiments

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Quantile Set Inversion

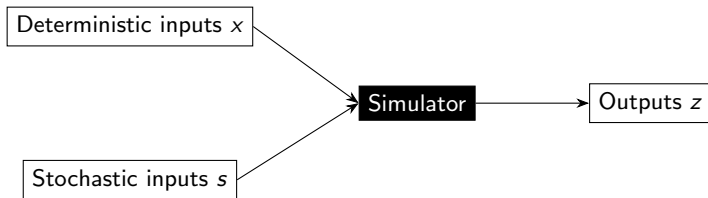
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Consider an **expensive-to-evaluate** numerical simulator  $f$ , with inputs in a set  $\mathbf{U} = \mathbf{X} \times \mathbf{S}$ :

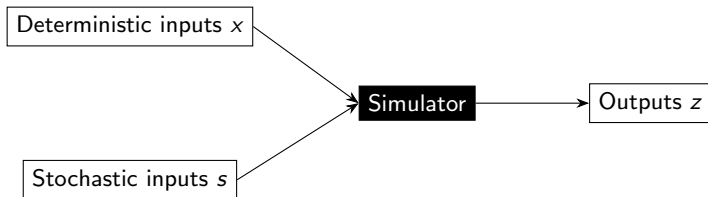
- ▶  $x \in \mathbf{X}$  (deterministic design choices).
- ▶  $s \in \mathbf{S}$  (stochastic factors).



For simplicity we assume a deterministic simulator  $f : \mathbf{U} = \mathbf{X} \times \mathbf{S} \mapsto \mathbb{R}^q$ .

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Given:

- ▶  $C \subset \mathbb{R}^q$  a subset of the outputs space  $\mathbb{R}^q$ .
- ▶  $\alpha \in (0, 1)$  a threshold.
- ▶  $\mathbb{P}_S$  a known distribution on  $\mathcal{S}$ .

We focus on the **quantile set inversion (QSI)** problem:

**Estimate the set** of all  $x \in \mathbb{X}$  such that

$$\mathbb{P}(f(x, S) \in C) \leq \alpha, \quad S \sim \mathbb{P}_S.$$

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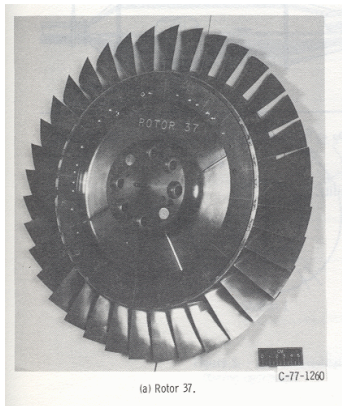
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## An example: the ROTOR37 compressor model

Function  $f : \mathbb{X} \times \mathbb{S} \mapsto \mathbb{R}^3$  with two kind of inputs:

- ▶  $x \in \mathbb{X}$ : design choice for the compressor
- ▶  $s \in \mathbb{S}$ : manufacturing uncertainties, with  $\mathbb{P}_S = \mathcal{U}(\mathbb{S})$ .



Simulator return three outputs:

- ▶  $f_1$ : mass flow
- ▶  $f_2$ : pressure ratio
- ▶  $f_3$ : isentropic efficiency

We can set, for example,  $\alpha = 5\%$  and

$$C = \left\{ z \in \mathbb{R}^3 : \frac{|z_1 - b_1|}{|b_1|} > 0.175 \quad \text{or} \quad \frac{|z_2 - b_2|}{|b_2|} > 0.175 \right\},$$

where  $b_1$  and  $b_2$  are baseline values for the mass flow and pressure ratio of the compressor.

For simplicity, we now assume  $f : \mathbb{U} = \mathbb{X} \times \mathbb{S} \mapsto \mathbb{R}$ , with  $C = (-\infty, T]$ .

The problem becomes

**Estimate the quantile set:**

$$\Gamma(f) = \{x \in \mathbb{X} : \mathbb{P}(f(x, S) \leq T) \leq \alpha\}.$$

**Remark:** With  $C = (-\infty, T]$ , the problem can be seen in term of quantile of  $f(x, S)$ . Indeed

$$x \in \Gamma(f) \quad \iff \quad q_\alpha(f(x, S)) > T,$$

with  $q_\alpha(f(x, S))$  the quantile of order  $\alpha$  of  $f(x, S)$  (with  $S \sim \mathbb{P}_S$ ).

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$$\Gamma(f) = \{x \in \mathbb{X} : \mathbb{P}(f(x, S) \leq T) \leq \alpha\},$$

Example of function and associated quantile set, with  $T = 7.5$  and  $\alpha = 5\%$ .

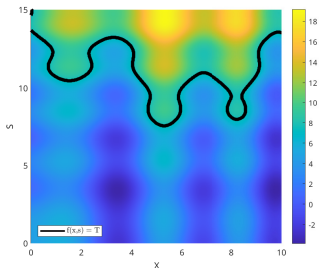


Figure: Representation of the function.

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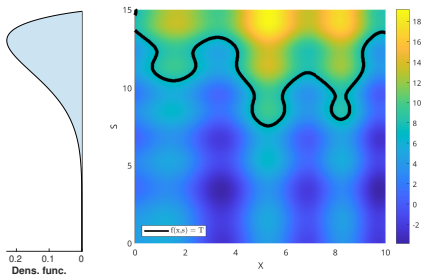


Figure: Representation of the function (right), the density of  $\mathbb{P}_S$  (left)

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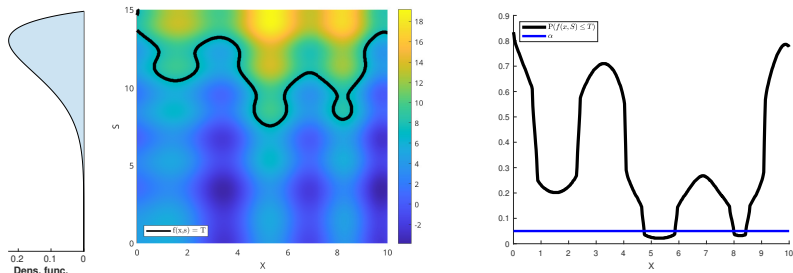


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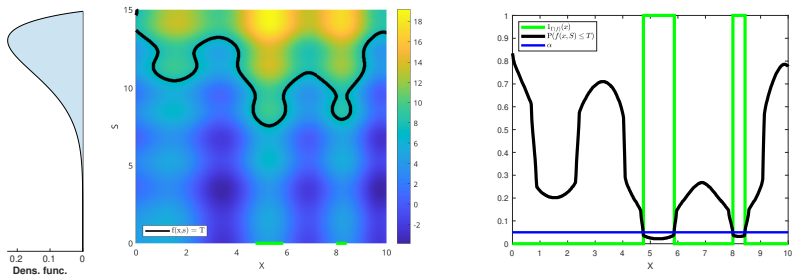


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Given the expensive-to-evaluate nature of the underlying function, it is necessary to evaluate the function at points chosen with attention.

**Active learning** (or sequential design of experiments) approach:

Consider

- ▶  $\mathcal{I}_n = \{(u_1, f(u_1)), \dots, (u_n, f(u_n))\}$  the current information,
- ▶  $a_n(u)$  a sampling criterion dependent on  $\mathcal{I}_n$ .

Until satisfied:

- ▶ Choose  $u \in \mathbb{U}$  as the minimizer (or maximizer) of  $a_n(u)$
- ▶ Evaluate  $f$  at  $u$
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## Bayesian framework & notations:

$f \sim$  GP prior  $\xi$  on  $\mathbb{U} = \mathbb{X} \times \mathbb{S}$ , with constant mean  $\mu$  and covariance  $k$ .

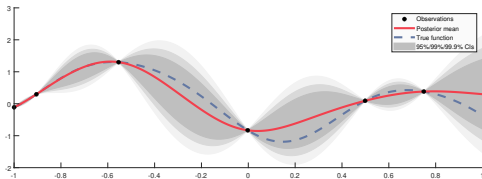


Figure: Illustration of a GP on an interval.

We denote:

- ▶  $\mathbb{P}_n$  and  $\mathbb{E}_n$ : conditional distribution and expectation given  $\mathcal{I}_n$ .
- ▶  $\mu_n$ ,  $\sigma_n$  and  $k_n$ : posterior mean, st. deviation and covariance of  $\xi$ .
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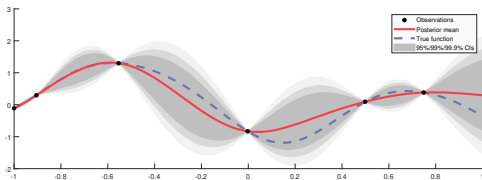


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## First approach: joint-space estimation

The QSI problem is related to the **estimation of the excursion set**

$$\Lambda(f) = \{u \in \mathbb{U} : f(u) \leq T\}$$

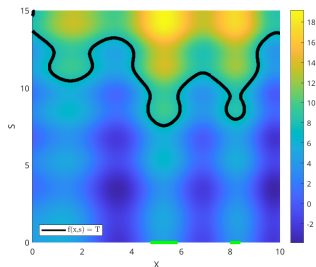


Figure: Example function. The black line delimits the set  $\Lambda(f)$ .

$$x \in \Gamma(f) \iff \mathbb{P}((x, S) \in \Lambda(f)) \leq \alpha,$$

Good approximation of  $\Lambda(f) \implies$  good approximation of  $\Gamma(f)$



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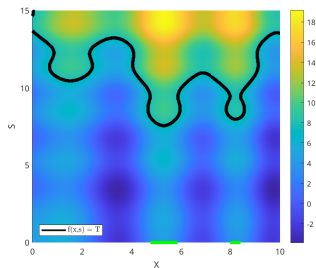


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**Good approximation of  $\Lambda(f)$   $\implies$  good approximation of  $\Gamma(f)$**

Several Bayesian methods focus on **estimating**  $\Lambda(f)$ . For example:

▶ **Maximal uncertainty sampling methods:**

- ▶ Maximum misclassification probability [Bryan et al. (2005)]:

$$U_{n+1} \in \operatorname{argmax}_{u \in \mathcal{U}} \min(p_n(u), 1 - p_n(u))$$

- ▶ [Ranjan et al. (2008); Echard et al. (2011), ... ]

▶ **Stepwise uncertainty reduction (SUR) methods:**

- ▶ For instance [Chevalier et al. (2014)] (Joint-SUR):

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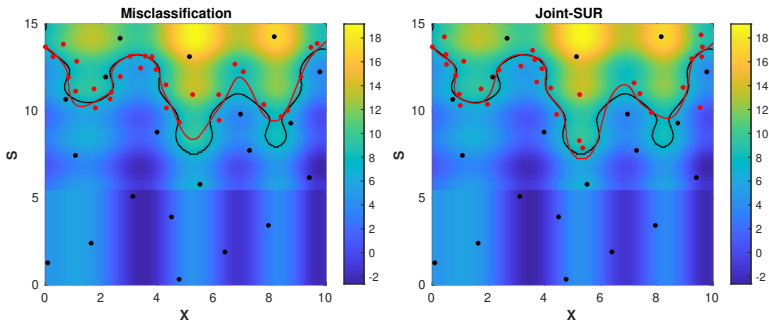


Figure: Examples of designs (red dots) obtained after  $n = 30$  steps with the maximum misclassification and the 'joint-SUR' criteria.

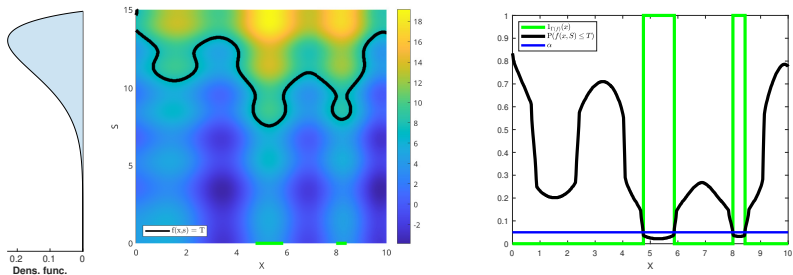


Figure: Representation of the function (middle), the density of  $\mathbb{P}_S$  (left) and associated quantile set (right).

## Second approach: Focusing directly on $\Gamma(f)$

To estimate  $\Gamma(f)$ , one only needs to focus on **'interesting parts'** of  $\Lambda(f)$ .

We denote:

- ▶  $\Gamma(\xi)$  the random quantile set associated to  $\xi$ .
- ▶  $\pi_n(x) = \mathbb{P}_n(x \in \Gamma(\xi))$ , the posterior probability that  $x$  belongs to the (random) quantile set generated by  $\xi$ .
- ▶  $\mathcal{Q}_n = \int_{\mathbb{X}} \min(\pi_n(x), 1 - \pi_n(x)) dx$ .

**QSI-SUR sampling criterion** [Ait Abdelmalek-Lomenech et al. (2024)]:

$$(X_{n+1}, S_{n+1}) \in \underset{(x,s) \in \mathbb{X} \times \mathbb{S}}{\operatorname{argmin}} \mathbb{E}_n(\mathcal{Q}_{n+1} \mid (X_{n+1}, S_{n+1}) = (x, s)),$$

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The implementation proposed in [Ait Abdelmalek-Lomenech et al. (2024)] produces good results on moderately difficult examples.

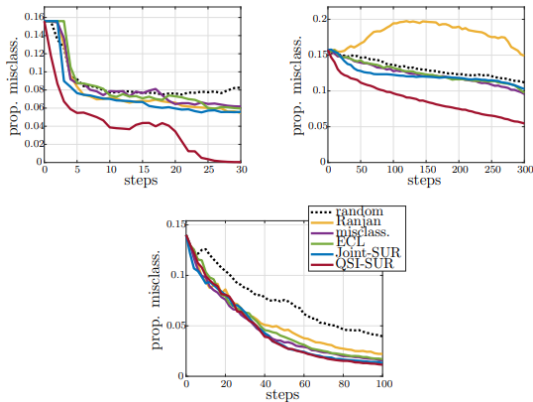


Figure: Median of the proportion of misclassified points vs. number of steps on several examples.

The QSI-SUR criterion focus on part of  $\Lambda(f)$  that gives relevant information on  $\Gamma(f)$ .

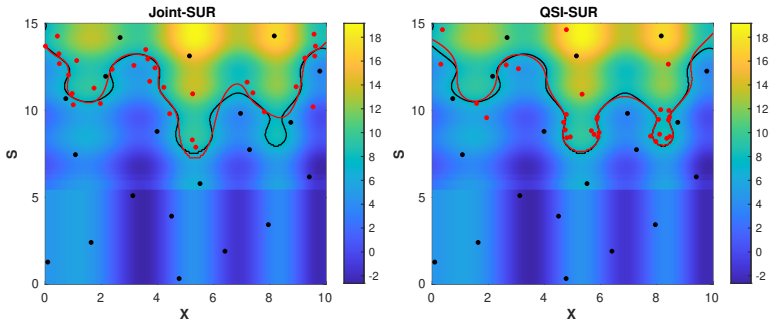


Figure: Example of design obtained with the QSI-SUR criterion and the Joint-SUR criterion.

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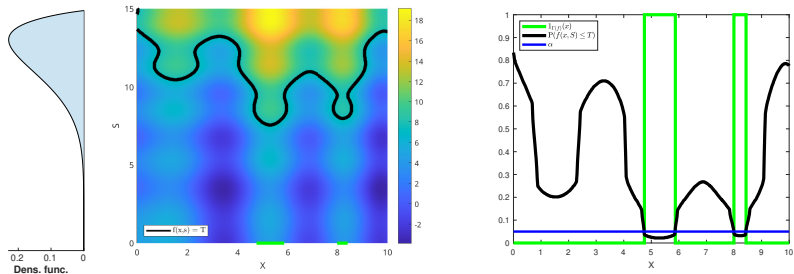


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The QSI-SUR criterion is based on

$$\mathcal{Q}_n = \int_{\mathbb{X}} \min(\pi_n(x), 1 - \pi_n(x)) dx.$$

Two main issues in the implementation:

- ▶ **First issue: Computational complexity.**
  - ▶  $\pi_n$  approximated using conditional sample paths of  $\xi(x, \cdot)$ .
  - ▶ Complexity is  $O(m^3)$ , where  $m$  is the number of points used for the approximation. Due to the Cholesky factorization of the covariance matrix.
  - ▶ Criterion too expensive for continuous optimization and batch design.
  
- ▶ **Second issue: Not adapted to "small"  $\Gamma(f)$ .**
  - ▶ Integral over  $\mathbb{X}$  in  $\mathcal{Q}_n$  is discretized.
  - ▶ Necessity of points (in  $\mathbb{X}$ ) close to the boundary of  $\Gamma(f)$
  - ▶ When  $\Gamma(f)$  is small, importance sampling can prove insufficient.

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To resolve the issues listed previously, we use a two part solution:

▶ **First issue: Computational complexity**

- ▶ We introduce a new type of method called "Maximum Expected Estimator Modification" (MEEM).
- ▶ We derive a MEEM criterion with complexity  $O(m^2)$ .
- ▶ This criterion allows continuous optimization and batch design of experiments.

▶ **Second issue: Not adapted to "small"  $\Gamma(f)$ .**

- ▶ We introduce a sequential Monte Carlo (SMC) framework.
- ▶ We estimate a sequence of decreasing quantile sets, converging towards the set of interest.

## Maximum Expected Estimator Modification (MEEM)

Let us consider a sequence of estimators  $(\widehat{\Gamma}_n)_n$  such that

$$\widehat{\Gamma}_n : \mathcal{I}_n \mapsto \mathcal{P}(\mathbb{X}),$$

and a "distance"

$$d : \mathcal{P}(\mathbb{X})^2 \mapsto \mathbb{R}^+.$$

**MEEM principle:** Choose the point that maximize the expected change in the estimation, i.e.

$$U_{n+1} \in \arg \max_{u \in \mathcal{U}} \mathbb{E}_n(d(\widehat{\Gamma}_{n+1}, \widehat{\Gamma}_n) | U_{n+1} = u)$$

**NB:** For convenience, we assume a batch size of 1.



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## "Duality" SUR / MEEM

Several SUR strategies in the litterature are equivalent to MEEM strategies.

For example:

▶ Optimization:

▶ Expected Improvement

- ▶ MEEM method with  $d(a, b) = |a - b|$  and estimator  $f_n^* = \min\{U_1, \dots, U_n\}$

▶ Function approximation:

- ▶ SUR method :  $U_{n+1} \in \arg \min \mathbb{E}_n \left( \int_{\mathbb{U}} \sigma_n^2(u) du \mid U_{n+1} = u \right)$ .

- ▶ MEEM method with  $d(h, g) = \int_{\mathbb{U}} (h(u) - g(u))^2 du$  and estimator  $f_n = \mu_n$

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▶ MEEM method with  $d(h, g) = \int_{\mathbb{U}} (h(u) - g(u))^2 du$  and estimator

$$f_n = \mu_n$$

## MEEM method for QSI

We choose the divergence  $d(\widehat{\Gamma}_{n+1}, \widehat{\Gamma}_n) = \lambda(\widehat{\Gamma}_{n+1} \Delta \widehat{\Gamma}_n)$  we have

$$d(\widehat{\Gamma}_{n+1}, \widehat{\Gamma}_n) = \int_{\mathbb{X}} |\mathbb{1}_{\widehat{\Gamma}_{n+1}}(x) - \mathbb{1}_{\widehat{\Gamma}_n}(x)| dx,$$

coupled with the sequence of plug-in estimators

$$\widehat{\Gamma}_n = \{x \in \mathbb{X} : \mathbb{P}(\mu_n(x, S) \leq T) \leq \alpha\},$$

We obtain the **QSI-MEEM** strategy:

$$U_{n+1} \in \arg \max_{u \in \mathbb{U}} \int_{\mathbb{X}} \mathbb{E}_n \left( |\mathbb{1}_{\widehat{\Gamma}_{n+1}}(x) - \mathbb{1}_{\widehat{\Gamma}_n}(x)| \mid U_{n+1} = u \right) dx.$$

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Criterion does not need to be approximated using conditional sample paths of  $\xi$ .

As a consequence of the kriging update formula, we have:

### Proposition

Given  $\mathcal{I}_n$  and  $U_{n+1}$ ,  $\hat{\Gamma}_{n+1}$  is a function of a standard Gaussian variable  $Z$ :

$$\hat{\Gamma}_{n+1}(z) = \{x \in \mathbb{X} : P(\mu_n(x, \mathcal{S}) + \kappa_n(x, \mathcal{S})z \leq T) \leq \alpha\},$$

with  $\kappa_n(x, s) = k_n(U_{n+1}, (x, s)) / \sigma_n(U_{n+1})$ .

$\implies$  Computational complexity of the criterion:  $O(m^2)$ , with  $m$  the number of points (in  $\mathbb{X} \times \mathbb{S}$ ) used for the approximation.

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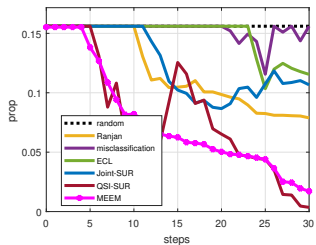
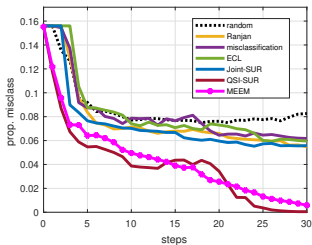
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## Sanity check

We observe that the QSI-MEEM method produces results similar (or better) than the QSI-SUR strategy on case with relatively large quantile sets.



**Figure:** Median (left) and quantile of order 0.9 of the proportion of misclassified points vs. number of steps, for 100 repetitions of the algorithms on the introductory example.

## Second issue: Estimation of small quantile sets

**Idea:** Multilevel splitting/**subset simulation** [Kahn and Harris (1951); Au and Beck (2001)] to efficiently sample points in  $\mathbb{X}$ .

- ▶ Sequentially estimate a sequence of **decreasing quantile sets**

$$\Gamma^0(f) \supset \Gamma^1(f) \supset \dots \supset \Gamma^K(f) = \Gamma(f),$$

using the MEEM strategy described previously.

- ▶ Such sets can be defined by setting

$$\Gamma^k(f) = \{x \in \mathbb{X} : \mathbb{P}(f(x, S) \leq T_k) \leq \alpha\},$$

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We propose a **SMC-based** algorithm inspired by **BSS** [Li (2012); Bect et al. (2017)]

It alternates two distinct phases:

▶ **Estimation phase**

- ▶ Define a new intermediary quantile set to estimate.
- ▶ Sample points  $U_n, \dots, U_{n+r}$  using the MEEM criterion.

▶ **Move phase**

- ▶ Concentrate the particles towards the previously estimated set.

For simplicity, we still assume  $C = (-\infty, T]$  and a batch size of 1.

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Let  $q_{n,k}$  a density targeting  $\Gamma^k(f) = \{x \in \mathbb{X} : \mathbb{P}(f(x, S) \in C_k) \leq \alpha\}$  at step  $n$ .

## Estimation phase:

- Set  $T_{k+1}$  such that

$$\text{ESS} \left( \frac{\mathbb{1}_{\hat{\Gamma}_n^{k+1}}(x)}{\mathbb{1}_{\hat{\Gamma}_n^k}(x)} \right) \approx 30\%.$$

- Sample point

$$U_{n+1} \in \operatorname{argmax} J_n(u),$$

with  $J_n$  the MEEM criterion targeting  $\Gamma^{k+1}(f)$ .

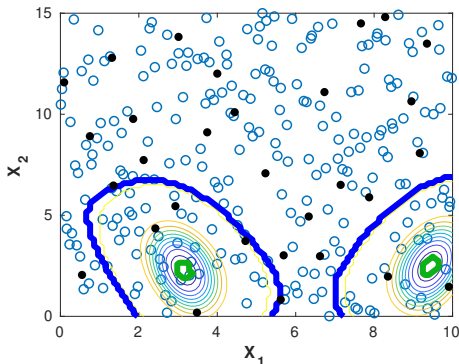


Figure: Temporary quantile set (blue line), final quantile set (green line), particles (blue dots). -  $n = 0$ .

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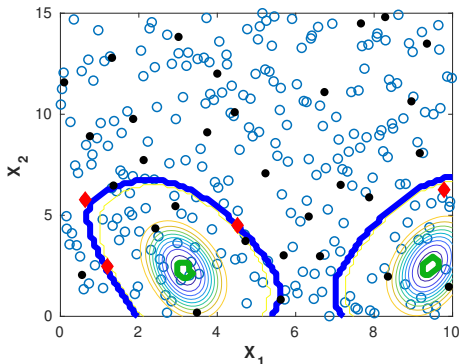


Figure: Temporary quantile set (blue line), final quantile set (green line), particles (blue dots) and projection of the sequential design (red dots). -  $n = 4$ .

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## Move phase:

When stopping condition is met:

- ▶ Residual resampling.
- ▶ Move particles in  $\Gamma^{k+1}(f)$  using MHRW with target density  $q_{n,k+1}$ .
- ▶ Adapt walk's variance to target acceptance rate 25%.

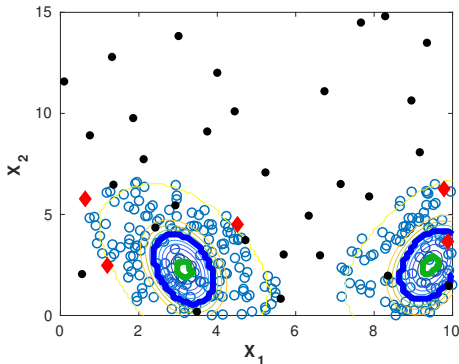


Figure: Temporary quantile set (blue line), final quantile set (green line), particles (blue dots) and projection of the sequential design (red dots). -  $n = 5$ .



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Quantile Set Inversion

SUR methods for Quantile Set Inversion

Estimation of small Quantile Sets

**Numerical experiments**

To evaluate the performances of the proposed strategy, we focus on the relative error

$$\frac{\lambda_{\mathbf{X}}(\Gamma(f)\Delta\hat{\Gamma}_n)}{\lambda_{\mathbf{X}}(\Gamma(f))}$$

obtained at the end of the strategy, and the number of steps required to obtain these results.

The functions considered are modeled by a GP with constant mean and Matern covariance kernel (with  $\nu \in \{1/2, 3/2, 5/2, +\infty\}$ ).

Covariance parameters are estimated at each step using ReML.

Each strategy is repeated 50 times, with different maximin LHS initial designs.

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## 1st example: 5-Trid function [Adorio and U.P. (2005)]

- ▶  $\mathbb{U} = \mathbb{X} \times \mathbb{S}$  with  $\mathbb{X} = [-25, 25]^3$  and  $\mathbb{S} = [-25, 25]^2$ ,
- ▶  $\mathbb{P}_{\mathbb{S}} = \mathcal{U}(\mathbb{S})$
- ▶  $\alpha = 5\%$  and  $C = [1098.5, +\infty)$
- ▶ Relative size of  $\Gamma(f)$ :  $\lambda_{\mathbb{X}}(\Gamma(f)) \sim 10^{-6}$ .

$$f(u) = \sum_{i=1}^5 (u_i - 1)^2 - \sum_{i=2}^5 u_i u_{i-1}$$

Size of the initial design: 50.

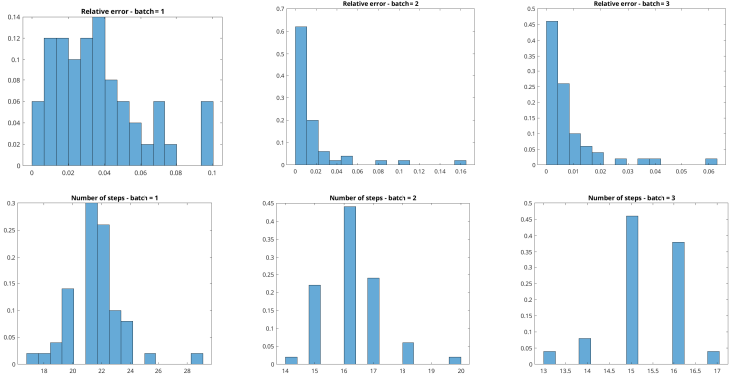


Figure: Distribution of the relative error (top) and the number of steps (bottom), for different batch size (left to right: 1, 2, 3). - (50 runs)

## 2nd example: OTL circuit function [E. N. and D. M. (2007)]

- ▶  $U = \mathbb{X} \times \mathbb{S}$
- ▶  $\mathbb{X} = [-50, 150] \times [25, 70] \times [0.5, 3] \times [1.2, 2.5] \times [0.25, 1.2]$
- ▶  $\mathbb{S} = [-50, 300]$ ,
- ▶  $\mathbb{P}_{\mathbb{S}} = \text{trunc}\mathcal{N}(175, 50)$
- ▶  $\alpha = 5\%$  and  $C = [2.65, +\infty)$
- ▶ Relative size of  $\Gamma(f)$ :  $\lambda_{\mathbb{X}}(\Gamma(f)) \sim 10^{-7}$ .

$f(x, s)$  represents the midpoint voltage of a circuit given the choice of resistances designs  $x$  and the current gain  $s$ .

Size of the initial design: 60.

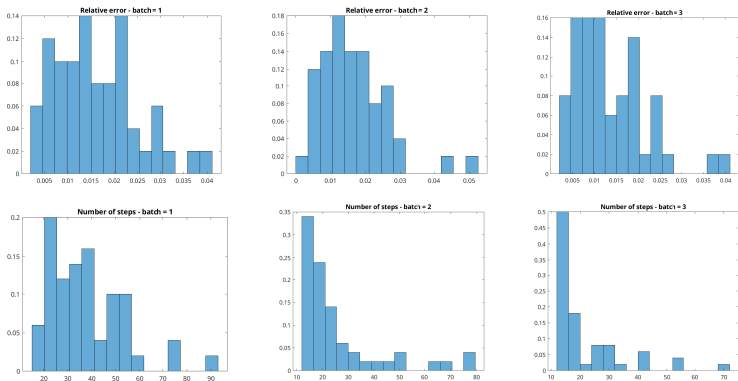


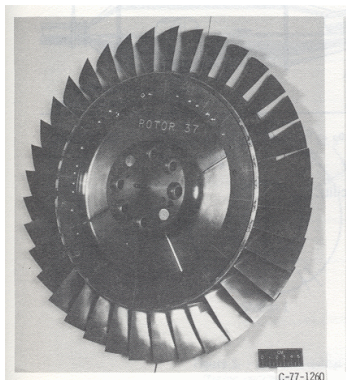
Figure: Distribution of the relative error (top) and the number of steps (bottom), for different batch size (left to right: 1, 2, 3). - (50 runs).

### 3rd example: ROTOR37 model [Reid and Moore (1978)]

Gaussian metamodel (provided by S. Da Veiga and SafranTech) of the ROTOR37 compressor model.

The function  $f : \mathbb{X} \times \mathbb{S} \mapsto \mathbb{R}^3$  takes two kind of inputs:

- ▶  $x \in \mathbb{X} = [0, 1]^{13}$ : design choice for the compressor
- ▶  $s \in \mathbb{S} = [0, 1]^5$ : manufacturing uncertainties ( $\mathbb{P}_{\mathbb{S}} = \mathcal{U}(\mathbb{S})$ )





Simulator returns three outputs:

- ▶  $f_1$ : the mass flow
- ▶  $f_2$ : the pressure ratio
- ▶  $f_3$ : the isentropic efficiency

Goal: finding the set of deterministic design choice leading to values of the mass flow and pressure ratio being close to baselines values ( $b_1, b_2$ ) with sufficiently high probability.

We consider:

- ▶  $C = \left\{ z \in \mathbb{R}^3 : \frac{|z_1 - b_1|}{|b_1|} > 0.175 \quad \text{or} \quad \frac{|z_2 - b_2|}{|b_2|} > 0.175 \right\}$
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We start the strategy from an initial design of size 90. A batch size of 5 is used.

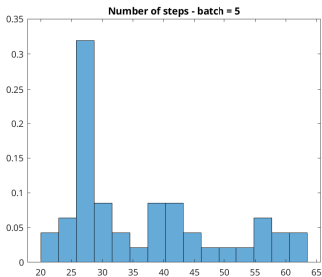
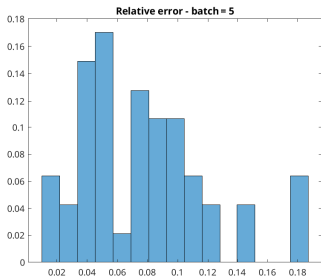


Figure: Distribution of the relative error (left) and the number of steps (right), for a batch size of 5. - (50 runs).

## Conclusion:

- ▶ Introducing the concept of MEEM allows to reproduce (or improve) the results obtained by using the QSI-SUR criterion on moderately difficult examples.
- ▶ The MEEM criterion permits a welcomed gain regarding the computational complexity, due to the absence of conditional Gaussian sample paths
- ▶ Coupled with a SMC framework, the criterion allows to accurately estimate small quantile sets (size of order  $10^{-6} - 10^{-8}$ ).

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Natural idea (in the spirit of [Dubourg et al. (2013); Bect et al. (2017)]):

$$q_{n,k}(x) \propto \pi_n^k(x) = \mathbb{P}_n(x \in \Gamma^k(\xi))$$

- ▶ Does not admit a closed-form expression.
- ▶ Expensive to estimate.

**Idea:** Replace  $\pi_n^k(x)$  by  $\mathbb{1}(x \in \Gamma_{n,k}^+)$ . How to define  $\Gamma_{n,k}^+$ ?

Given  $x_0 \in \mathbb{X}$ ,  $\mu_n$  and  $\sigma_n$  the posterior mean and standard deviation of  $\xi$  and  $\beta \in (1/2, 1)$ , consider the **quantile function**:

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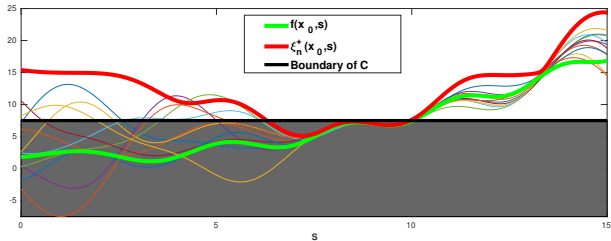


Figure: Example of quantile function  $\xi_n^+(x_0, \cdot)$ , with a fixed  $x_0$ .

Setting  $\Gamma_{n,k}^+ = \Gamma_n^k(\xi_n^+)$  eliminates  $x_0$  if  $\{x_0 \in \Gamma^k(\xi)\}$  is **very improbable**.

We define the target densities as

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**NB:** The MHRW step becomes a constrained random walk.

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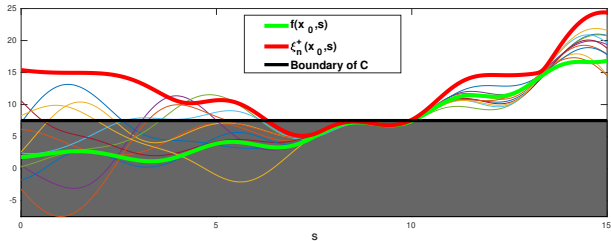


Figure: Example of quantile function  $\xi_n^+(x_0, \cdot)$ , with a fixed  $x_0$ .

Setting  $\Gamma_{n,k}^+ = \Gamma_n^k(\xi_n^+)$  eliminates  $x_0$  if  $\{x_0 \in \Gamma^k(\xi)\}$  is **very improbable**.

We define the target densities as

$$q_{n,k}(x) \propto \mathbb{1}(x \in \Gamma_n^k(\xi_n^+))$$

**NB:** The MHRW step becomes a constrained random walk.

$C = (-\infty, T]$  and  $\xi(x_0, \cdot)$  is a **high quantile**

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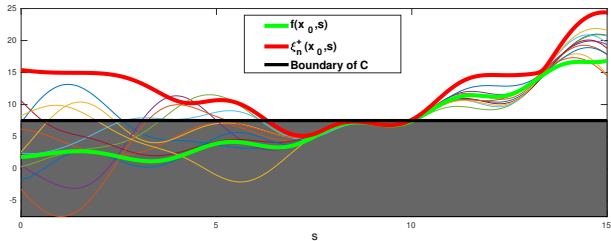


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