

# Variable importance for random forests: a sensitivity analysis perspective

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- State-of-the-art learning algorithms = black-boxes (e.g. random forests, neural networks)

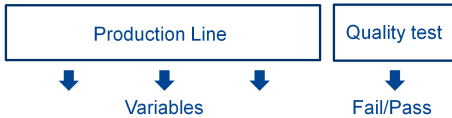
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- Consequence: impossible to grasp how inputs are combined to generate predictions
- Strong limitation for applications with critical decisions at stake (e.g. healthcare, industrial processes)

# Industrial processes

- Context  
Manufacturing process driven by controllable variables.



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Identify production conditions generating defects: variable settings.



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- Objective  
Identify production conditions generating defects: variable settings.
- Method
  - 1 Fit a learning algorithm
  - 2 Use variable importance to detect influential variables
  - 3 Explore associated physical phenomenon with domain experts





# Random forests

- Random forests are an efficient approach
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- Our objective (Bénard et al., 2021)
  - Theoretical analysis of the MDA
    - First convergence result for the original MDA (Ishwaran, 2007; Zhu et al., 2015)
    - Theoretical understanding of MDA bias
  - Design of Sobol-MDA algorithm to fix the MDA flaws

- Regression setting
  - input vector  $\mathbf{X} = (X^{(1)}, \dots, X^{(p)}) \in \mathbb{R}^p$
  - output  $Y \in \mathbb{R}$
  - dataset  $\mathcal{D}_n = \{(\mathbf{X}_i, Y_i), i = 1, \dots, n\}$ ,  
where  $(\mathbf{X}_i, Y_i) \sim \mathbb{P}_{\mathbf{X}, Y}$ .

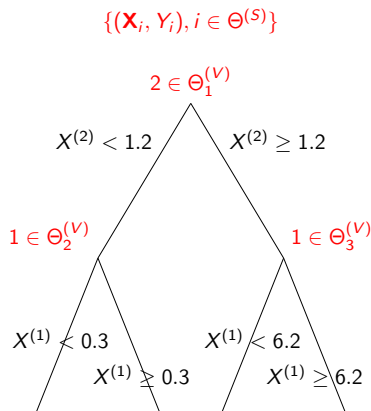
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- Random forest algorithm

- Aggregation of  $\Theta$ -random trees  
 $\Theta = (\Theta^{(S)}, \Theta^{(V)})$
- $M$ : number of trees
- $m_{M,n}(\mathbf{X}, \Theta_M)$ : the forest estimate at  $\mathbf{X}$



- 1 Introduction
- 2 MDA Theoretical Limitations
  - MDA definition
  - MDA convergence
- 3 Sobol-MDA

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break the dependence between  $X^{(j)}$  and  $Y$
- 4 compute the decrease of accuracy of the forest with the permuted data

$X^{(1)}$	$X^{(2)}$	...	$X^{(j)}$	...	$X^{(p)}$	$Y$
2.1	4.3	...	0.1	...	2.6	2.3
1.7	4.1	...	9.2	...	3.8	0.4
3.4	9.2	...	3.2	...	3.6	10.2
5.6	1.2	...	8.2	...	4.2	9.1
8.9	6.8	...	6.7	...	2.9	4.5

**Table:** Example of the permutation of a dataset  $\mathcal{D}_n$  for  $n = 5$ .

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Explained variance of Y = 16.4

Explained variance of Y = 13.7

$$\text{MDA}(X^{(j)}) = 16.4 - 13.7 = 2.7$$

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Question: Can I use  $\mathcal{D}_n$  to both fit the forest and compute accuracy ?

No: overfitting and inflated accuracy.

How to handle this in practice?



The explained variance estimate of MDA algorithms differ across implementations

**Train-Test MDA:** train data to fit the forest, and test data for accuracy

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**Out-of-bag (OOB) samples:**  $\mathcal{D}_n$  is bootstrap prior to the construction of each tree, leaving aside a portion of  $\mathcal{D}_n$ , which is not involved in the tree growing and defines the “out-of-bag” sample.

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Selected samples:  $\Theta_\ell^{(S)} = \{1, 3, 4\}$

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OOB samples:  $\{1, \dots, n\} \setminus \Theta_\ell^{(s)} = \{2, 5\}$

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MDA Version	Package	Error	Data
Train-Test	scikit-learn randomForestSRC	Forest	Testing dataset
Breiman-Cutler	randomForest (normalized) ranger / randomForestSRC	Tree	OOB sample
Ishwaran-Kogalur	randomForestSRC	Forest	OOB sample

Table: Summary of the different MDA algorithms.

- $i \in \{1, \dots, n\} \setminus \Theta_\ell^{(S)} = \{2, 5\}$ : OOB sample of the  $\ell$ -th tree

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# Breiman-Cutler MDA

- $i \in \{1, \dots, n\} \setminus \Theta_\ell^{(S)} = \{2, 5\}$ : OOB sample of the  $\ell$ -th tree
- $N_{n,\ell} = \sum_{i=1}^n \mathbb{1}_{i \notin \Theta_\ell^{(S)}} = 2$ : size of the OOB sample of the  $\ell$ -th tree

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$\mathbf{X}_i$

$\mathbf{X}_{i,\pi_{j\ell}}$

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$$\widehat{\text{MDA}}_{M,n}^{(BC)}(\mathbf{X}^{(j)}) = \frac{1}{M} \sum_{\ell=1}^M \frac{1}{N_{n,\ell}} \sum_{i=1}^n [(Y_i - m_n(\mathbf{X}_{i,\pi_{j\ell}}, \Theta_\ell))^2 - (Y_i - m_n(\mathbf{X}_i, \Theta_\ell))^2] \mathbb{1}_{i \notin \Theta_\ell^{(S)}}$$



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Quadratic risk of the  $\ell$ -th tree

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Inflated quadratic risk of the  $\ell$ -th tree where  $\mathbf{X}^{(j)}$  is permuted

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Risks are computed over the OOB sample of each tree

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Average over all trees

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(A1)

The response  $Y \in \mathbb{R}$  follows

$$Y = m(\mathbf{X}) + \varepsilon$$

where

- $\mathbf{X} = (X^{(1)}, \dots, X^{(p)}) \in [0, 1]^p$
- $\mathbf{X}$  admits a density  $f$  such that  $c_1 < f(\mathbf{x}) < c_2$ , with constants  $c_1, c_2 > 0$
- $m$  is continuous
- the noise  $\varepsilon$  is sub-Gaussian and centered

# Assumptions

(A2): the theoretical tree is consistent  
(always true with slight modifications of the forest algorithm)

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*The randomized theoretical CART tree built with the distribution of  $(\mathbf{X}, Y)$  is consistent, that is, for all  $\mathbf{x} \in [0, 1]^p$ , almost surely,*

$$\lim_{k \rightarrow \infty} \Delta(m, A_k^*(\mathbf{x}, \Theta)) = 0.$$



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(A3): tree partition is not too complex with respect to  $n$

(A3)

*The asymptotic regime of  $a_n$ , the size of the subsampling without replacement, and the number of terminal leaves  $t_n$  is such that  $a_n \leq n - 2$ ,  $a_n/n < 1 - \kappa$  for a fixed  $\kappa > 0$ ,  $\lim_{n \rightarrow \infty} a_n = \infty$ ,  $\lim_{n \rightarrow \infty} t_n = \infty$ , and  $\lim_{n \rightarrow \infty} t_n \frac{(\log(a_n))^9}{a_n} = 0$ .*

## Theorem (Bénard et al. (2021))

If Assumptions (A1), (A2), and (A3) are satisfied, then, for all  $M \in \mathbb{N}^*$  and  $j \in \{1, \dots, p\}$  we have

$$\widehat{MDA}_{M,n}^{(BC)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} \mathbb{E}[(m(\mathbf{X}) - m(\mathbf{X}_{\pi_j}))^2]$$

$\mathbf{X}_{\pi_j}$ :  $\mathbf{X}$  where the  $j$ -th component is replaced by an independent copy, i.e.

$$\mathbf{X}_{\pi_j} = (X^{(1)}, \dots, X^{(j)}, \dots, X^{(p)})$$

**Limit interpretation?**

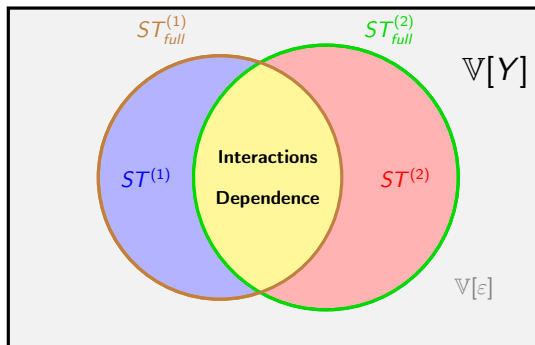


Figure: Standard and full total Sobol indices for  $Y = m(\mathbf{X}^{(1)}, \mathbf{X}^{(2)}) + \varepsilon$ .

**Total Sobol index** (Sobol, 1993)

$$ST^{(1)} = \frac{\mathbb{E}[V(m(\mathbf{X})|\mathbf{X}^{(-1)})]}{V(Y)}$$

**Full total Sobol index** (Mara et al., 2015; Benoumechiara, 2019)

$$ST_{full}^{(1)} = \frac{\mathbb{E}[V(m(\mathbf{X}_{\pi_j})|\mathbf{X}^{(-1)})]}{V(Y)}$$

## Proposition (Bénard et al. (2021))

If Assumptions (A1), (A2) and (A3) are satisfied, then for all  $M \in \mathbb{N}^*$  and  $j \in \{1, \dots, p\}$  we have

$$\widehat{MDA}_{M,n}^{(BC)}(\mathbf{X}^{(j)}) \xrightarrow{\mathbb{L}^1} \mathbb{V}[Y] \times ST^{(j)} + \mathbb{V}[Y] \times ST_{full}^{(j)} + MDA_3^{*(j)}.$$

The term  $MDA_3^{*(j)}$  is not an importance measure and is defined by

$$MDA_3^{*(j)} = \mathbb{E}[(\mathbb{E}[m(\mathbf{X})|\mathbf{X}^{(-j)}] - \mathbb{E}[m(\mathbf{X}_{\pi_j})|\mathbf{X}^{(-j)}])^2].$$

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$$\begin{aligned} (i) \quad & \widehat{MDA}_{M,n}^{(TT)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} \mathbb{V}[Y] \times ST^{(j)} + \mathbb{V}[Y] \times ST_{full}^{(j)} + MDA_3^{*(j)} \\ (ii) \quad & \widehat{MDA}_{M,n}^{(BC)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} \mathbb{V}[Y] \times ST^{(j)} + \mathbb{V}[Y] \times ST_{full}^{(j)} + MDA_3^{*(j)}. \end{aligned}$$

If additionally  $M \rightarrow \infty$ , then

$$(iii) \quad \widehat{MDA}_{M,n}^{(IK)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} \mathbb{V}[Y] \times ST^{(j)} + MDA_3^{*(j)}.$$

**If inputs  $\mathbf{X}$  are independent:**  $\text{MDA}_3^{*(j)} = 0$  and  $ST^{(j)} = ST_{full}^{(j)}$ .

Corollary (Bénard et al. (2021))

*If  $\mathbf{X}$  has independent components, and if Assumptions (A1)-(A3) are satisfied, for all  $M \in \mathbb{N}^*$  and  $j \in \{1, \dots, p\}$  we have*

$$\widehat{\text{MDA}}_{M,n}^{(TT)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} 2\mathbb{V}[Y] \times ST^{(j)}$$
$$\widehat{\text{MDA}}_{M,n}^{(BC)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} 2\mathbb{V}[Y] \times ST^{(j)}.$$

*If additionally  $M \rightarrow \infty$ , then*

$$\widehat{\text{MDA}}_{M,n}^{(IK)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} \mathbb{V}[Y] \times ST^{(j)}.$$

This Corollary completes the result from (Gregorutti, 2015).

# Additive regression function

If  $m$  is additive:  $\text{MDA}_3^{*(j)} = 0$ .

Corollary (Bénard et al. (2021))

If the regression function  $m$  is additive, and if Assumptions (A1)-(A3) are satisfied, for all  $M \in \mathbb{N}^*$  and  $j \in \{1, \dots, p\}$  we have

$$\widehat{\text{MDA}}_{M,n}^{(TT)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} \mathbb{V}[Y] \times ST^{(j)} + \mathbb{V}[Y] \times ST_{full}^{(j)}$$
$$\widehat{\text{MDA}}_{M,n}^{(BC)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} \mathbb{V}[Y] \times ST^{(j)} + \mathbb{V}[Y] \times ST_{full}^{(j)}.$$

If additionally  $M \rightarrow \infty$ , then

$$\widehat{\text{MDA}}_{M,n}^{(IK)}(X^{(j)}) \xrightarrow{\mathbb{L}^1} \mathbb{V}[Y] \times ST^{(j)}.$$



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- We develop the Sobol-MDA: a fast and consistent estimate of  $ST^{(j)}$  for random forests

## 1 Introduction

## 2 MDA Theoretical Limitations

- MDA definition
- MDA convergence

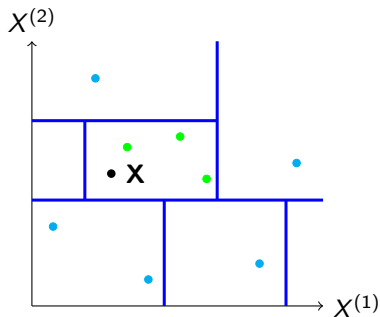
## 3 Sobol-MDA

Principle: **project** the partition of each tree along the  $j$ -th direction to remove  $X^{(j)}$  from the prediction process.

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$$\widehat{\text{S-MDA}}_{M,n}(X^{(j)}) = \frac{1}{\hat{\sigma}_Y^2} \frac{1}{n} \sum_{i=1}^n \left[ Y_i - m_{M,n}^{(-j, OOB)}(\mathbf{X}_i^{(-j)}, \Theta_M) \right]^2 - \left[ Y_i - m_{M,n}^{(OOB)}(\mathbf{X}_i, \Theta_M) \right]^2$$

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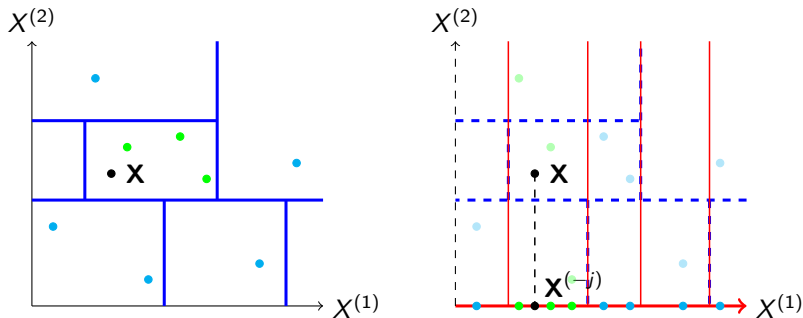


**Figure:** Example of the partition of  $[0, 1]^2$  by a random tree (left side) projected on the subspace span by  $\mathbf{X}^{(-2)} = X^{(1)}$  (right side), for  $p = 2$  and  $j = 2$ .

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The Sobol-MDA recovers the appropriate theoretical counterpart for variable selection: the total Sobol index

Theorem (Bénard et al. (2021))

*If Assumptions (A1), (A2'), and (A3') are satisfied, for all  $M \in \mathbb{N}^*$  and  $j \in \{1, \dots, p\}$*

$$\widehat{S\text{-MDA}}_{M,n}(X^{(j)}) \xrightarrow{P} ST^{(j)}.$$

Settings (Archer and Kimes, 2008; Gregorutti et al., 2017)

- $p = 200$  input variables
- 5 independent groups of 40 variables
- each group is a Gaussian vector, strongly correlated

# Sobol-MDA Experiments

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- 1 variable from each group involved in  $m$

$$m(\mathbf{X}) = 2X^{(1)} + X^{(41)} + X^{(81)} + X^{(121)} + X^{(161)}.$$

- independent Gaussian noise with  $\mathbb{V}[\varepsilon] = 10\% \mathbb{V}[Y]$

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- $n = 1000$  observations
- $M = 300$  trees

# Sobol-MDA Experiments

$\widehat{\text{S-MDA}}$		$\widehat{\text{BC-MDA}}/2\mathbb{V}[Y]$		$\widehat{\text{IK-MDA}}/\mathbb{V}[Y]$	
$\mathbf{X}^{(1)}$	0.035	$\mathbf{X}^{(1)}$	0.048	$\mathbf{X}^{(1)}$	0.056
$\mathbf{X}^{(161)}$	0.005	$\mathbf{X}^{(25)}$	0.010	$\mathbf{X}^{(5)}$	0.009
$\mathbf{X}^{(81)}$	0.004	$\mathbf{X}^{(31)}$	0.008	$\mathbf{X}^{(81)}$	0.007
$\mathbf{X}^{(121)}$	0.004	$\mathbf{X}^{(14)}$	0.008	$\mathbf{X}^{(41)}$	0.005
$\mathbf{X}^{(41)}$	0.002	$\mathbf{X}^{(40)}$	0.007	$\mathbf{X}^{(161)}$	0.005
$\mathbf{X}^{(179)}$	0.002	$\mathbf{X}^{(3)}$	0.007	$\mathbf{X}^{(15)}$	0.005
$\mathbf{X}^{(13)}$	0.001	$\mathbf{X}^{(17)}$	0.006	$\mathbf{X}^{(121)}$	0.005
$\mathbf{X}^{(25)}$	0.001	$\mathbf{X}^{(26)}$	0.006	$\mathbf{X}^{(7)}$	0.005
$\mathbf{X}^{(73)}$	0.001	$\mathbf{X}^{(41)}$	0.006	$\mathbf{X}^{(4)}$	0.004
$\mathbf{X}^{(155)}$	0.001	$\mathbf{X}^{(121)}$	0.006	$\mathbf{X}^{(28)}$	0.004

Table: Sobol-MDA, normalized BC-MDA, and normalized IK-MDA estimates with influential variables in blue.

Additional experiments are available in B nard et al. (2021)  
(non-linear data with interactions and dependence)

- analytical example
- backward variable selection with real data

# Conclusion

- Strong connections between the MDA and Sobol indices
- MDA does not target the appropriate quantity



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- MDA does not target the appropriate quantity
- Sobol-MDA fixes the flaws of original MDA
- R/C++ package SobolMDA, available online on Gitlab (<https://gitlab.com/drti/sobolmda>), and based on the package `ranger`
- Sobol-MDA can be associated with any black-box algorithm
  - fit a black box  $\hat{f}$  on  $\mathcal{D}_n$
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- Perspectives: generalization to Shapley effects

# Questions ?



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